

Coastal fish dynamics across scales: roles of warming and restoration, and interactions with
fisheries

by

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1 **Abstract**

2 Coastal fish population dynamics are highly variable in time and space, and in the Mid-
3 Atlantic Bight (MAB) region of the U.S. Northeast Continental Shelf, many species migrate
4 seasonally and ontogenically between the coastal zone and offshore shelf habitats. The seasonal
5 timing of migrations and spawning, the survivorship of larvae and juveniles, the locations of
6 inshore habitats that fishes move into, and the presence and quality of nursery habitats like
7 seagrasses impact the abundances and species diversities of inshore fish communities and have
8 implications for the small-scale commercial fisheries that depend on them. In Chapter 1 of this
9 dissertation, I explored how ocean warming on the continental shelf due to climate change has
10 impacted the abundances of fishes in inshore habitats across the MAB. I found that, on the
11 spatial scale of the MAB, the population responses of fishes to warming were weakly mediated
12 by life history characteristics, but that smaller-bodied, faster-lived species whose populations
13 responded positively to warming tended to be warm-adapted, while populations of larger-bodied,
14 slower-lived species responding negatively were cool-adapted. In Chapter 2, I evaluated how
15 seasonal species asynchrony—the temporal offset in the population dynamics of species
16 biomasses—related to seasonal asynchrony in the harvests of those species, and how harvest
17 asynchrony contributed to the stability of small-scale commercial fisheries operating in the
18 Chesapeake Bay (USA).

19 I showed that harvest asynchrony and stability were associated with both social and
20 ecological dynamics, including fisher behaviors in response to management restrictions and
21 changes in effort, and to the availability and variability of targeted species populations that
22 contributed to species asynchrony. In Chapter 3, I studied the impacts of seagrass restoration on

23 the abundance and diversity of juvenile coastal fishes. I showed that fish communities within
24 restored meadows were more abundant and diverse compared to nearby unvegetated habitats,
25 further bolstering the case that seagrass restoration is a worthy endeavor. Together, this body of
26 work provides novel insights into the impacts of continental shelf warming on inshore fish
27 populations in the MAB, the consequences of population variability on small-scale commercial
28 fisheries, and how seagrass restoration can enhance the function of inshore ecosystems.

29 Introduction

30 Marine ecosystems are changing rapidly in response to climate change, leading to shifts
31 in the distribution (Kleisner et al., 2016; Pinsky et al., 2013), productivity (Free et al., 2019;
32 Friedland, Ganley, et al., 2023), and phenology (Asch et al., 2019; Walsh et al., 2015) of fishes
33 worldwide. These impacts are well-documented on regional scales in continental shelf
34 ecosystems (Asch et al., 2019; Free et al., 2019; Simpson et al., 2011), but have been less well
35 studied on regional scales among inshore habitats like estuaries and coastal bays (Fujiwara et al.,
36 2022; Oke et al., 2022). In general, the causes and consequences of variability in inshore fish
37 population dynamics are important to understand because many species of commercially
38 harvested fish inhabiting shelf habitats have migratory life histories, spending a substantial
39 portion of their lives inshore where they also support large commercial and recreational fisheries
40 (Schonfeld et al., 2022). A critical part of the life history of many such species that influences the
41 variability of harvestable biomass in both inshore and offshore systems is the presence of inshore
42 nursery habitats like seagrass meadows that enhance the growth and survival of larval and
43 juvenile fishes (Lefcheck et al., 2019a).

44 In this dissertation, I explored the relationships between inshore fish populations, their
45 habitats and environmental drivers, and the commercial fisheries that target them on spatial
46 scales ranging from the tens to hundreds of kilometers. In Chapter 1, I tested the hypothesis that
47 the responses of inshore fish populations to continental shelf sea-surface temperature (SST)
48 anomalies were mediated by their positioning along the fast-slow continuum of life history
49 strategies, an important biological characteristic of fishes that influences species' capacities to
50 respond to climate change (Beukhof et al., 2019). This analysis considered inshore fish
51 populations along the entire length of the Mid-Atlantic Bight (MAB) region of the Northeast
52 U.S. Continental Shelf (NES) in the western Atlantic Ocean. In Chapter 2, I evaluated how the
53 within-year asynchrony of inshore fish populations influenced the asynchrony and stability of
54 commercial fisheries targeting those populations and their overall value in Chesapeake Bay,
55 USA. Lastly, in Chapter 3, I examined the habitat provisioning function of restored seagrass for
56 juvenile fishes relative to unvegetated habitats in the coastal bays of Virginia, USA.

57 The NES is warming rapidly (Friedland, Morse, et al., 2020) and as a result, populations
58 of many fishes have shifted poleward and deeper (Friedland et al., 2021; Kleisner et al., 2016),
59 while also becoming smaller-bodied and more abundant (Friedland, Ganley, et al., 2023). How
60 the impacts of a warming ocean are reflected in the population dynamics of fishes in inshore
61 habitats in the region is largely unknown, although recent work has found inshore community
62 abundances and diversities to have increased with oceanic SSTs (Oke et al., 2022). In Chapter 1,
63 I tested the hypothesis that the position of species along the fast-slow life history continuum
64 (which I quantified as a composite index of species-specific natural mortalities, asymptotic
65 biomasses, and maximum ages) mediated the inshore population responses of fishes to seasonal
66 SST anomalies in offshore habitats in the MAB. I based this hypothesis on research from

67 continental shelf ecosystems, which has shown that population distributions of faster-lived
68 (smaller-bodied, shorter-lived) species have a greater capacity to respond to warming waters, in
69 terms of distribution shifts or changes in abundances, than those of slower-lived species (larger-
70 bodied, longer-lived) (Free et al., 2019; Perry et al., 2005; Simpson et al., 2011).

71 My results from this study showed that, first, there was a weakly mediating effect of the
72 fast-slow continuum on inshore population responses to seasonal warming in offshore habitats.
73 This meant that fish populations responding positively to warming tended to be faster-lived, and
74 slower-lived responding negatively. This effect was weak, and further analyses indicated that
75 there was high bay to bay variability in population responses to warming within species.
76 However, in evaluating the populations of species who were faster-lived, we found that faster-
77 lived species whose populations were responding positively to seasonal warming tended to be
78 warm-adapted, while slower-lived fishes responding negatively to seasonal warming tended to be
79 cool-adapted. This finding suggests that increased fish abundances in the inshore habitats of the
80 MAB in response to SST anomalies (Oke et al., 2022) could be due to the population responses
81 of primarily faster-lived species in the southerly, subtropical portion of the system. This work is
82 among the first to explore how the responses of inshore fish populations to offshore SST
83 anomalies are mediated by fish life history characteristics on a macroecological scale and
84 provides important insights into how climate change is likely to shape future inshore fish
85 assemblages in the MAB.

86 A key finding from this chapter was that the responses of inshore fish populations to
87 offshore warming were highly variable across habitats within species. This suggests that, while
88 some species were predicted to benefit from a warming ocean, other processes that went
89 uncharacterized in this study (e.g., environmental conditions and species interaction) were likely

90 highly important to “realized” population responses to warming. Understanding these dynamics
91 has large implications for small-scale commercial fisheries in the MAB (White & Scheld, 2021)
92 that directly or indirectly depend on these populations (e.g., through trophic linkages). This is
93 because, for fishers operating on small spatial scales, declines in the local abundances of target
94 species may be difficult to overcome if fishers lack the operational capacity to diversify their
95 targets to different stocks that are abundant outside of historically fished areas (Thorson et al.,
96 2018). However, fish populations in the inshore habitats of the Mid-Atlantic tend to fluctuate
97 asynchronously within years according to species-specific seasonal migratory dynamics, and
98 fishers may be able to stabilize their yields over the course of the year by asynchronously
99 targeting these stocks as their availabilities fluctuate seasonally.

100 I explored these dynamics in Chapter 2 from the level of the regional commercial fishing
101 harvest portfolio. Specifically, I evaluated how the temporal offset of seasonal fish population
102 dynamics among species, or species asynchrony, in the Maryland and Virginia regions of the
103 Chesapeake Bay (USA) contributed to asynchrony in the commercial fishing harvests drawn
104 from these species. I then evaluated how seasonal harvest asynchrony contributed to the within-
105 year economic stability and value of commercial harvest portfolios. This analysis ultimately
106 linked species asynchrony occurring due to the seasonal migratory patterns of biomass-dominant
107 species and interannual changes in their relative abundances to the economic stability of regional
108 commercial fisheries.

109 Results from this study showed that both ecological and social dynamics were important
110 contributing factors to harvest asynchrony and harvest portfolio stability. In Maryland, where the
111 harvest of the most valuable species, striped bass *Morone saxatilis*, was restricted for part of the
112 year, we found that seasonal harvest asynchrony and stability emerged due to effort

113 compensation, or fishing effort being directed towards other species during the striped bass
114 closure. However, these harvest portfolio effects did not stabilize the within-year economic value
115 of the fishing portfolio likely due to the relatively low value of targeted stocks outside of striped
116 bass in MD waters.

117 In Virginia, there were no analogous seasonal closures that significantly impacted
118 harvests, resulting in the emergence of harvest asynchrony as fishers preferentially targeted
119 species as they became seasonally abundant. Despite this seasonal harvest compensation that
120 supported harvest stability, a decline in the abundance of fish entering into the Bay seasonally
121 from the continental shelf (Schonfeld et al., 2022) counteracted seasonal harvest compensation
122 and, in tandem with declining fishing effort, drove down the stability of the Virginia harvest
123 portfolio. However, there exists a high diversity of commercial fishing opportunities outside of
124 the demersal finfish species discussed here, as well as in related industries like seafood sales and
125 processing, that are viable alternatives. Frequently, fishers already retain the necessary licenses
126 that would allow them to take advantage of these opportunities (White & Scheld, 2021). For
127 instance, wild caught shellfish and shellfish aquaculture, the blue crab *Callinectes sapidus*
128 fishery, or the emerging commercial fishery surrounding the invasive blue catfish *Ictalurus*
129 *furcatus* (Fabrizio et al., 2021; White & Scheld, 2021).

130 Ultimately, fisheries require fish populations to draw from, and the abundances of these
131 populations are enhanced by inshore nursery habitats that support the survival and growth of
132 larvae and juveniles prior to joining the adult population (Lefcheck et al., 2019a). In the MAB,
133 seagrasses form undersea meadows that serve as important nursery habitats for larval and
134 juvenile fishes. In the coastal bays of Virginia—the focus of my third chapter—seagrass was
135 extirpated in the 1930s following wasting disease and hurricane disturbance (Orth &

136 McGlathery, 2012). Over six decades later, in the late 1990s, a seed-based restoration program
137 was undertaken that by 2021 led to the re-establishment of ~36 km² of seagrass meadow in the
138 bays (Orth et al., 2012). In Chapter 3, I examined the habitat provisioning function of the
139 restored seagrass meadow for juvenile fish relative to nearby unvegetated areas and the direct
140 and indirect effects of landscape-scale environmental variables on seagrass presence and fish
141 community abundance and diversity. Using long-term survey data, I discovered that juvenile fish
142 abundance and diversity during summer were substantially higher inside the meadow than in
143 unvegetated areas, and that this effect was consistent despite strong interannual variation in fish
144 counts. I also found that water residence time and depth were important predictors of seagrass
145 presence, showing how environmental variability contributed to restoration success and the
146 subsequent ecosystem function of restored habitat.

147 In this body of work, I explored how climate and habitat influence the dynamics of
148 inshore fish populations along the coastline of the MAB, with implications for the impacts of
149 warming oceans on fish communities and the stability of the fisheries that depend upon them. I
150 showed how emergent properties of commercial fisheries in Chesapeake Bay, i.e., stability and
151 asynchrony, are related to species asynchrony in the fish community and the behaviors of fishers
152 in response to seasonal harvest restrictions. Lastly, I showed that restored seagrass habitat is
153 important habitat for a diverse and abundant community of coastal fishes.

154 Chapter 1

155

156 *Evidence for a weakly mediating effect of life history characteristics on inshore fish population*

157 *responses to offshore warming in the Mid-Atlantic Bight*

158

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166 Abstract

167 Marine ecosystems are experiencing rapid and unprecedented changes in response to
168 ocean warming. The consequences of such warming for fishes, including altered population
169 distributions, productivities, and phenologies, are usually studied in the context of continental
170 shelf ecosystems where large-scale commercial fisheries operate and where governments support
171 spatially comprehensive bottom trawl surveys. However, fishes in temperate continental shelf
172 ecosystems often migrate seasonally to and from the coastal zone where population responses to
173 warming are less well-understood. We explored how seasonal offshore sea-surface temperature
174 (SST) anomalies across the Mid-Atlantic Bight region of the Northeast U.S. Continental Shelf
175 impacted the dynamics of autumn fish assemblages in inshore habitats across 42 years (1982–
176 2023), and how these effects were mediated by species’ positioning along the “fast-slow” life
177 history continuum. We found that the responses of fish populations to SST anomalies in the
178 spring and summer months were weakly mediated by species’ positioning along the fast-slow
179 continuum ($P_{\text{Spring}} = 0.08$, $P_{\text{Summer}} = 0.09$), meaning that inshore populations of smaller-bodied
180 and shorter-lived species tended to respond positively to seasonal warming, and larger-bodied
181 and longer-lived species negatively. However, there was uncertainty in these relationships
182 following bay-to-bay variability in population SST responses. We also found that the direction of
183 population responses to warming (positive or negative) tended to fall along a gradient of species
184 thermal affinities—faster-lived populations who responded positively to warming tended to be
185 warm-adapted species, and slower-lived populations responding negatively, cool-adapted. These
186 findings showed that life history characteristics contributed to how inshore fish populations
187 responded to warming, but also that the effects of warming are unlikely to be consistent across

188 inshore populations subjected to highly variable environments and species interactions impacting
189 population abundances.

190 **Introduction**

191 Globally, oceans are warming rapidly due to climate change with far-reaching
192 consequences for the structure and function of marine ecosystems and the human systems that
193 depend upon them (Baudron et al., 2020; Kleisner et al., 2017). For instance, ocean warming is
194 responsible for expansions and contractions of fish species distributions into more suitable
195 thermal habitats (Dulvy et al., 2008; Kleisner et al., 2016; Pinsky et al., 2013), contributing to
196 changes in species richness (Batt et al., 2017; Friedland, Langan, et al., 2020), productivity (Free
197 et al., 2019; Friedland, Langan, et al., 2020), and body size (Baudron et al., 2014; Friedland,
198 Ganley, et al., 2023; Ikpewe et al., 2021). While most macroecological studies of marine fish
199 responses to ocean warming have focused on continental shelf ecosystems given data availability
200 and shelf contributions to fisheries (Free et al., 2019; Friedland, Ganley, et al., 2023; M. McLean
201 et al., 2019; Simpson et al., 2011), many of the dominant species inhabiting continental shelves
202 exhibit complex life histories with seasonally mediated migrations to and from coastal habitats,
203 particularly at mid and lower latitudes (Pinsky et al., 2020; Walsh et al., 2015). Despite the
204 ubiquity of migratory behavior in marine fishes, the macroecological patterns of inshore fish
205 community responses to ocean warming in offshore habitats (Oke et al., 2022; Schonfeld et al.,
206 2022) and how these effects may be mediated by fish functional traits remains understudied.

207 In the Northeast U.S. Continental Shelf (NES) region of the Northwest Atlantic Ocean,
208 surface and bottom water temperatures are increasing at rates that far exceed the global average
209 (Thomas et al., 2017). The NES is a transition zone between subtropical and subpolar
210 biogeographic regions (Friedland et al., 2021), characterized by a steep latitudinal temperature

211 gradient and marine fish populations whose distributions track local climate velocities, and
212 oceanographic and physiographic features (Kleisner et al., 2016, 2017; Pinsky et al., 2013). In
213 the southern region commonly referred to as the Mid-Atlantic Bight (MAB, Fig. 1a)—the focus
214 of this study—population centers have shifted poleward along the predominant latitudinal
215 temperature gradient (Kleisner et al., 2016). Fish distributions on the NES have also expanded,
216 leading to increased local species richness and habitat overlap among species, as well as higher
217 productivity and a shift towards smaller body sizes overall (Fisher et al., 2010a; Friedland,
218 Ganley, et al., 2023). While there are some species whose biomasses have declined in recent
219 years, biomass trends across diverse taxa on the NES have been mostly positive and synchronous
220 with increasing water temperatures (Friedland, Tanaka, et al., 2023).

221 Given these trends on the NES, one might expect similar patterns of increased
222 productivity and diversity to emerge across fish populations in MAB estuaries and coastal bays.
223 Indeed, inshore fish communities along the U.S. east coast have become more abundant and
224 diverse in response to increasing sea-surface temperatures (SSTs) in the region, suggesting that
225 there may be a joint response to warming across species in both the inshore and offshore habitats
226 of the MAB (Friedland, Ganley, et al., 2023; Oke et al., 2022). However, the effects of warming
227 may not be consistent across or within species and inshore habitats in part due to interactions
228 between warming and local environmental and trophic dynamics, but also due to the diversity of
229 species' life history characteristics and thermal affinities (Kleisner et al., 2017) across latitudes
230 in the NES.

231 Fish life history characteristics along the latitudinal temperature gradient of the west
232 Atlantic continental shelf are not uniformly distributed (Fisher et al., 2010b). Instead, the
233 maximum lengths of fish species tends to increase with latitude from the Southeast US

234 Continental Shelf, whose northern biogeographic boundary is the southern end of the MAB at
235 Cape Hatteras, NC (Pappalardo et al., 2015, Fig. 1a), to the northern reaches of the NES and the
236 Scotian Shelf (Fisher et al., 2010b). This gradient is important because body size and the life
237 history parameters that covary with body size, e.g., longevity, age-at-maturity, growth, natural
238 mortality, and generation time are mediating characteristics of population responses to
239 environmental variability (Free et al., 2019; Fujiwara et al., 2022; M. McLean et al., 2019; M. J.
240 McLean et al., 2019; Perry et al., 2005). Frequently, populations of smaller-bodied, shorter-lived
241 fishes are shown to have greater capacities for tracking environmental variability, in terms of
242 their abundance responses and population distributions, due to their rapid turnover of biomass
243 following high natural mortalities and short generation times (M. McLean et al., 2019; Perry et
244 al., 2005). Given the increased favorability of water temperatures in the MAB for more southerly
245 distributed, smaller-bodied species, we would expect that these “faster-lived” species would be
246 those whose population dynamics respond positively to warming SSTs.

247 Here, we explored how seasonally- and regionally-variable ocean SST anomalies from
248 1982 to 2023 impacted inshore fish populations across the MAB, and whether these effects were
249 mediated by species’ positioning along a continuum of life history parameters characterized by a
250 composite index of body size, natural mortality, and maximum age (i.e., characterizing the “fast-
251 slow” continuum of life history characteristics). We also evaluated how the species contributing
252 most to relationships between population responses to seasonal SSTs and the fast-slow
253 continuum differed according to their environmental temperatures (an average of typical water
254 temperatures encountered by species across their ranges) (Pauly, 1980; Thorson et al., 2017),
255 under the expectation that faster-lived species whose populations increased with positive SST
256 anomalies would be associated with warmer waters and slower-lived species whose populations

257 declined with SST anomalies with cooler waters. Our analysis integrated long-term bottom trawl
258 surveys across five estuaries and coastal bays in the MAB (North Carolina to Massachusetts),
259 providing important context for understanding how increasing SSTs impact the communities of
260 fish occupying these habitats.

261 **Methods**

262 *Methods overview*

263 The analysis for this work involved four major steps: 1) We developed abundance indices
264 for commonly collected species in each of five estuaries and coastal bays in the MAB (hereafter
265 referred to as “bays”, and each bay-species combination a “population”). These indices were
266 point estimates of inshore relative abundance for each population during the month of
267 September. 2) We used satellite-derived SST measurements to create season-and bay-specific
268 time series of offshore SST anomalies (within a 150 km radius of trawl survey locations on the
269 continental shelf). We then related SST anomaly time series to abundance indices across all
270 populations using mixed-effects models that included population-specific random slopes
271 representing the effects of seasonal (winter, spring, and summer) offshore SST anomalies on
272 September inshore population abundances. 3) We used principal components analysis (PCA) to
273 ordinate species along three trait axes (natural mortality, M , asymptotic biomass, W_{Inf} , and
274 maximum age, T_{Max}). We then extracted PC1 from this analysis, which we refer to as the “fast-
275 slow index”, to relate to the population-specific random slopes from the previous step. 4) Lastly,
276 we identified the species whose contributions drove relationships between population responses
277 to seasonal SST anomalies and the fast-slow index and tested the hypothesis that the
278 environmental temperatures of those species differed according to whether their populations

279 were fast-lived and responding positively to warming, or slow-lived and responding negatively
280 to warming.

281 *Trawl survey data sources*

282 Data were drawn from September fishery-independent bottom trawl surveys from
283 lagoons and estuaries across the MAB. These bays included Buzzards Bay, Massachusetts (MA),
284 Delaware Bay, New Jersey (NJ), a collection of connected bays in Maryland (MD; Sinepuxent,
285 Assawoman, Chincoteague, and Isle of Wight Bays), South Bay, Virginia (VA), and Pamlico
286 Sound, North Carolina (NC) (Fig. 1).

287 Massachusetts trawls were collected by the Massachusetts Department of Marine
288 Fisheries between 1978 and 2019. These surveys used a 11.9 m otter trawl, were 20 minutes in
289 duration, and typically targeted 4–7 sites per year in Buzzards Bay (mean annual number of
290 Buzzards Bay trawls = 5.9). New Jersey trawls were collected by the New Jersey Bureau of
291 Marine Fisheries as part of the Delaware Bay Juvenile Finfish Trawl Survey between 1997 and
292 2018. This survey used a 4.9 m otter trawl towed for 10 minutes and targeted 11 sites per year in
293 September (except 6 sites in 2017). In Maryland, trawl surveys were performed using a 4.9 m
294 otter trawl by the Maryland Department of Natural Resources between 1990 and 2018 (19–20
295 trawls per year). Trawls in this survey were 6 minutes in duration and targeted four connected
296 coastal bays that fell along roughly 50 km of coastline.

297 Virginia bottom trawls were collected by the Virginia Institute of Marine Science
298 between 2012 and 2023 using a 4.9 m otter trawl. Trawls in this survey were 2 minutes in
299 duration and surveyed 5–6 sites per year. Lastly, the North Carolina bottom trawl survey was
300 conducted between 1987 and 2019 in Pamlico Sound by the North Carolina Department of
301 Marine Fisheries using double-rigged 9.1 m head rope mongoose trawls. We excluded five of 34

302 survey years in Pamlico Sound where the September survey was disrupted by hurricanes and
303 vessel repairs. The remaining 29 years of data included between 36 and 44 trawls annually, each
304 lasting 20 minutes in duration. Prior to developing abundance indices from these data, catches
305 across surveys were standardized to catch per minute by dividing catch numbers by trawl
306 duration (we refer to this metric as “tCPUE”).

307 *Abundance indices*

308 We chose to analyze September bottom trawls because surveys within all states occurred
309 during the month of September (other months were not sampled in all states). For each survey,
310 we used statistical models to estimate the relative abundances of each collected species across
311 years. These models varied according to the complexity of the trawl survey, but all treated the
312 year of sampling as a fixed-effect variable for which abundances were derived. For surveys
313 where information regarding survey locations were available (i.e., “stations” in Buzzards Bay,
314 Delaware Bay, Maryland Coastal Bays, and Pamlico Sound), we developed a model selection
315 routine that fitted six models of varying observation error distribution and random effects
316 structure for each species. For the Maryland Coastal Bays, we pooled catches together prior to
317 fitting these models and did not model catches from the four bays separately because they were
318 adjacent to one another within a relatively small area (roughly 50 km distance).

319 Three of these models were generalized linear mixed models (GLMMs) with either a
320 Tweedie, negative binomial, or Poisson observation error distribution and a random intercept
321 term for “station”. The other three models were generalized linear models (GLMs)
322 corresponding to Tweedie, negative binomial, and Poisson observation error distributions and no
323 random intercept term for station. For each model of the six that converged, we evaluated
324 residual patterns using the DHARMA R package (Hartig, 2020), excluding models from further

325 analysis that displayed significant ($P < 0.05$) deviations from expected residual behaviors
326 (specifically, regarding residual uniformity, dispersion, and zero-inflation). We then used AIC to
327 select the best-fitting model that passed these thresholds for residual behavior. Lastly, we used
328 the selected model to predict September abundance for each year of the trawl survey (excluding
329 random intercepts from this prediction to generate a single annual index for each population). In
330 the case of the South Bay survey that had the fewest years of data and lacked “stations” for
331 which to model using random intercepts, we repeated the above analytical workflow but only for
332 the three GLMs of varying observation error distribution. GLMMs/GLMs were fitted using
333 maximum likelihood with the *glmmTMB* R package (Brooks et al., 2017a). This process yielded
334 77 time series of abundance corresponding to 45 different species across the five bays.

335 There are several species of fish found in the inshore habitats of the MAB that are
336 considered inshore residents and are not found in high abundances offshore. To ensure that our
337 analysis relating offshore SSTs to inshore population abundances was biologically relevant, we
338 cross-referenced the list of species that we developed abundance indices for with a list of species
339 collected offshore in the MAB by the Northeast Fisheries Science Center Bottom Trawl Survey
340 (NEFSC BTS) (Wigley et al., 2003). If species were infrequently collected in the NEFSC BTS,
341 meaning that their abundance or biomass fell outside the 75th percentile of abundances or
342 biomasses across species collected by the survey between 1963-2023, then we excluded them
343 from further analysis.

344 *The fast-slow continuum and environmental temperatures*

345 We generated a composite index of life history parameters to serve as our measure of
346 individual species’ positioning along the fast-slow continuum. The life history index was the first
347 principal component (PC1) extracted from a principal component analysis (PCA) of species-

348 specific life history parameters including natural mortality (M ; the fraction of the population lost
349 to natural causes annually), asymptotic mass (W_{Inf}), and maximum age (T_{Max}). Life history
350 parameters were derived from the R package *FishLife* (Thorson et al., 2017). *FishLife* leverages
351 the public database *FishBase* when life history characteristics are provided in published lab or
352 field studies (Froese et al., 2010). When life history parameters are unknown for a particular
353 species, *FishLife* provides predicted values based on empirically derived parameters from other
354 species belonging to the same taxonomic grouping (e.g., genus, family, order) (Thorson et al.,
355 2017). We found that PC1 explained 89.9% of the total variance across species, and how species
356 fell along PC1 therefore served as our measure of the fast-slow continuum (Fig. 2). There were
357 several different life history parameters that we could have used in creating this index (see
358 Thorson et al., 2017), but they were highly correlated with those chosen here and did not greatly
359 alter the proportion of total variance explained by PC1 or the interpretation of the index.
360 Switching between correlated variables, such as exchanging the von Bertalanffy growth
361 parameter (K) with M and age-at-maturity with maximum age, did not alter our conclusions from
362 later analyses involving this index, nor did only exchanging age-at-maturity with maximum age
363 in the PCA. However, exchanging K for M while leaving maximum age in the PCA increased
364 uncertainty in our final conclusion, indicating this analysis was sensitive to the inclusion of K vs.
365 M in the creation of the fast-slow index.

366 Among the life history parameters available for each species in *FishLife* is a measure of
367 environmental temperature, which is an average of water temperatures used within population
368 growth studies listed on *FishBase*, reflecting mean water temperatures within the regions where a
369 particular population was studied. These values are derived from maps of global water
370 temperatures (Pauly, 1980), and represent environmental temperatures encountered by fishes in

371 the habitats where they are commonly found. When environmental temperatures are unavailable
372 on *FishBase*, *FishLife* provides predictions as for other life history characteristics.

373 *Ocean temperature data and anomaly calculation*

374 We related population abundance indices to seasonal SST anomalies on the MAB
375 continental shelf across the study period. To calculate seasonal SST anomalies, we first drew a
376 150 km radius circle around each survey location that we then intersected with a polygon of the
377 NES whose eastern boundary was defined by the 100 m isobath (this radius was arbitrary, but re-
378 analyses using radii of 100 km and 200 km did not alter conclusions regarding inshore fish
379 population responses to seasonal SSTs). Using these polygons, we cropped rasters of daily SSTs
380 between 1982 and 2023 (using the NOAA Optimum Interpolation SST V2 data set [Reynolds et
381 al., 2007]) to specific regions “local” to inshore habitats (Fig. 1a). We then split this set of daily
382 raster layers into seasons, where we defined winter as January–March, spring as April–June, and
383 summer as July–September. We calculated baseline SSTs for each region as the average of all
384 raster cells and days within that season between 1982 and 2010. We then calculated the seasonal
385 anomaly each year by subtracting this seasonal baseline from the average of all raster cells and
386 days within that buffer area, year, and season (Fig. 1b). We then used trend analysis to evaluate
387 how SST anomalies within the vicinities of surveyed habitats on the continental shelf have
388 changed over the study period. We tested for linear trends in the SST time series using linear
389 regression or generalized least squares (GLS) with various error structures if model residuals
390 were correlated (Table 1).

391 *Species responses to seasonal SST anomalies*

392 We fitted three GLMMs to model September population responses to seasonal SST
393 anomalies (the SSTs in the preceding winter, spring, and summer; summer included SST

394 anomalies from the month of September). For each season, we fitted the following Tweedie
395 GLMM:

$$\begin{aligned} \mu_{ijk} &= \exp(\text{bay}_i + \text{species: bay}_{0ij} + \\ &\quad \text{species: bay}_{1ij} \cdot \text{SSTA}_{ik}), \\ \text{species: bay}_{0ij} &\sim N(0, \sigma_{\text{species: bay}_0}^2), \\ \text{species: bay}_{1ij} &\sim N(0, \sigma_{\text{species: bay}_1}^2), \end{aligned} \quad (1)$$

397 where μ_{ijk} is the abundance of species j in bay i and year k , bay_i is a bay-specific fixed effect
398 intercept, $\text{species: bay}_{0ij}$ is the population-specific random intercept, $\text{species: bay}_{1ij}$ is the
399 population-specific random effect of seasonal SST anomalies, SSTA_{ik} (random slope). Measures
400 of summer SST anomalies included September data, and so the summer model partly reflected
401 species responses to September SST conditions when sampling occurred. These models were
402 fitted using restricted maximum likelihood (REML) (Brooks et al., 2017). After modeling
403 population responses to seasonal SST anomalies, we extracted population-specific random slopes
404 ($\text{species: bay}_{1ij}$) to relate to the fast-slow continuum index described above using a Gaussian
405 GLMM with random intercepts on bay and species. If the inclusion of these intercepts prevented
406 the model from converging, they were removed.

407 Lastly, for each of these latter models relating population-specific SST anomaly effects to
408 the fast-slow continuum, we identified the species driving these patterns by isolating the
409 populations whose random slopes exceeded the 75th percentile of slopes and whose value for
410 PC1 (the fast-slow index) was less than the 25th percentile of PC1. In terms of life history
411 characteristics and direction of responses to SST anomalies, these species represented the
412 “fast/positive” group. Alternatively, we selected the populations whose random slopes were
413 ≤ 25 th percentile of slopes and those whose value for PC1 was ≥ 75 th percentile of PC1,
414 representing the “slow/negative” group. We tested the hypothesis that the environmental

415 temperatures of species comprising these groups were different using a *t* test. Our results for this
416 analysis were robust to the choice of percentiles used to classify species groupings. Selecting the
417 median or 20th/80th percentiles rather than the 25th/75th percentiles as cut-off values for slope
418 and PC1 did not change our conclusions. Selecting the 10th/90th percentiles altered our
419 conclusions, but greatly reduced the number of species within each group (to between 2-4), so
420 we suggest that the 25th/75th percentile cut-offs are reasonable (providing 6-7 species in both
421 “fast/positive” and “slow/negative” groups.)

422 **Results**

423 *Seasonal warming trends in the offshore vicinities of inshore habitats*

424 Our analysis of seasonal offshore SST anomaly trends within the offshore vicinities of
425 surveyed inshore habitats revealed that SST anomaly trend magnitude was seasonally and
426 spatially variable (Fig. 1b, Table 1). In the winter months, SST anomalies were increasing only
427 in the offshore vicinity of Buzzards Bay (linear trend = $0.02\text{ }^{\circ}\text{C yr}^{-1}$, $P = 0.031$). During the
428 spring, the magnitude of warming trends followed a latitudinal gradient (from North to South):
429 near Buzzards Bay (trend = $0.034\text{ }^{\circ}\text{C yr}^{-1}$, $P = 0.06$), Delaware Bay (trend = $0.028\text{ }^{\circ}\text{C yr}^{-1}$, $P =$
430 0.004), and the Maryland Coastal Bays (trend = $0.023\text{ }^{\circ}\text{C yr}^{-1}$, $P = 0.023$), but spring trends were
431 not evident near South Bay or Pamlico Sound. Warming trends were most spatially consistent
432 during the summer months with positive trends occurring across all regions, and trend
433 magnitudes declined with latitude (from North to South): near Buzzards Bay (trend = $0.059\text{ }^{\circ}\text{C}$
434 yr^{-1} , $P < 0.001$), Delaware Bay (trend = $0.042\text{ }^{\circ}\text{C yr}^{-1}$, $P < 0.001$), the Maryland Coastal Bays
435 (trend = $0.04\text{ }^{\circ}\text{C yr}^{-1}$, $P < 0.001$), South Bay (trend = $0.034\text{ }^{\circ}\text{C yr}^{-1}$, $P < 0.001$), Pamlico Sound
436 (trend = $0.034\text{ }^{\circ}\text{C yr}^{-1}$, $P < 0.001$). We include a discussion of these trends to show that inshore
437 fish population responses to SST anomalies reflected population responses to increasing water

438 temperatures, particularly in the summer months when significant warming occurred offshore
439 within the vicinities of all bays, rather than solely interannual variability in water temperatures
440 around a long-term mean.

441 *Inshore fish population responses to seasonal SST anomalies*

442 Our analyses relating the inshore abundances of fishes in the MAB to seasonal SST
443 anomalies offshore revealed that population responses (meaning positive or negative abundance
444 responses within species and bays) were weakly mediated by species' positioning along the fast-
445 slow continuum (Fig. 3, Table 2, $P_{\text{Spring}} = 0.08$, $P_{\text{Summer}} = 0.087$). These relationships were
446 uncertain due to population-level variability in species responses to seasonal SST anomalies—
447 for species that were found in least three of five bays over the study period (whose abundances
448 we modeled, in total, nine species), the range of population responses to SST anomalies included
449 zero for 56% of species ($n = 5$) in the spring SST anomaly model and for 89% of species ($n = 8$)
450 in the summer SST anomaly model (Fig. 4). This suggested that the mediating effects of life
451 history on how populations responded to seasonal SST anomalies were highly variable across
452 bays.

453 Lastly, we identified the faster-lived species responding positively to seasonal SST
454 anomalies (the “fast/positive” group) and the slower-lived species responding negatively to
455 seasonal SST anomalies (the “slow/negative” group). In comparing the environmental
456 temperatures of these species, we found that the fast/positive group had significantly higher
457 environmental temperatures ($P_{\text{Spring}} = 0.008$, $P_{\text{Summer}} = 0.051$) than the slow/negative group (Fig.
458 5). While the life history effect mediating inshore population responses to seasonal SST
459 anomalies was weak, the fastest-lived species responding positively and slowest-lived species
460 responding negatively to SST anomalies were organized along an environmental temperature

461 gradient. This finding aligns with the expectation that populations of warm-adapted and faster-
462 lived species benefit more from a warming MAB than cool-adapted and slower-lived species, as
463 the spatial gradient of warming trends in the region shows that SSTs in cooler, northern regions
464 are increasing more rapidly than in warmer, southern regions.

465 **Discussion**

466 Our study shows that SST anomalies increased across the offshore habitats of the MAB
467 (1982–2023), and that this changing environment tended to promote the autumn population
468 abundances of faster-lived species, while the abundances of slower-lived species tended to
469 decline in response to increasing SSTs in the spring and summer. However, these relationships
470 were weak ($P_{\text{Spring}} = 0.08$, $P_{\text{Summer}} = 0.087$), indicating that population-level responses to warming
471 differed in sign (positive/negative) between the inshore habitats of the MAB within species. This
472 bay-to-bay variability in population responses to warming could have been due to differential
473 population responses to region-specific environmental variability or species interactions that
474 went uncharacterized in this study. In addition, we found that the fastest-lived species whose
475 abundances were positively associated response with warming were warm-adapted, whereas the
476 slowest-lived species whose abundances were negatively associated with warming were cool-
477 adapted. Together, our work shows that the autumn population dynamics of inshore fishes in the
478 MAB are being altered by warming waters, and that the magnitude of population responses to
479 warming were weakly dependent on species' positioning along the fast-slow life history
480 continuum, thermal preferences, and local scale processes mediating the integration of these
481 characteristics into population dynamics.

482 Life history-mediated changes in the population dynamics and distributions of fishes in
483 response to warming are well-documented in continental shelf habitats (Free et al., 2019; M.

484 McLean et al., 2019; Perry et al., 2005). In the North Sea, those species whose populations
485 shifted or became more abundant in response to warming waters were found to be shorter-lived
486 and smaller-bodied than those populations that did not, suggesting that the faster turnover of
487 these species served as an adaptive response to a warming environment (Perry et al., 2005;
488 Simpson et al., 2011). However, the evidence provided from these studies that fast-lived species
489 shifted their distributions (Perry et al., 2005) and increased in abundances (Simpson et al., 2011)
490 in response to warming to a greater degree than slow-lived species identified these patterns by
491 comparing life history parameters (e.g., maximum body size) between groups of species that
492 shifted or were more abundant following warming. In the latter study of abundance changes, the
493 authors found no relationship between the magnitude of abundance changes and maximum body
494 sizes (Simpson et al., 2011), and this hypothesis is most like the one tested in our study—that the
495 magnitudes of abundance responses to warming fell along a gradient of fast-slow life histories.

496 The differential within-species responses to warming in our study contributed to the
497 uncertainty around the relationship between magnitudes of warming responses and life history
498 characteristics, but this response variability could shed light on why the potentially mediating
499 effects of life history characteristics on the magnitude of abundance changes in response to
500 warming is elusive on shelf-wide spatial scales. That is, region-specific dynamics may confound
501 patterns of warming responses mediated by life history characteristics evaluated at larger spatial
502 scales. Of course, fish population distributions that are not constrained by the physiographic
503 features of inshore environments will reasonably be expected to respond more homogeneously to
504 warming. However, a more explicit consideration of population structures and the spatial
505 variability of population responses to warming in continental shelf habitats, where available data
506 sources are more highly resolved spatially than in inshore environments, may yield insights into

507 how abundance responses to warming fall along life history gradients like the fast-slow gradient
508 evaluated in this study.

509 While our analysis did not touch on population distribution shifts, other work in the
510 adjacent NES has shown that suitable habitat area has expanded for many commonly collected
511 species (Friedland, Langan, et al., 2020; Friedland, Tanaka, et al., 2023) and that marine fish
512 population centers in the MAB have generally shifted poleward (Kleisner et al., 2016). Our
513 findings that the species contributing most positively to the relationship between population
514 responses to warming and life history characteristics were warm-adapted while those
515 contributing most negatively were cool adapted supports an analogous interpretation to those
516 findings from the NES—an increase in suitable thermal habitat for warm-adapted species in the
517 MAB may have promoted increased abundances of those species in inshore habitats, whereas a
518 decline in suitable thermal habitat for cool-adapted species promoted population declines for
519 those species. However, this interpretation does not speak to the mechanisms through which
520 changes in autumn population abundances in inshore habitats occurs in response to warming.
521 While discerning the specific mechanisms through which warming impacts coastal fish
522 populations falls outside the scope of this study, there are several avenues through which inshore
523 population dynamics may hypothetically respond to warming on the continental shelf.

524 For instance, an important mechanism that is sensitive to offshore warming and could
525 explain patterns of increased September abundances of warm-adapted species following
526 warming is the phenology of migratory and spawning behaviors of coastal fishes. Many fishes
527 common to the MAB will leave coastal embayments for the continental shelf to spawn in the
528 summer and/or fall months, e.g., spot *Leiostomus xanthurus* and Atlantic croaker *Micropogonias*
529 *undulatus* (Bucheister et al. 2013; Schonfeld et al. 2022). An explanation for why September

530 population abundances in southerly inshore habitats tend to respond positively to warming could
531 be related to the timing of this migration. That is, a later onset of seasonal cooling during the fall
532 could promote fishes to remain in inshore habitats for longer, therefore increasing September
533 abundances of inshore fishes.

534 Water temperatures also influence the timing of spawning (Norcross & Austin, 1988) and
535 larval ingress (Thaxton et al., 2020) from the continental shelf into inshore habitats. The
536 availability of food resources is critical for the survival of larval fishes (Cushing, 1990), and so
537 changes in spawning and ingress phenology may threaten the spatiotemporal overlap between
538 larvae and their food, ultimately impacting their recruitment and population abundances in
539 inshore habitats. Warming in offshore habitats is also expected to be correlated with warming in
540 inshore habitats, and warming has been documented in several embayments along the NES
541 (Nixon et al., 2009; Rheuban et al., 2016; Rice et al., 2015; Woodland et al., 2021). However,
542 regardless of spawning location, regional warming may enhance fish population abundances up
543 to a point but may also eventually lead to declines in abundances as temperatures increase
544 outside of species' optimal thermal ranges; non-linear relationships that we did not quantify in
545 our work (Auth et al. 2020). This may help to explain the declining patterns of abundances
546 among cool-adapted species found in our study. Future work may better understand the
547 mechanisms behind population responses to SST anomalies by considering these non-linear
548 effects between abundances and temperatures, the age-structure of collected fishes, as well as a
549 broader range of months when trawl surveys occurred to characterize interactions between
550 warming and migratory and spawning phenologies.

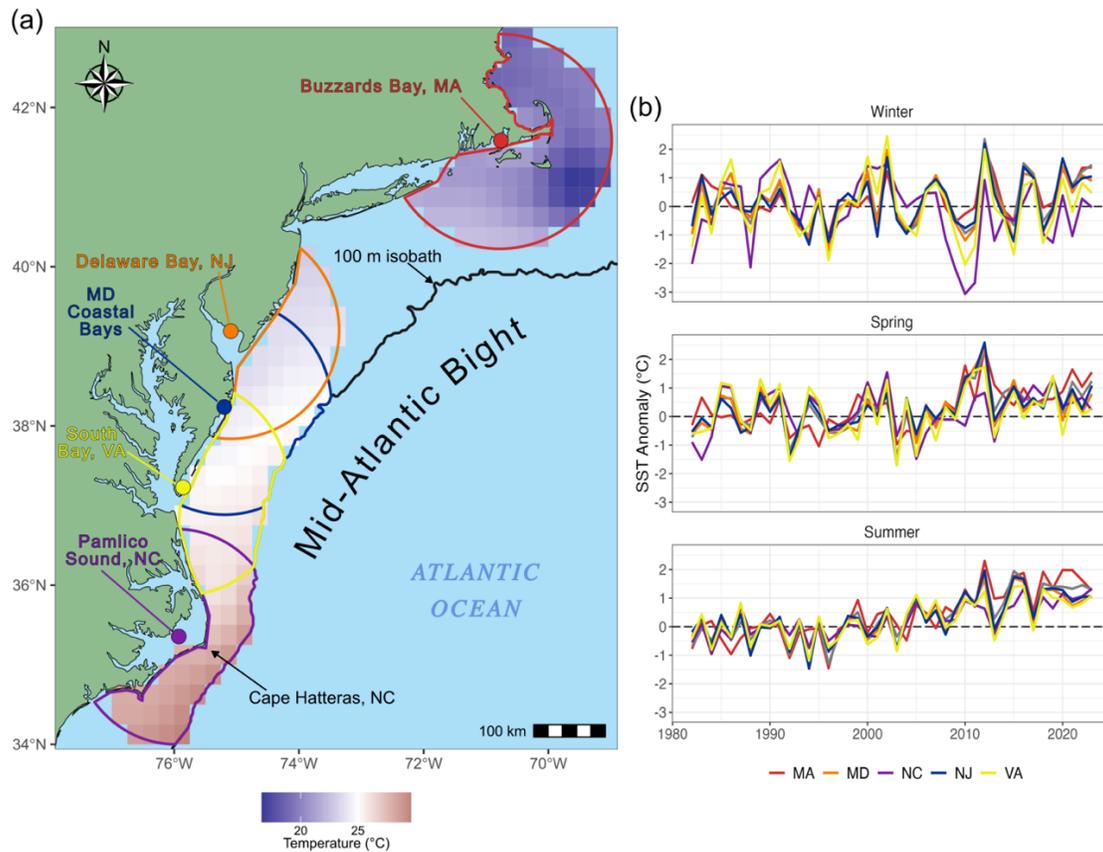
551 In this work, we analyzed five disparate bottom trawl surveys from across the Mid-
552 Atlantic Bight to understand how autumn inshore fish populations responded to SST anomalies

553 in offshore habitats, and how these responses were mediated by species' positioning along the
554 fast-slow life history continuum. Our results showed that there was high bay-to-bay variability in
555 population responses to warming, but also that population responses to warming were mediated
556 by their thermal affinities and weakly by their life history characteristics. September population
557 abundances of faster-lived, smaller-bodied species that are warm-adapted have tended to increase
558 in response to spring and summertime warming, whereas populations of larger-bodied, longer-
559 lived species that are cool-adapted have tended to decline in response. This work advances our
560 understanding of how climate change influences the dynamics of inshore fish populations in the
561 MAB, and provides insights into the species likely to become more abundant in the region under
562 continued warming.

563 **Acknowledgements**

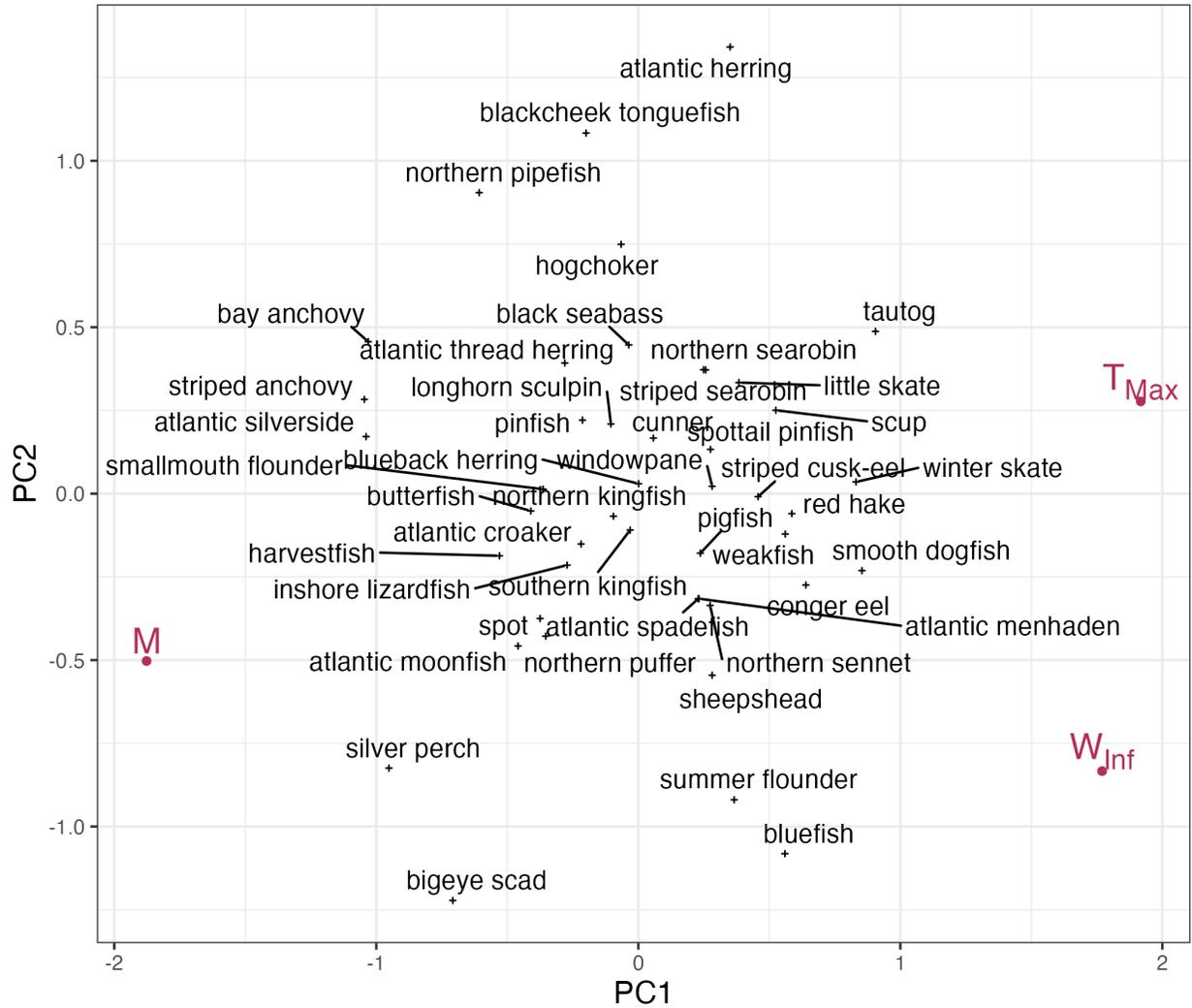
564 This project was made possible by financial support from Virginia Sea Grant through their
565 graduate research fellowship. We are extremely grateful to the research staff and scientists who
566 collected and organized the trawl survey data used in this study.

567 **Figures**



568

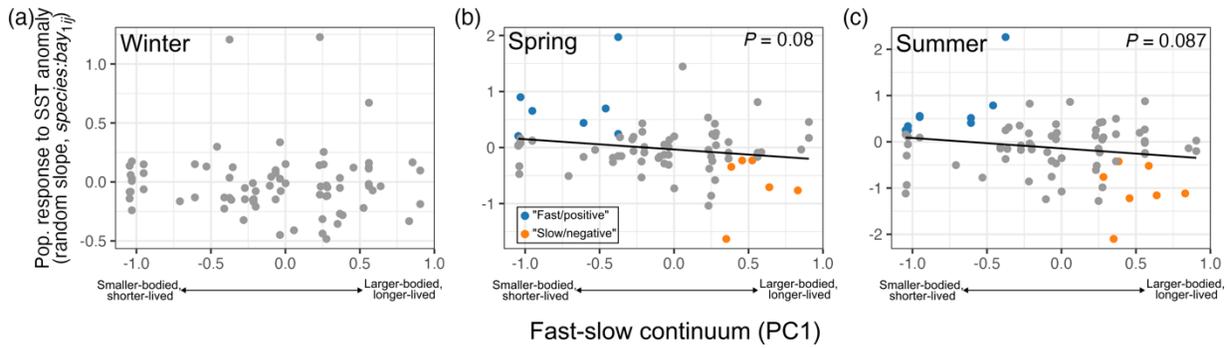
569 *Figure 1.* Map of the study region shown with the centroid locations of bottom trawl surveys and
570 the names of the bays where trawls occurred (a). Abbreviations are for US states: MA =
571 Massachusetts, NJ = New Jersey, MD = Maryland, VA = Virginia, NC = North Carolina. “MD
572 Coastal Bays” refers to the centroid of trawls performed across four connected coastal bays
573 spanning ~50 km of coastline. The colored polygons associated with survey locations are the 150
574 km radius circles on the NES that were intersected with the 100 m isobath boundary. The filled
575 grid cells within those polygons show summer 2023 SSTs (°C) from the NOAA OISST V2 data
576 set. Data corresponding to these grid cells were used to calculate bay-specific seasonal SST
577 anomalies, the time series of which are shown in (b), where line colors correspond to the
578 different regions identified on the map.



579

580 *Figure 2.* PCA biplot showing the multivariate positioning of species considered in our SST
 581 anomaly analysis in relation to their life history parameters: natural mortality (M), maximum age
 582 (T_{Max}), and asymptotic biomass (W_{Inf}).

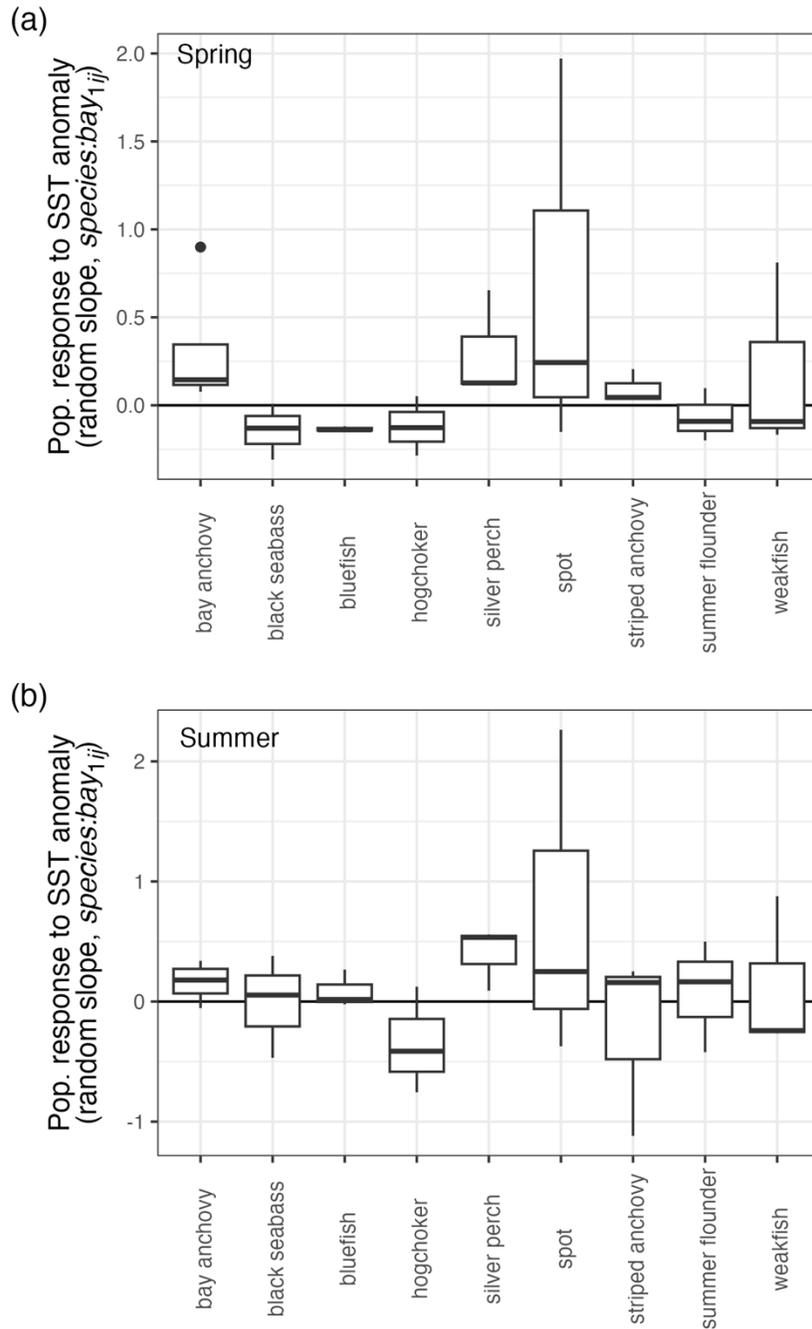
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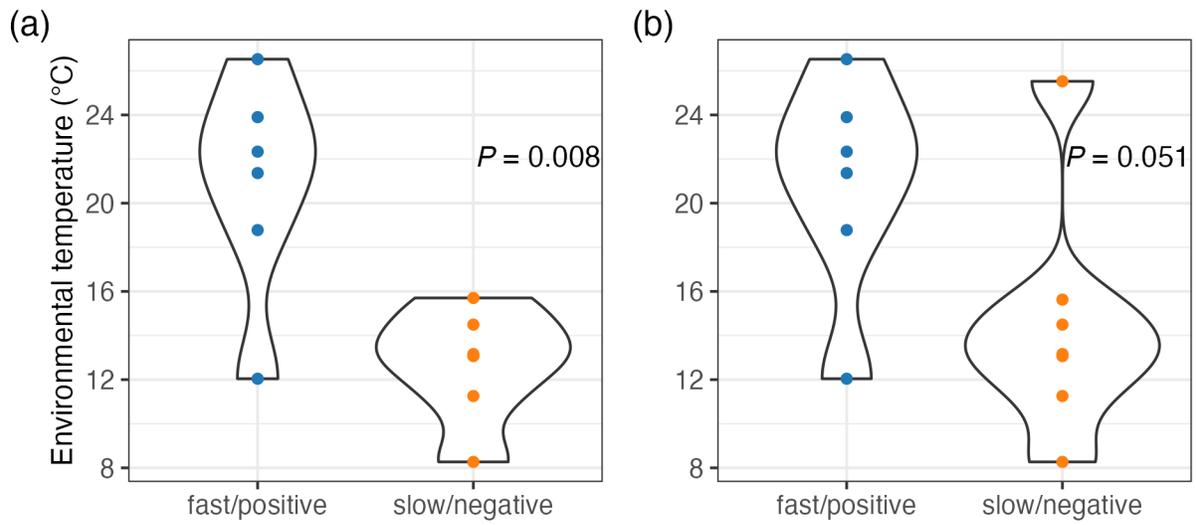
585

586 *Figure 3.* The relationship between autumn inshore fish population responses to seasonal SST
587 anomalies in winter (a), spring (b), and summer (c) and the fast-slow continuum. The blue points
588 represent “fast/positive” populations whose SST responses exceed the 75th percentile for slopes
589 and whose PC1 value is in the lower 25th percentile of PC1. Alternatively, the orange points
590 correspond to “slow/negative” populations whose SST responses fall inside the bottom 25th
591 percentile of slopes and whose PC1 values exceed the 75th percentile for PC1.



592

593 *Figure 4.* Population responses to spring (a) and summer (b) shown by species to visualize their
 594 variability around zero effect (black horizontal line). Only species with population estimates in
 595 three or more bays were included here.



596

597 *Figure 5.* Violin plots show the distribution of environmental temperatures associated with
 598 species driving patterns between population responses to spring (a) and summer (b) SST
 599 anomalies on the continental shelf and the fast-slow continuum (Fig. 3). P values are from t tests
 600 between the “fast/positive” and “slow/negative” groups.

601

602 *Table 1:* Trend statistics for seasonal SST anomalies on the continental shelf by region,
 603 reflecting trends within the boundaries on the continental shelf identified in Figure 1. Trends
 604 with $P < 0.01$ are highlighted. “iid” refers to independent and identically distributed error.

<i>State</i>	<i>Season</i>	<i>Trend ($^{\circ}\text{C yr}^{-1}$)</i>	<i>SE</i>	<i>T statistic</i>	<i>P</i>	<i>Error structure</i>
MA	Winter	0.020	0.009	2.239	0.031	iid
MA	Spring	0.034	0.018	1.930	0.061	AR(2)
MA	Summer	0.059	0.007	8.267	0.000	iid
NJ	Winter	0.015	0.011	1.433	0.160	iid
NJ	Spring	0.028	0.009	3.071	0.004	iid
NJ	Summer	0.042	0.008	5.462	0.000	iid
MD	Winter	0.013	0.011	1.162	0.252	iid
MD	Spring	0.023	0.010	2.363	0.023	iid
MD	Summer	0.040	0.008	5.260	0.000	iid
VA	Winter	0.000	0.014	-0.011	0.991	iid
VA	Spring	0.010	0.009	1.088	0.283	AR(3)
VA	Summer	0.034	0.007	4.563	0.000	iid
NC	Winter	-0.009	0.024	-0.395	0.695	AR(1)
NC	Spring	0.017	0.009	1.853	0.071	iid
NC	Summer	0.034	0.005	6.470	0.000	iid

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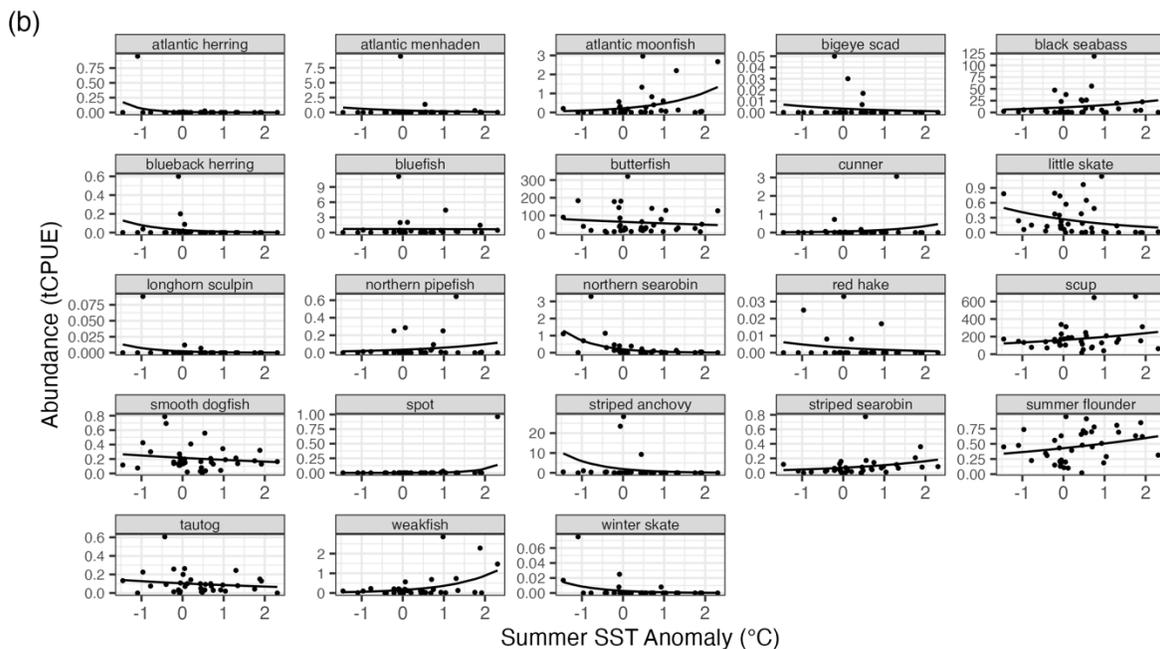
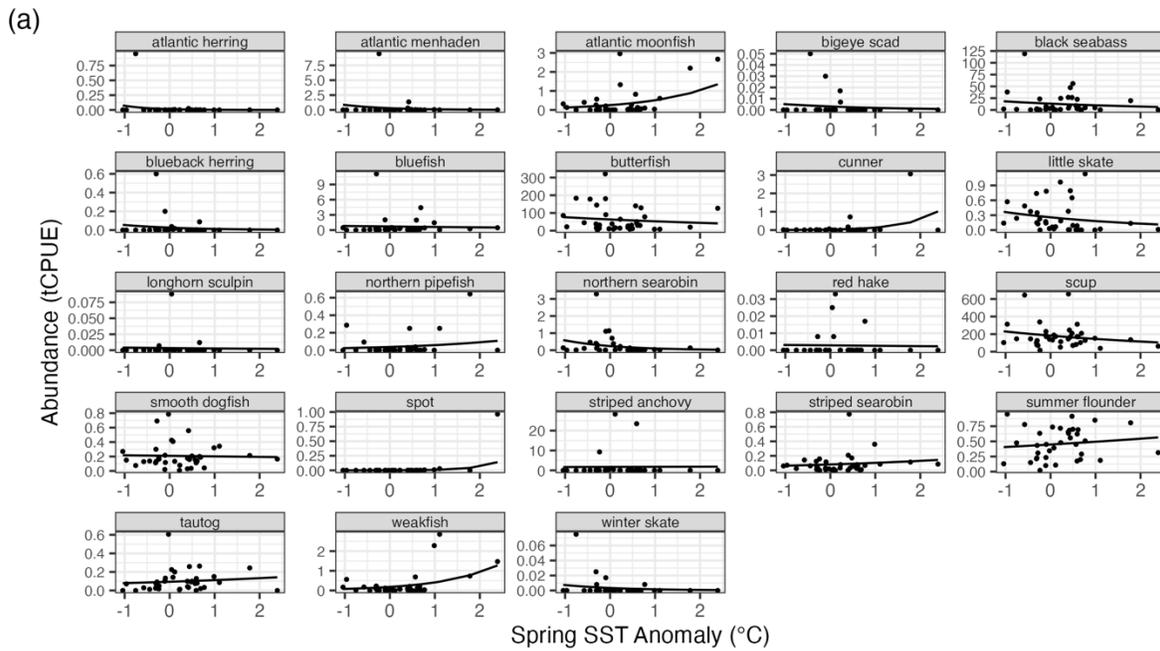
606 *Table 2.* Model summary from mixed effects model relating inshore fish population responses to
607 winter, spring, and summer SST anomalies (within 150 km of the centroid of trawl survey
608 locations on the continental shelf and within the 100 m isobath) to the fast-slow continuum
609 (PC1). Here, $\tau_{00,bay}$ and $\tau_{00,species}$ are random intercepts for bay and species respectively, and N_{bay}
610 indicates the number of bays and $N_{species}$ the number of species. When random intercepts are
611 excluded, their inclusion meant that they prevented the model from converging.

<i>SST anomaly season</i>	<i>Predictors</i>	<i>Estimates</i>	<i>CI</i>	<i>P</i>
Winter	(Intercept)	-0.00	-0.08– 0.05	0.676
	Fast-slow continuum (PC1)	-0.01	-0.13 – 0.10	0.829
Random Effects				
	σ^2	0.08		
	$\tau_{00,bay}$	0.00		
	$\tau_{00,species}$	0.00		
	N_{bay}	5		
	$N_{species}$	45		
	Observations	77		
	Marginal R^2	0.001		
<i>SST anomaly season</i>	<i>Predictors</i>	<i>Estimates</i>	<i>CI</i>	<i>P</i>
Spring	(Intercept)	-0.02	-0.15 – 0.08	0.551
	Fast-slow continuum (PC1)	-0.18	-0.39 – 0.02	0.080
Random Effects				
	σ^2	0.22		
	$\tau_{00, bay}$	0.00		
	$\tau_{00, species}$	0.01		

	N_{bay}	5		
	N_{species}	45		
	Observations	77		
	Marginal R^2	0.045		
<hr/>				
<i>SST anomaly season</i>	<i>Predictors</i>	<i>Estimates</i>	<i>CI</i>	<i>P</i>
<hr/>				
Summer	(Intercept)	-0.14	-0.29 – -0.01	0.002
	Fast-slow continuum (PC1)	-0.22	-0.48 – -0.03	0.087
	Random Effects			
	σ^2	0.38		
	$\tau_{00, \text{species}}$	0.00		
	N_{bay}	5		
	N_{species}	45		
	Observations	77		
	Marginal R^2	0.038		

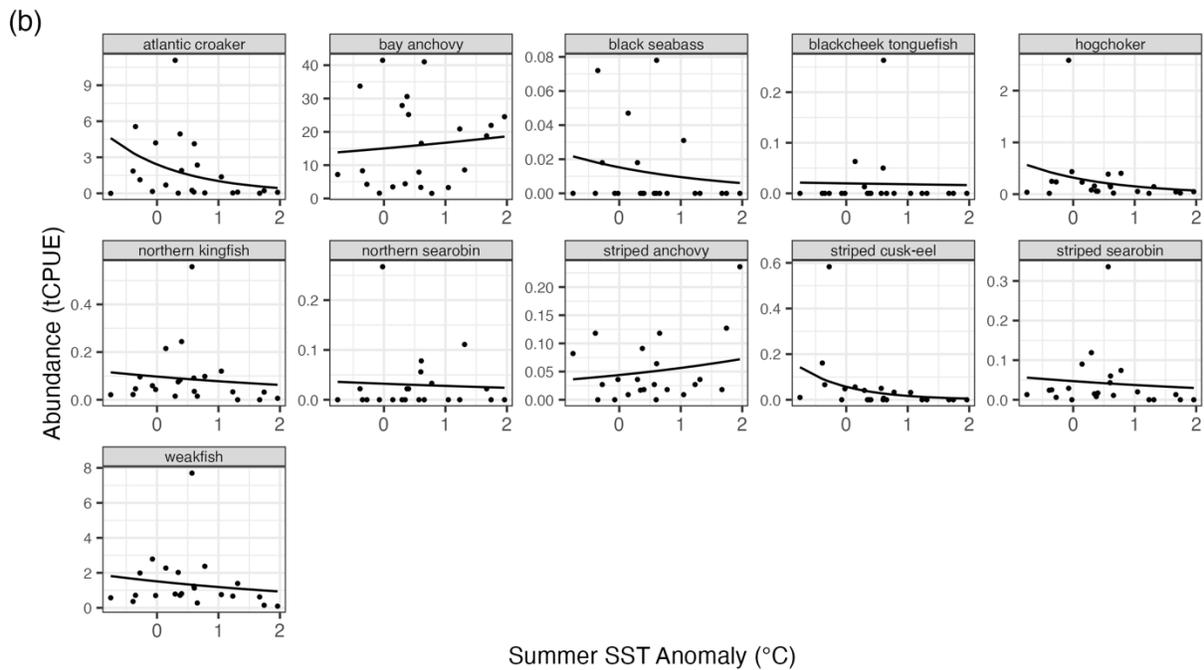
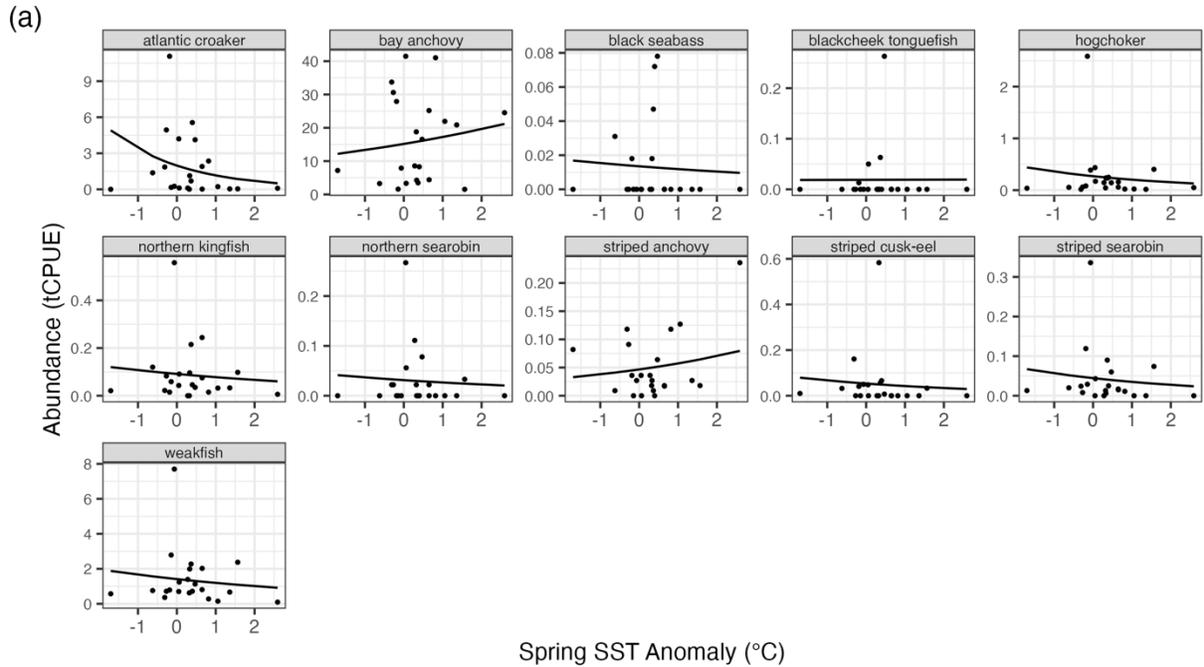
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613 Appendix



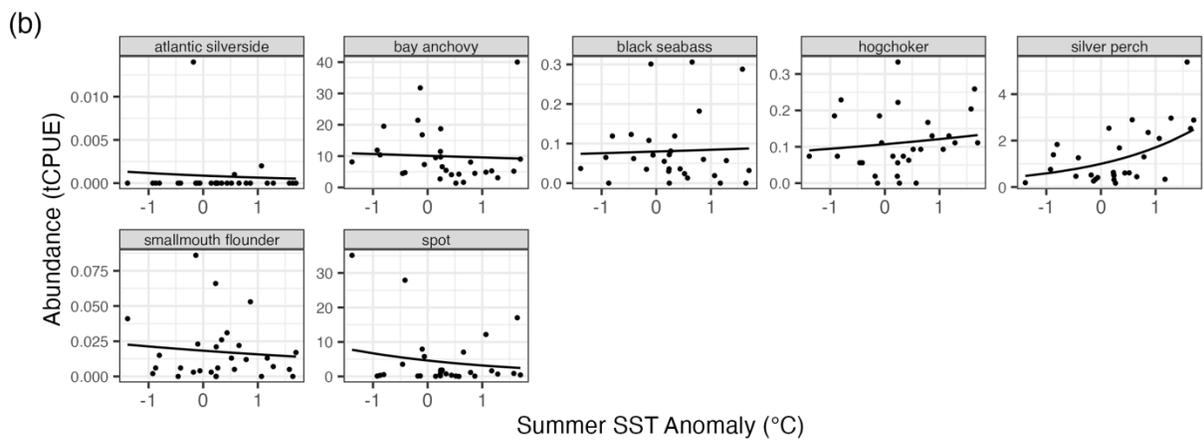
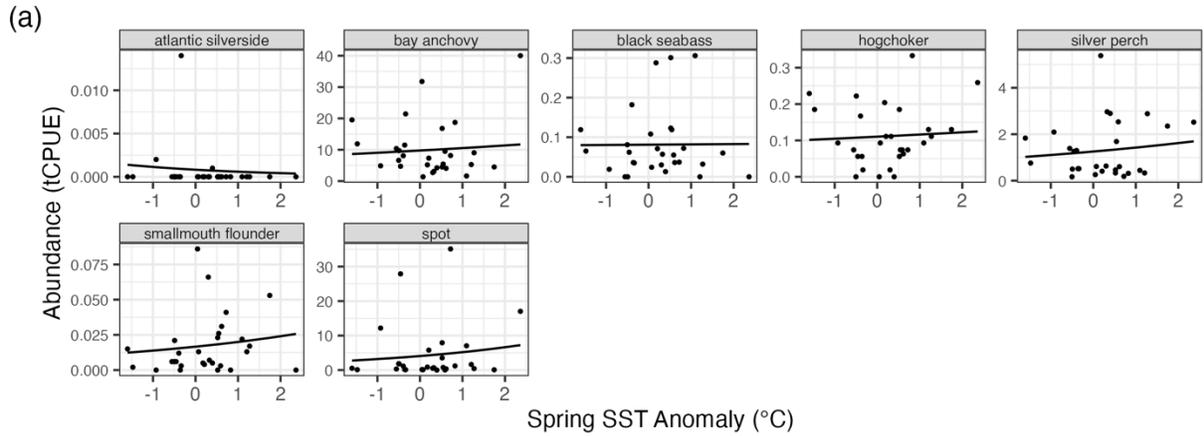
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615 Appendix 1. Species-level predictions for the relationship between autumn inshore tCPUE
 616 (modeled time series of catch per trawling minute) and spring (a) and summer (b) SST anomalies
 617 in the offshore habitat near Buzzards Bay, Massachusetts.



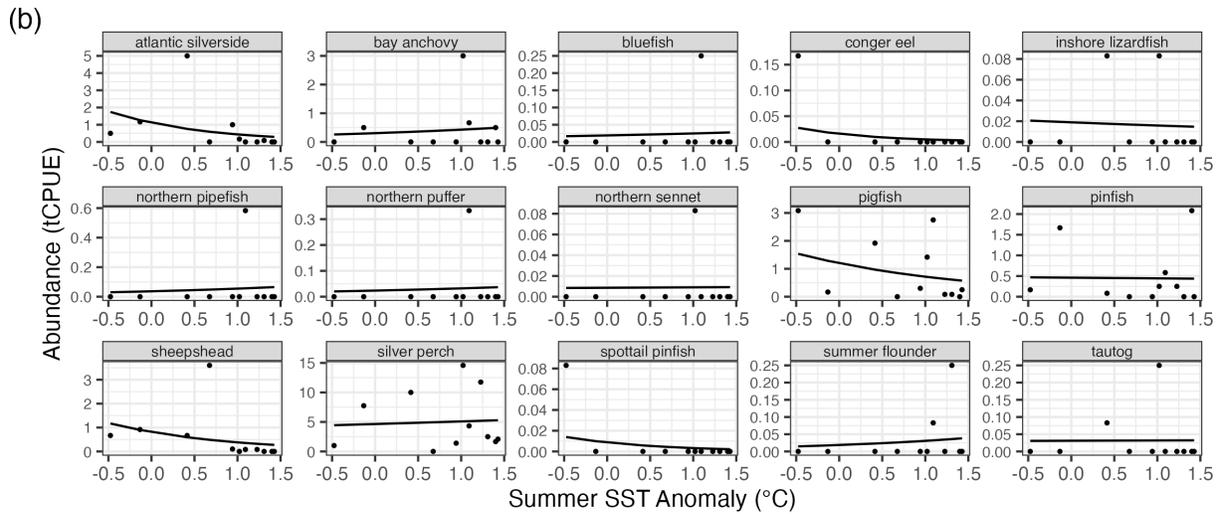
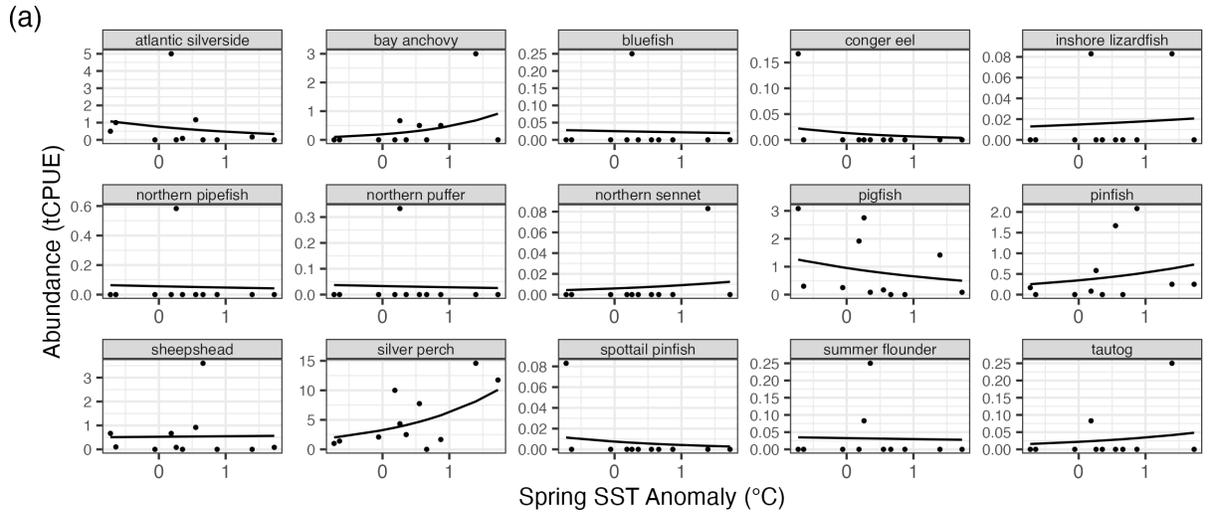
618

619 *Appendix 2.* Species-level predictions for the relationship between autumn inshore tCPUE
 620 (modeled time series of catch per trawling minute) and spring (a) and summer (b) SST anomalies
 621 in the offshore habitat near Delaware Bay, New Jersey.



622

623 *Appendix 3. Species-level predictions for the relationship between autumn inshore tCPUE*
 624 (modeled time series of catch per trawling minute) and spring (a) and summer (b) SST anomalies
 625 in the offshore habitat near the coastal bays of Maryland.

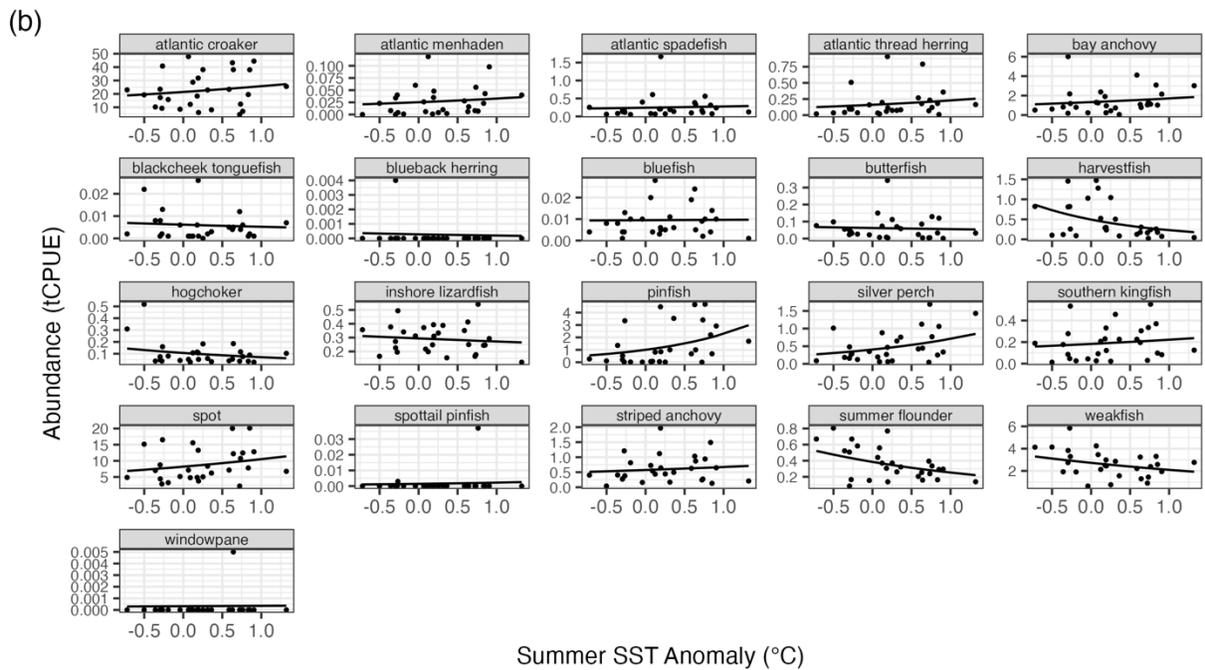
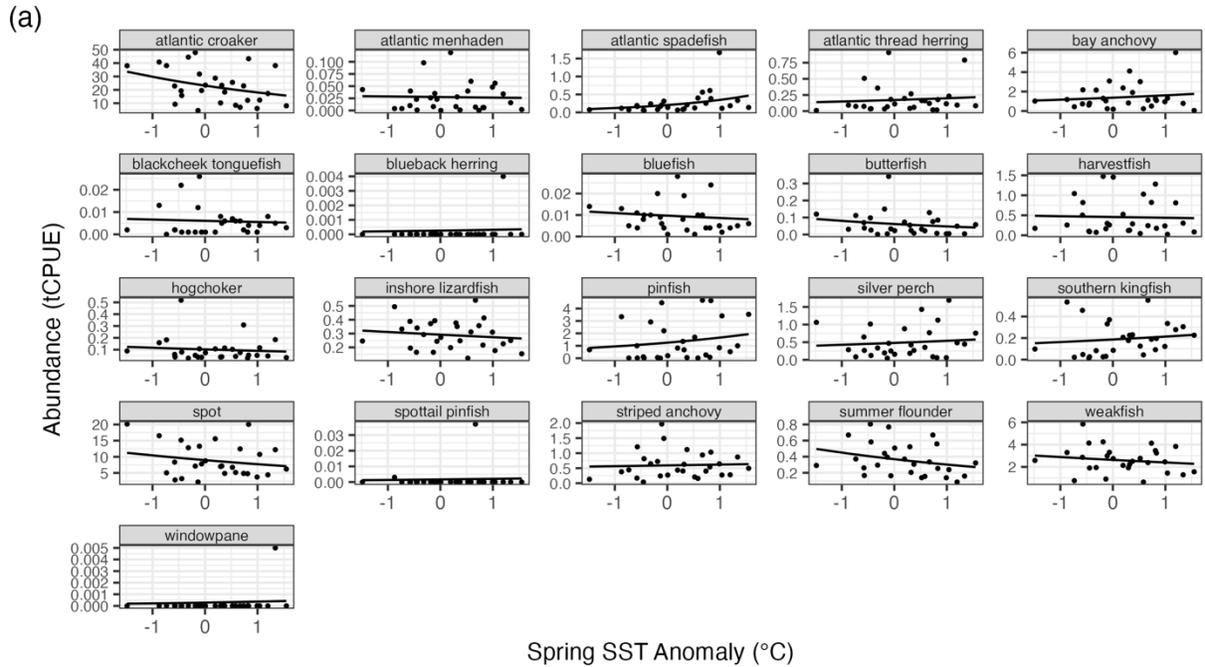


626

627 *Appendix 4. Species-level predictions for the relationship between autumn inshore tCPUE*

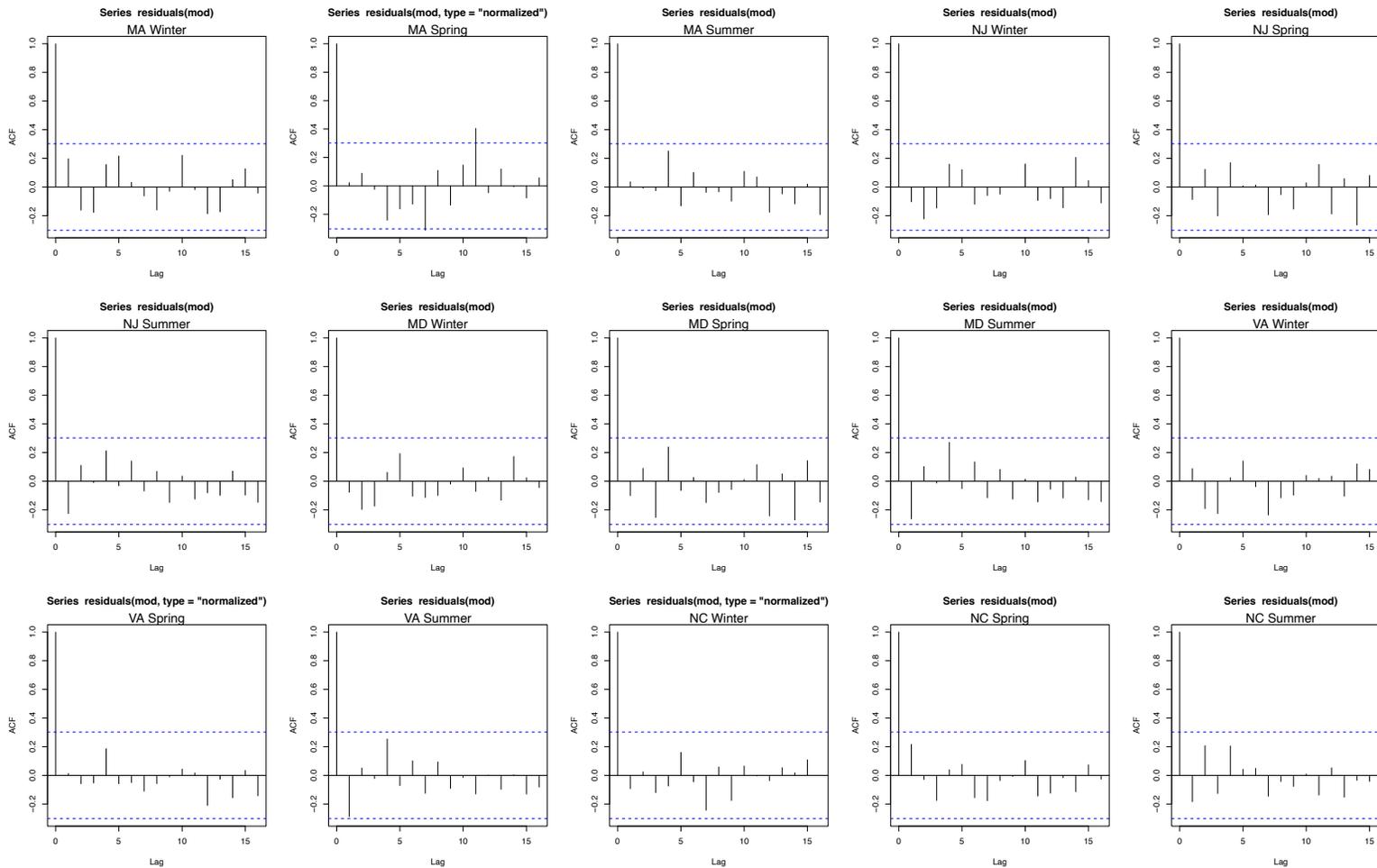
628 (modeled time series of catch per trawling minute) and spring (a) and summer (b) SST anomalies

629 in the offshore habitat near South Bay, Virginia.



630

631 *Appendix 5. Species-level predictions for the relationship between autumn inshore tCPUE*
 632 *(modeled time series of catch per trawling minute) and spring (a) and summer (b) SST anomalies*
 633 *in the offshore habitat near Pamlico Sound, North Carolina.*



634

635 *Appendix 6.* Autocorrelation of residuals from trend models of seasonal SST anomalies. Figures with “type = normalized” show the
 636 autocorrelation of normalized residuals after fitting the trend model using generalized least squares with various error structures (see
 637 Table 1).

638 *Appendix 7. Summary statistics from the three mixed-effects models used to model population*
 639 *responses to warming in the winter, spring, and summer months. “CI” is the 95% confidence*
 640 *interval for the estimates.*

<i>Season</i>	<i>Predictors</i>	<i>Estimates</i>	<i>CI</i>	<i>P</i>
Winter	(Intercept)	0.15	0.05 – 0.48	0.001
	<i>bay</i> [Delaware Bay]	0.94	0.12 – 7.16	0.951
	<i>bay</i> [MD Coastal Bays]	1.2	0.11 – 13.16	0.88
	<i>bay</i> [Pamlico Sound]	0.8	0.15 – 4.27	0.794
	<i>bay</i> [South Bay]	0.59	0.09 – 3.78	0.58

Random Effects

(intercept) <i>species: bay</i> _{0ij}	7.846
(slope) <i>species: bay</i> _{1ij}	0.164
N _{species}	45
N _{bay}	5
Observations	2049

<i>Season</i>	<i>Predictors</i>	<i>Estimates</i>	<i>CI</i>	<i>P</i>
Spring	(Intercept)	0.13	0.04 – 0.43	0.001
	<i>bay</i> [Delaware Bay]	1.14	0.15 – 8.96	0.9
	<i>bay</i> [MD Coastal Bays]	1.36	0.12 – 15.38	0.805
	<i>bay</i> [Pamlico Sound]	0.94	0.17 – 5.17	0.948
	<i>bay</i> [South Bay]	0.59	0.09 – 3.89	0.584

Random Effects

(intercept) <i>species: bay</i> _{0ij}	8.066
(slope) <i>species: bay</i> _{1ij}	0.426
N _{species}	45
N _{bay}	5
Observations	2049

<i>Season</i>	<i>Predictors</i>	<i>Estimates</i>	<i>CI</i>	<i>P</i>
Summer	(Intercept)	0.12	0.04 – 0.40	0.001
	<i>bay</i> [Delaware Bay]	1.31	0.16 – 10.57	0.801
	<i>bay</i> [MD Coastal Bays]	1.47	0.13 – 17.09	0.76
	<i>bay</i> [Pamlico Sound]	0.99	0.18 – 5.55	0.995
	<i>bay</i> [South Bay]	0.71	0.10 – 4.88	0.731

Random Effects

(intercept) <i>species:bay_{0ij}</i>	8.263
(slope) <i>species:bay_{1ij}</i>	0.681
N _{species}	45
N _{bay}	5
Observations	2049

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646 Chapter 2

647 *Asynchrony drives stability across coupled social-ecological systems: an example from*

648 *Chesapeake Bay demersal fisheries*

649

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664

665 **Abstract**

666 Ecological diversity can confer temporal stability of ecosystem processes through
667 asynchrony in species' abundances. This stability can also translate to stability of ecosystem
668 services, for example in fisheries, by distributing harvest pressure among asynchronous
669 populations that are also exploited. Coupled investigations that consider both naturally- and
670 anthropogenically-induced asynchrony are rare, but have the potential to greatly inform the
671 management of resource systems. Here, we used harvest data and fisheries-independent trawl
672 surveys in two regions (Maryland and Virginia) of the Chesapeake Bay, USA from 2002–2018
673 to explore how seasonal (within-year) asynchrony among targeted fish species contributed to: 1)
674 seasonal asynchrony in the harvests of these species, and 2) the within-year stability and
675 economic value of harvests. In Maryland, seasonal closure of striped bass (*Morone saxatilis*)
676 fishing resulted in asynchrony by forcing the harvest of alternative stocks. In Virginia, the
677 natural emigration of target species from the system was positively associated with harvest
678 asynchrony, although this effect was driven largely by the increase in relative evenness among
679 remaining species that constituted a much lower exploitable biomass. Our findings show that
680 both social (management) and ecological factors influence asynchrony in the harvests of fished
681 populations, which goes onto affect the stability of regional harvest portfolios.

682

683 **Introduction**

684 The management of social-ecological systems like commercial fisheries is challenged by
685 their immense complexity that limits the comprehension of the drivers of system dynamics
686 (Levin et al., 2013; Link, 2018) and ultimately their predictability (Boettiger & Hastings, 2013).
687 However, in commercial fishing harvest portfolios, defined as dynamic collections of harvests
688 drawn from multiple species and/or populations within a region (i.e., “stocks”), as well as in the
689 ecosystems those portfolios depend upon, it is interestingly the fostering and conservation of
690 complexity that has long been advocated for as a strategy to reduce variability in aggregate
691 system properties like total harvest yield and revenue (Cline et al., 2017; Hilborn et al., 2001;
692 Schindler et al., 2010) or community biomass production (Tilman, 1996; Tilman et al., 1998).
693 When commercial fishing portfolios or biological communities are composed of diverse species
694 assemblages, the asynchronous dynamics of these “assets” leads to reduced variability (enhanced
695 stability) over time in the aggregate system relative to the average asset, otherwise known as a
696 “portfolio effect” (S. C. Anderson et al., 2017; Link, 2018; Schindler et al., 2015). Despite the
697 dependence of commercial fisheries on the population dynamics of the species they target, the
698 relationships between asynchrony among species populations in the environment (species
699 asynchrony) and the asynchrony of harvests drawn from those species (harvest asynchrony) are
700 not well integrated. Clarifying these relationships will shed light on the stabilizing role of
701 biodiversity in social-ecological systems like commercial fisheries.

702 The stabilizing effect of asynchrony is a common feature of complex, hierarchical, and
703 dynamic systems that are composed of species populations whose abundances fluctuate through
704 time (Link, 2018). Understanding how asynchrony in harvest portfolios and the biological
705 communities they depend upon relate to each other benefits from first considering how

706 asynchrony emerges from within these systems. While there are many ways to quantify
707 asynchrony (Doak et al., 1998; J. W. Moore et al., 2021; Siple et al., 2020; Thorson et al., 2018),
708 we adopt an approach developed in community ecology where, in the context of a community of
709 species, asynchrony (ϕ) is quantified as the ratio comparing community stability ($S_{Community}$) to
710 an average across-species stability ($S_{Species}$), such that $S_{Community} = \phi S_{Species}$ (Thibaut &
711 Connolly, 2013; Wang & Loreau, 2016; Zhao et al., 2022). While this approach has only
712 previously been applied in an ecosystem context, we extend its application to fishery harvest
713 portfolios that share similar hierarchical structures to communities, i.e., populations (stocks) that
714 comprise a portfolio (harvest) (Link, 2018). In all cases, stability is the inverse temporal
715 variability (i.e., the inverse coefficient of variation = $1/CV$) of system properties. In this
716 formulation, ϕ cannot be < 1 , meaning that stability of the community will almost always be
717 higher than average across-species stability due to asynchrony (Doak et al., 1998; Link, 2018).
718 This measure of asynchrony integrates two related processes whose effects can be partitioned
719 mathematically: statistical averaging and compensation (Fig. 1, Zhao et al., 2022).

720 Statistical averaging is the enhancement of system stability that occurs when system
721 components fluctuate independently of one another through time, thereby reducing variability of
722 the system in aggregate (Doak et al., 1998; Link, 2018). In the mathematical partitioning of ϕ ,
723 Zhao et al. (2022) showed that statistical averaging depends on the quantity and evenness of
724 species variances. For example, the contribution of statistical averaging to ϕ will be low in a
725 community characterized by few species whose population dynamics are independent and with
726 highly unequal variances. This unevenness in variances is a common feature of communities,
727 which are almost universally skewed in abundances among species (Avolio et al., 2019;
728 Pedersen et al., 2019) and where population variances tend to increase non-linearly with their

729 means (S. C. Anderson et al., 2013; Mikkelsen et al., 2011). In contrast, compensation effects
730 also contribute to ϕ , but increase when species' dynamics in a community are negatively
731 correlated. In other words, different species may rise to (and fall from) dominance through time
732 in response to, for example, changing conditions in the environment or availability of resources,
733 such that any given ecosystem property is steadily provided by a rotating cast of characters.
734 Compensation effects can occur in ecosystems due to species interactions (Del Río et al., 2017)
735 or due to the differential responses of species to their environments (Brown et al., 2016).

736 While asynchrony among species dynamics within a community—or among harvest
737 dynamics within a harvest portfolio—will stabilize the aggregate properties of these systems,
738 how asynchrony within a community relates to asynchrony in the harvests derived from that
739 community depends on both ecological and social factors (J. W. Moore et al., 2021; Oken et al.,
740 2021). Asynchrony in stock population dynamics can enhance the stability of the fishery
741 portfolios derived from those stocks, as fishing effort can switch among individual stocks as they
742 become more or less productive from year to year (Hilborn et al., 2003; J. W. Moore et al., 2021;
743 Nesbitt & Moore, 2016; Oken et al., 2021; Schindler et al., 2010). Similarly, seasonal (within-
744 year) compensatory dynamics among stocks extend fishing seasons as harvestable biomass
745 remains accessible to fishers for longer (Nesbitt & Moore, 2016; Schindler et al., 2010).
746 However, the stability imparted to the fishery portfolio by stock population asynchrony is
747 mediated by fisher behaviors in response to regulatory and economic factors. For example, if the
748 harvest of a particular stock is limited due to management regulations, if many fishers are unable
749 to attain licenses to target a particular stock, or if diversifying across stocks is cost prohibitive or
750 not profitable (Kasperski & Holland, 2013; Sethi et al., 2014), stability may arise not from
751 inherent asynchrony in the population dynamics of targeted stocks, but rather from societal

752 changes in how fishing effort is distributed across stocks. We note that these processes are not
753 mutually exclusive, and in fact may operate in tandem to further bolster fisheries stability.

754 In this study, we tested the hypothesis that within-year asynchrony in commercial fishing
755 harvest portfolios is driven by within-year asynchrony among the species comprising the harvest
756 portfolio. To accomplish this, we partitioned asynchrony into compensatory and statistical
757 averaging effects among species and then among harvests. We predicted that seasonal
758 compensation in population dynamics would be positively associated with the seasonal
759 compensation of harvests. We then extended this analysis to test the hypothesis that within-year
760 harvest asynchrony increased within-year fishery economic stability and the overall value of the
761 harvest portfolio.

762 We explored these hypotheses within communities of demersal finfish and the associated
763 commercial fisheries in the mainstem of Chesapeake Bay, USA (Fig. 2). The Chesapeake Bay
764 (hereafter, ‘the Bay’) is among the largest and most economically and ecologically valuable
765 estuaries in the world and supports numerous economically and culturally valuable fisheries
766 (Sanchirico et al., 2008). Like many other temperate estuaries, fishes in the Bay are transient and
767 migrate into and out of the estuary seasonally (Buchheister et al., 2013; Schonfeld et al., 2022).
768 The fish community in the Bay is also heavily structured by salinity, with distinct communities
769 dominating more saline waters near the mouth versus those in the comparably fresher headwaters
770 (Lefcheck et al., 2014), and commercial fisheries operating in these regions fall under different
771 management regimes, although several species within the portfolio of stocks are collectively
772 overseen by a regional authority. The relatively low estuarine diversity, naturally asynchronous
773 dynamics among harvested species through time, and well-documented regulatory structuring
774 along a prominent spatial environmental gradient make the Chesapeake Bay an ideal test case for

775 evaluating how asynchrony among populations of exploited species relates to asynchrony in
776 dependent fisheries.

777 **Methods**

778 *Methods overview*

779 To test the hypotheses that within-year species asynchrony among targeted fishes in
780 Chesapeake Bay led to within-year harvest asynchrony and harvest portfolio yield, stability, and
781 economic value, we first developed monthly time series of species-specific total biomass
782 estimates (“biomass indices”) for six species of finfish commonly targeted by commercial
783 fishing in the Maryland and Virginia regions of the Bay mainstem (Fig. 2). We used these
784 indices to calculate within-year species asynchrony and its components—species compensation
785 and species statistical averaging effects—within both regions.

786 We then used the total monthly harvests of these six species from the mainstem of the
787 Bay in Maryland and Virginia to calculate harvest asynchrony and its components—harvest
788 compensation and harvest statistical averaging effects. Next, we related seasonal species
789 compensation and statistical averaging effects to seasonal harvest compensation and statistical
790 averaging effects using structural equation models (SEMs). We predicted that harvest and
791 species compensation would be positively associated given strong seasonal patterns in fish
792 abundances in the Bay and the expectation that seasonal harvests from these species would
793 correlate with their seasonal abundances. In Maryland, the striped bass (*Morone saxatilis*) fishery
794 is closed for part of the year, so we expected that harvest diversification outside of striped bass
795 during this closure would increase harvest asynchrony.

796 Lastly, we summed harvest yields and harvest values across species within-years to
797 create monthly indices of portfolio harvests and value. We used these indices to calculate annual

798 harvest yield and value stability, as well as total annual harvest yields and value. We extended
799 the SEMs from the previous analysis to evaluate the indirect linkages between species
800 asynchrony (partitioned as species compensation and statistical averaging effects) and harvest
801 portfolio yield, stability, and economic value. We identified that commercial fishing effort
802 declined over the study period, and so we related an index of overall fishing effort to the
803 components of harvest asynchrony in SEMs.

804 *Fish community dynamics and commercial fishing in the Bay*

805 In the northern mainstem of the Chesapeake Bay (blue region in Fig. 2), fisheries
806 management falls under the jurisdiction of the US state of Maryland (MD; via the Maryland
807 Department of Natural Resources [MD-DNR]), and long-term average salinities within this
808 region vary from fresh to mesohaline (0–15 PSU) with proximity to freshwater inputs to the
809 north and west (Harding et al., 2019). Demersal fish biomass in MD waters is dominated by
810 anadromous white perch (*Morone americana*) and striped bass (*Morone saxatilis*), especially in
811 spring prior to moving into Bay tributaries to spawn and in autumn when these fishes aggregate
812 in deeper waters (Buchheister et al., 2013). In the mainstem alone, striped bass harvests in 2002–
813 2018 were valued at \$3.1 million USD per year, making them the most valuable finfish harvested
814 in this region, and white perch \$276,000 USD per year, the third most valuable species (this
815 study; data from MD-DNR). In the MD waters of Chesapeake Bay, commercial striped bass
816 harvests are prohibited during the spring spawning season (March–May).

817 In the southern mainstem of the Bay (orange region in Fig. 2), fisheries management falls
818 to the commonwealth of Virginia (VA; via the Virginia Marine Resources Commission
819 [VMRC]). Salinities in this region increase to euhaline at the interface of the Bay and Atlantic
820 Ocean, and demersal fish biomass is frequently dominated by species that seasonally migrate

821 onto the adjacent continental shelf to spawn (Buchheister et al., 2013). These species include the
822 major fishery targets Atlantic croaker (*Micropogonis undulatus*) and spot (*Leiostomus*
823 *xanthurus*), whose biomasses tend to peak in the spring and summer following the migration and
824 recruitment of adults and young-of-the-year fish into the Bay from the continental shelf, as well
825 as the ontogenic movement of juveniles from shallow Bay nurseries to the Bay mainstem
826 (Buchheister et al., 2013). In the VA mainstem, the three most valuable commercial finfish
827 fisheries include Atlantic croaker (average harvest value in 2002–2018 was \$2.3 million
828 USD/year), striped bass (\$1.65 million USD/year), and spot (\$1.27 million USD/year) (this
829 study; data from VMRC). Regulations in the Bay prevent commercial harvest of Atlantic croaker
830 prior to January 15 of each calendar year, striped bass prior to January 16, and spot prior to April
831 15 or after December 8. Given the unique regulatory, environmental, and ecological
832 characteristics of the northern (MD) and southern (VA) regions of the Bay, we considered the
833 ecosystems and fisheries in these regions to be distinct and analyzed them separately.

834 *Bottom-trawl survey and commercial harvest data*

835 In order to characterize demersal finfish population dynamics, we derived within-year
836 biomass indices from the Chesapeake Bay Multispecies Monitoring and Assessment Program
837 (ChesMMAP) bottom trawl survey (Latour et al., 2003). During 2002–2018, ChesMMAP cruises
838 usually surveyed the entire Bay mainstem five times annually (targeting March, May, July,
839 September, and November), sampling 300–400 locations per year. However, surveys in some
840 month-year combinations did not occur, and in others sample coverage was relatively sparse
841 (detailed below). The survey followed a stratified random design within three depth ranges (3.0–
842 9.1 m, 9.1–15.2 m, and >15.2 m) (Buchheister et al., 2013).

843 We estimated biomass indices (see below) for demersal finfishes that were both well-
844 sampled by the bottom trawl survey (see Buchheister et al., 2013) and for whose cumulative
845 commercial harvests made up >90% of total harvest yield over the study period. In MD, these
846 species included Atlantic croaker, striped bass, white perch, gizzard shad (*Dorosoma*
847 *cepedianum*), and channel catfish (*Ictalurus punctatus*). While gizzard shad and channel catfish
848 were only found in ~5% and ~2% of ChesMMAAP trawls, respectively, these species are
849 considered well-sampled in the Bay mainstem by the bottom trawl (Buchheister et al. 2013).
850 After excluding blue crabs and Atlantic menhaden (a pelagic species poorly-sampled by the
851 bottom trawl [Bucheister et al., 2013]), the combined harvests of these five species contributed to
852 ~94% of total harvests in the MD mainstem over the study period. In VA, these species included
853 Atlantic croaker, spot, and striped bass. We selected these species because their combined
854 harvests accounted for ~93% of total harvests over the study period after excluding blue crab and
855 Atlantic menhaden.

856 We evaluated commercial harvest dynamics for these species in the regions of MD and
857 VA that overlapped with ChesMMAAP surveys using monthly harvest yield and value data
858 provided by MD-DNR and VMRC (Fig. 3). We also received data on fishing effort in the form
859 of the number of trips per month that fishers encountered each species in both MD and VA. We
860 created a combined index of trips across species within each region that was the mean number of
861 trips across species annually. To understand how the closure of the striped bass fishery, the most
862 valuable fishery in the MD harvest portfolio, impacted the stability and value of the portfolio, we
863 created an index called the seasonal harvest ratio (SHR). The SHR was the ratio of the average
864 monthly harvests occurring during the striped bass fishery closure to the average monthly
865 harvests occurring outside the striped bass fishery closure. When the SHR was high, harvests of

866 species other than striped bass were compensating for the absence of striped bass harvests during
867 the closure.

868 *Biomass index modeling*

869 We used generalized additive mixed models (GAMMs) to develop within-year biomass
870 indices for these species in MD and VA waters to compare with harvests in these regions. The
871 GAMMs included penalized smooth terms and spatial and spatiotemporal Gaussian Markov
872 random fields. We modeled biomass indices with Tweedie observation error and a log link
873 (Tweedie, 1984). All models included categorical predictors corresponding to each of five target
874 sampling months for the ChesMMAV survey. We allowed for interannual variability in month-
875 year predictions by including factor-smooth interactions (bs = “fs” in the R package *mgcv* [S. N.
876 Wood, 2006]) for numeric month by year; treating month as numeric because surveys sometimes
877 occurred in months adjacent to the “target” month. For the spot biomass model, we included a
878 thin-plate regression smoother for the numeric year-month combination (S. N. Wood, 2006), and
879 for the gizzard shad and channel catfish models, we included categorical predictors for year
880 factors. These changes improved convergence and residual behavior for these models.

881 There were four instances where cruises did not occur in 2002–2018 (out of 85 possible;
882 4.7%). In fitting each model, we also excluded trawls collected from regions where a given
883 species was not found throughout all surveys (Commander et al., 2022). This reduced the
884 number of observations available for fitting these models, especially for anadromous fishes
885 whose distributions were limited to the northern mainstem. However, the inclusion of smooth
886 terms allowed us to interpolate reasonable biomass estimates representing typical seasonal
887 densities for each species during these periods and for when there were relatively few
888 observations (e.g., in 3 of 85 month-year combinations, there were fewer than 30 trawls available

889 for fitting the white perch model). To calculate area-weighted biomass indices, we predicted
 890 from each model over a grid with 2 km² cells derived from the boundary of the ChesMMA
 891 survey (Fig. 2) and summed these predictions within months. The model for species besides spot,
 892 gizzard shad, and channel catfish can be represented as

$$\begin{aligned}
 \mathbb{E}[y_{s,t}] &= \mu_{s,t}, \\
 \mu_{s,t} &= \exp(\alpha_{\text{season}} + f_{\text{year}}(\text{month}) + \omega_s + \epsilon_{s,t}), \\
 \omega_s &\sim \text{MVN}(\mathbf{0}, \mathbf{\Sigma}_\omega), \\
 \epsilon_{s,t} &\sim \text{MVN}(\mathbf{0}, \mathbf{\Sigma}_\epsilon),
 \end{aligned}
 \tag{1}$$

894 where $\mu_{s,t}$ represents the mean at location s and time t , α_{season} an overall mean for each target
 895 month, $f_{\text{year}}(\text{month})$ the numeric month-year factor-smooth interaction with common
 896 wiggleness, ω_s a spatial random field value (varies by space but is constant with time), and $\epsilon_{s,t}$ a
 897 spatiotemporal random field value that is independent each year. The Gaussian random fields
 898 were approximated as Gaussian Markov random fields using the Stochastic Partial Differential
 899 Equation (SPDE) approach (Lindgren et al., 2011) with covariance matrices $\mathbf{\Sigma}_\omega$ and $\mathbf{\Sigma}_\epsilon$, each
 900 constrained by a Matérn covariance function (Cressie & Huang, 1999).

901 We fitted these models in R using the package *sdmTMB* (S. C. Anderson et al., 2022) via
 902 maximum marginal likelihood. *sdmTMB* integrates SPDE matrices from R-INLA (Lindgren &
 903 Rue, 2015) with marginal log likelihood calculations and random effect integration via the
 904 Laplace approximation with TMB (Kristensen et al., 2015). Our ‘mesh’ for the SPDE
 905 calculations was constructed in R-INLA with an inner mesh near the data and an outer mesh
 906 further away from the data to reduce boundary effects (Lindgren & Rue, 2015). Our inner and
 907 outer meshes had an ‘offset’ of 3 km and 8 km from the data, maximum triangle edge lengths of
 908 3 km and 15 km, and a minimum triangle edge length of 3 km. We assessed model convergence
 909 by checking that the maximum absolute gradient of the marginal log likelihood with respect to

910 all fixed effects was < 0.001 and the Hessian matrix was positive definite. The time series
911 derived from these models are shown in Figure 4.

912 *Partitioning stability and asynchrony and hypothesis testing*

913 We applied the stability and asynchrony partitioning framework developed by Thibault &
914 Connolly (2013; who considered the related quantities of variability and synchrony) that is
915 commonly used in ecological contexts (Lamy et al., 2019; Wang et al., 2019) to both the
916 communities of species being targeted by commercial fisheries in the Bay and to the harvest
917 portfolios in MD and VA. When applied to the fishery harvest portfolio, the stability of intra-
918 annual portfolio harvests is $S_{Portfolio}$, or the inverse of total harvest variability, given by

$$919 S_{Portfolio} = \left(\frac{\sigma_R}{\mu_R} \right)^{-1}. \quad (2)$$

920 Here, σ_R is the temporal standard deviation of total harvests and μ_R the temporal mean of total
921 harvests in a region R . $S_{Portfolio}$ is equivalent to the product of harvest asynchrony ($\phi_{Harvest}$)
922 and species stability ($S_{Harvest}$), the latter term being given by

$$923 S_{Harvest} = \left(\sum_i^n CV_i \times \frac{\mu_i}{\mu_R} \right)^{-1}, \quad (3)$$

924 where CV_i is the temporal coefficient of variation of harvests of species i , and CV_i is weighted by
925 species harvest contributions to μ_R . Stated plainly, $S_{Harvest}$ is the inverse of weighted average
926 harvest variability (Zhao et al., 2022). We calculated $S_{Portfolio}$ twice: once corresponding to the
927 5 months for which biomass estimates were available ($S_{Portfolio,S}$), and again for all 12 months
928 of the year ($S_{Portfolio,L}$). We similarly calculated the stability of portfolio value by calculating the
929 inverse temporal variability of summed harvest values (all months included).

930 We next partitioned species asynchrony ($\phi_{Species}$) and harvest asynchrony ($\phi_{Harvest}$)
931 into compensation (CPE) and statistical averaging effects (SAE) (Zhao et al., 2022). Continuing

932 in the context of the commercial harvest portfolio, harvest asynchrony is given by $\phi_{Harvest} =$

933 $\frac{\sum_i \sigma_i}{\sigma_R} = SAE_{Harvest} \times CPE_{Harvest}$. In the theory of Zhao et al. (2022),

934
$$SAE_{Harvest} = \frac{\sum_i \sigma_i}{\sqrt{\sum_i \sigma_i^2}}, \quad (4)$$

935 where the $SAE_{Harvest}$ is the ratio comparing the portfolio harvest stability if all species harvests

936 were uncorrelated and independent to the portfolio harvest stability if species harvests were

937 perfectly synchronous. $CPE_{Harvest}$ is then

938
$$CPE_{Harvest} = \frac{\sqrt{\sum_i \sigma_i^2}}{\sigma_R}. \quad (5)$$

939 Zhao et al. (2022) note that the CPE is equivalent to the inverse of the square-rooted variance

940 ratio (Schluter, 1984), a commonly used measure of synchrony/asynchrony in the ecological

941 literature. $CPE_{Harvest} > 1$ indicates greater portfolio harvest stability than what would be

942 expected if species harvests fluctuated independently of one another. We partitioned $\phi_{Species}$

943 identically to $\phi_{Harvest}$ (Table 1).

944 We used piecewise structural equation models (SEMs) to relate asynchrony and stability

945 indices across ecosystem and fishery systems (Lefcheck, 2016). Within each SEM, we evaluated

946 sub-models for residual independence, normality, and homogeneity of variances. When residual

947 autocorrelation was present, we modeled residual error as an AR(1) process using generalized

948 least squares (Pinheiro et al., 2017). All stability and asynchrony indices were log-transformed

949 prior to fitting models relating these indices across systems given their multiplicative

950 relationships within-systems. To provide context to these results, we used linear models to model

951 interannual trends in population and harvest dynamics. We used generalized least squares models

952 with varying error structures to account for residual autocorrelation when evident.

953 **Results**

954 *Relationships between the components of species and harvest asynchrony*

955 Across the Chesapeake Bay, we found that harvest compensation was an important driver
956 of harvest asynchrony and the stability of fisheries yields. However, the drivers of harvest
957 asynchrony differed geographically. In MD (Fig. 5, MD SEM 1), harvest compensation was
958 positively associated with the seasonal harvest ratio (SHR) —or the degree to which species
959 other than striped bass dominated the fishery during closed season. This result therefore
960 represents the diversification of harvests during the annual striped bass fishery closure that
961 compensate for the engineered loss of harvestable biomass of this fish (Fig. 6). In contrast, there
962 were no relationships between compensation and statistical averaging in the MD community and
963 fishery portfolio, indicating that seasonal compensation among populations of target species was
964 not responsible for enhancing the portfolio effect in this state. Ultimately, then, asynchrony in
965 MD fisheries harvests was found to be largely promoted by fisher behaviors occurring in
966 response to management strategies, rather than natural fluctuations among exploited species
967 here.

968 During the study period in VA, the total annual biomass of all three dominant species
969 (Atlantic croaker, spot, and striped bass) declined significantly in the lower mainstem of the Bay
970 (Table 1, Fig. 7). However, these biomass declines were not equivalent in magnitude: Atlantic
971 croaker saw the largest declines overall (linear trend = $-1.22 \times 10^5 \text{ kg year}^{-1}$, $P < 0.001$), followed
972 by spot ($-1.91 \times 10^4 \text{ kg year}^{-1}$, $P = 0.003$) and lastly striped bass ($-3.5 \times 10^3 \text{ kg year}^{-1}$, $P = 0.04$).
973 After the substantial loss of Atlantic croaker biomass, the variabilities of species biomasses
974 became more evenly distributed, thereby increasing the species statistical averaging effect over

975 the study period (Fig. 8a). Increases in the species statistical averaging effect due to increased
976 evenness of biomass variabilities among species were then positively associated with species
977 compensation ($P < 0.001$, Fig. 5, VA SEM 1). In other words, following the steep declines in
978 Atlantic croaker biomass, species population dynamics became more seasonally compensatory
979 and therefore asynchronous (Fig. 8b), albeit with substantially lower biomass overall.

980 While species statistical averaging and compensation increased in the VA fish
981 community over the study period, harvest compensation effects showed a weak negative trend (P
982 = 0.079, Fig. 8d, Table 2), and harvest compensation was most strongly associated with species
983 statistical averaging ($\beta_{std} = -0.834$, $P = 0.017$, Fig. 5, VA SEM 1). This suggests that declines
984 in the availability of target species biomass in the Bay were associated with shifts in the timing
985 of harvests that resulted in a decline of harvest compensation. This was supported by the finding
986 that the timing of Atlantic croaker and striped bass harvests shifted to later and earlier the year,
987 respectively (both $P < 0.001$, Fig. 9, Table 3), such that the timing of Atlantic croaker harvests
988 became more similar to that of spot, whereas striped bass harvests became more dominated by
989 harvests in the spring months. In addition to species statistical averaging being associated with a
990 shift in harvest timing, we also found a strong correlative association between species statistical
991 averaging and the average number of fishing trips across species ($P < 0.001$, VA SEM 1, Fig. 5),
992 which declined throughout the 2002-2018 period across stocks (Fig. 10a). We characterized this
993 relationship as correlated error because effort changes in response to biomass declines may be
994 related to exogenous factors like harvest prices, which we did not consider. Interestingly, despite
995 steep declines in biomass availability, catches per trip did not show significant trends across the
996 study period (Fig. 10b). This suggests that changes in harvest effort played a substantial role in
997 the decline in harvests overall in VA.

998 We found that harvest compensation was positively associated with species compensation
999 ($\beta_{std} = 0.636, P = 0.032, \text{Fig. 5, VA SEM 1}$), suggesting that seasonal compensatory dynamics
1000 in population biomasses drove seasonal harvest compensation. However, the magnitude of the
1001 standardized partial effect of species statistical averaging on harvest compensation was greater
1002 than that of species compensation on harvest compensation and opposite in sign ($SAE_{\beta_{std}} =$
1003 -0.834 vs $CPE_{\beta_{std}} = 0.636$). This result means that the stabilizing effects of seasonal species
1004 dynamics on harvest compensation were overwhelmed by the destabilizing effects of species
1005 statistical averaging, representing the decline in target species biomasses that was associated
1006 with more temporally synchronous harvests (Fig. 8d, 9).

1007 *Portfolio harvest stability and economic performance*

1008 We find that portfolio harvest stability was enhanced by harvest asynchrony, which the
1009 previously discussed results show was driven by fisher behavioral responses to the striped bass
1010 closure (MD) and to the relative biomasses and variabilities of fish populations as they naturally
1011 migrate into and out of the Bay mainstem seasonally (VA). Using two additional SEMs, we
1012 expanded this analysis to evaluate how portfolio harvest stability associated with portfolio value
1013 stability and overall value.

1014 In MD (Fig. 5, MD SEM 2), we found that portfolio harvest stability quantified at the
1015 same temporal resolution as the population biomass indices (i.e., $S_{\text{Portfolio}, S}$) was positively
1016 associated with portfolio harvest stability quantified using all months of harvest data ($S_{\text{Portfolio}, L}$).
1017 We quantified harvest stability at both resolutions because biomass indices were only resolved at
1018 the five-month timescale, whereas fishing harvest data were available for all 12 months. This
1019 finding shows that harvest asynchrony induced by fisher target diversification in response to the
1020 striped bass fishery closures (Fig. 6) was positively linked to the within-year harvest stability of

1021 the portfolio. However, portfolio harvest stability was not linked to the economic stability of
1022 harvests. In other words, active switching by fishers buffered the stability of the portfolio against
1023 regulatory measures, in the form of the seasonal closure, but that this stabilizing effect did not
1024 translate to enhanced economic stability. Lastly, we identified a strong positive relationship
1025 between total yield and total value ($P < 0.001$), showing that years with greater harvests were
1026 associated with greater portfolio value (Fig. 5, MD SEM 2).

1027 In VA (Fig. 5, VA SEM 2), where we also identified a positive link between portfolio
1028 harvest stability at the 5-month and 12-month temporal resolutions, portfolio harvest stability
1029 ($S_{\text{Portfolio, L}}$) declined substantially over the study period (Fig. 11, Table 4, $P < 0.001$) in response
1030 to diminished harvest compensation (Figs. 8b) and species harvest stability, the latter of which
1031 was associated with fishing effort (Fig. 10). In other words, the loss of harvestable biomass,
1032 primarily of Atlantic croaker, was associated with more temporally synchronous harvests, and
1033 this coincided with decreasing fishing effort that drove more temporally variable and lower
1034 harvests within-years. Both of these factors contributed to interannual declines in harvest
1035 portfolio stability overall. Portfolio harvest stability was positively associated with the stability
1036 of portfolio value, as well as total harvest yield (Fig. 5, VA SEM 2).

1037 **Discussion**

1038 Asynchronous dynamics among different species can promote community stability over
1039 time, but this effect could be translated to—or compromised by—the active exploitation of these
1040 species by humans. We approached this problem by analyzing 17 years of data from fisheries-
1041 independent surveys (i.e., community data) and fisheries landings across one of the largest and
1042 most economically valuable estuaries in the world, the Chesapeake Bay. We found that
1043 asynchrony in Chesapeake Bay fish harvests within a given year emerged from two sources:

1044 fisher-behavior responses to management strategies (i.e., the closure of the striped bass fishery in
1045 Maryland), and from the components of asynchrony among fish species that emerged due to
1046 seasonal population dynamics (compensation) and the unevenness of population variabilities
1047 (statistical averaging effects). However, while species compensation was a positive contributor
1048 to harvest compensation, these effects were counteracted by the negative effect of species
1049 statistical averaging on harvest compensation, which reflected a decline in overall biomass
1050 availability and increased evenness of variabilities in the ecosystem. In Virginia, we also found
1051 that the stability of the harvest portfolio was positively associated with the stability of the value
1052 of the harvest portfolio. These findings demonstrate that conserving asynchronously fluctuating
1053 exploited natural resources can enhance the within-year stability of their associated harvest
1054 portfolios and their value, but that these stabilizing effects can be overwhelmed by the relative
1055 availability and variability of targets in the environment, human behavior, and management
1056 regulations.

1057 Our application of this framework to a coupled social-ecological system is highly
1058 relevant, as it more fully considers the mechanisms through which asynchrony among species in
1059 a biological community contributes to stability while the system is simultaneously being both
1060 exploited and actively managed. Several authors have explored similar relationships between
1061 community dynamics and harvest asynchrony (J. W. Moore et al., 2021; Nesbitt & Moore, 2016;
1062 Schindler et al., 2010), but have not treated the populations targeted by fisheries and their
1063 associated harvests as hierarchically structured portfolios that are linked by asynchrony. By
1064 conceptualizing these systems separately, we were able to identify the specific components of
1065 asynchrony that contributed to (in the case of MD) or diminished (as in VA) these coupled
1066 systems.

1067 Cross-system asynchrony relationships were limited in MD, where harvest asynchrony
1068 was strongly related to the adaptive response of fishers to management strategies limiting access
1069 to the valuable striped bass fishery. During the closure, fishers targeted species like Atlantic
1070 croaker, white perch, and gizzard shad. In this sense, the striped bass closure ultimately enhanced
1071 the stability of the regional harvest portfolio that became more diversified as a result, although
1072 this stabilizing effect did not translate to stability of the value of the harvest portfolio, possibly
1073 due to the lower value of species being targeted during the striped bass closure compared to
1074 striped bass.

1075 Portfolio harvest stability in VA waters declined over the study period, and this decline
1076 was related to ecological and social factors impacting species and harvest asynchrony and
1077 species harvest stability. First, through the declines in the biomass of Atlantic croaker and spot in
1078 the lower Bay (also documented in Buchheister et al., 2013; Schonfeld et al., 2022), a
1079 relationship that was captured in VA SEM 1 (Fig. 5) through the effect of species statistical
1080 averaging on harvest compensation. When species biomasses were most uneven (high biomass
1081 availability), their harvests tended to be more seasonally compensatory, contributing to the
1082 stability of the harvest portfolio overall. However, as population variabilities became more even
1083 (low biomass availability), species harvests became more synchronous, and this contributed to
1084 declining harvest stability interannually.

1085 Second, we identified a clear decline in the number of commercial fishing trips occurring
1086 in VA over the study period. While understanding the complete suite of factors influencing this
1087 decline is beyond the scope of our study, our finding is supported by recent work showing
1088 declines in finfish fishery participation in VA (in terms of the number of licensed fishers) (White
1089 & Scheld, 2021), although that study found that these declines may be due largely to

1090 demographic factors. Declining trends in effort may also be due to market factors that we did not
1091 characterize (e.g., prices), with commercial fishers being tempted to other fisheries and marine-
1092 related industries that may be more profitable and with greater promise of future reliability in
1093 harvest, including but not limited to the growing aquaculture industry, as well as fisheries for
1094 blue crabs (*Callinectes sapidus*) and hard clams (*Mercenaria mercenaria*) (White & Scheld,
1095 2021). Increasing participation in these industries by fishers at the expense of effort in the finfish
1096 fisheries documented here could help explain effort declines, but further work is needed to better
1097 understand the drivers of effort declines.

1098 Like many fishes inhabiting temperate ecosystems, the life histories of Atlantic croaker,
1099 spot, and striped bass involve a seasonal migratory phase wherein fishes seasonally move into
1100 and out of inshore habitats from the coastal ocean (Buchheister et al., 2013; Murdy & Musick,
1101 2013). Importantly, the peak abundances of these species in VA waters are typically seasonally
1102 compensatory, implying that their harvest dynamics may also be seasonally compensatory, and
1103 that this compensation would stabilize the harvest portfolio. We found evidence for this effect,
1104 although the positive contribution of seasonal species compensation to harvest asynchrony was
1105 overwhelmed by the negative effect of species statistical averaging. This finding suggests that
1106 seasonal species compensation was and is not effective in stabilizing within-year harvests in the
1107 face of fewer fish entering into the Bay (Schonfeld et al., 2022). Such biomass declines challenge
1108 the capacity for fishers to diversify among finfish species since fishers in the Bay typically do
1109 not operate outside of Bay waters where these species may be found in higher numbers
1110 (Schonfeld et al., 2022).

1111 However, one diversification opportunity that may prove important for providing stability
1112 to commercial fishers in the Bay that would also not require leaving the Bay to fish is the

1113 emerging blue catfish (*Ictalurus furcatus*) fishery. Blue catfish is an invasive species whose
1114 abundances have increased dramatically in Bay tributaries in recent years, and in VA waters,
1115 harvests of blue catfish have exceeded those of striped bass since 2015 (Fabrizio et al., 2021).
1116 The governments of Maryland and Virginia have recently put forth legislative efforts supporting
1117 the development of a commercial fishery around this species (Senate Joint Resolution 4, 2021;
1118 Blue Catfish Processing, Flash Freezing, and Infrastructure Grant Program, 2023).

1119 The substantial yield of blue catfish in Bay tributaries in recent years demonstrates the
1120 increasing importance of this fishery and raises a limitation of the presented analysis. First, the
1121 Bay is home to several species of finfish and invertebrates that are not considered within this
1122 study. For instance, beyond blue catfish whose low abundances in the Bay mainstem were
1123 inconsequential during our study period (2002–2018), the Bay supports relatively enormous blue
1124 crab fisheries that, at a regional scale, are valued at \$80–\$100 million USD annually (Scheld et
1125 al., 2021). Additionally, in the MD portion of the Bay mainstem, the pelagic Atlantic menhaden
1126 (*Brevoortia tyrannus*) is the second most valuable finfish after striped bass, the harvests of which
1127 were valued at ~\$700,000 USD annually over the study period. While we focused our analyses
1128 on fishes that were well-sampled in ChesMMAAP bottom trawl surveys and so excluded these
1129 important species, if comparable biomass estimates can be made to those presented here, then it
1130 would be reasonable to include these species in a future analysis.

1131 Similarly, our analyses were geographically constrained to ensure the overlap between
1132 commercial harvest data and the ChesMMAAP survey. Future work should consider expanding
1133 this approach to Bay tributaries and their associated commercial and recreational fisheries.
1134 Lastly, we assumed a unidirectional relationship between species abundances and species
1135 harvests, although harvesting will certainly influence the population dynamics of harvested

1136 species (Arlinghaus et al., 2017; J. W. Moore et al., 2021). Future work should explore these
1137 bidirectional dynamics, but doing so here is beyond the scope of this paper.

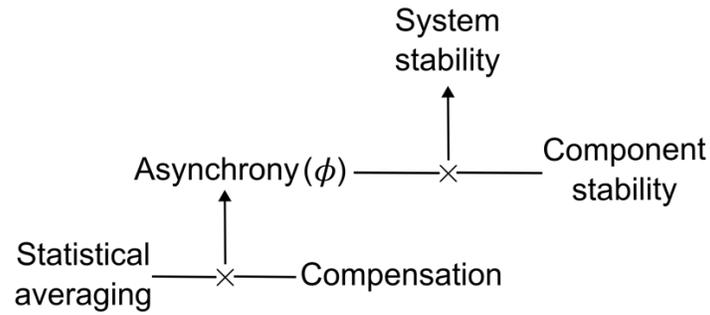
1138 In the present analysis, we have presented novel insights into how species asynchrony
1139 relates to harvest asynchrony to stabilize commercial fishing harvest portfolios. In a system like
1140 MD where the dominant fishery target was seasonally closed to harvesting, we showed that
1141 harvest asynchrony arose due to the diversification of harvests occurring during the closure, and
1142 that this diversification stabilized the within-year harvest portfolio overall, but harvest stability
1143 was not associated with the stability of portfolio value. In VA, declines in the biomass of target
1144 species in the Bay played a key role in how within-year species asynchrony related to harvest
1145 asynchrony. With biomass declines came a greater evenness of population variabilities that was
1146 associated with declines in harvest compensation, and a coincident decline in effort that was
1147 associated with declines in species harvest stability. These relationships drove a negative trend in
1148 the stability of the harvest portfolio, although periods of high portfolio stability were associated
1149 with high portfolio value stability.

1150 Understanding how stability arising from ecosystem processes contributes to the stability
1151 of dependent social-ecological systems has the potential to unlock novel management strategies
1152 that leverage these stabilizing features while concurrently conserving biodiversity.

1153 **Acknowledgements**

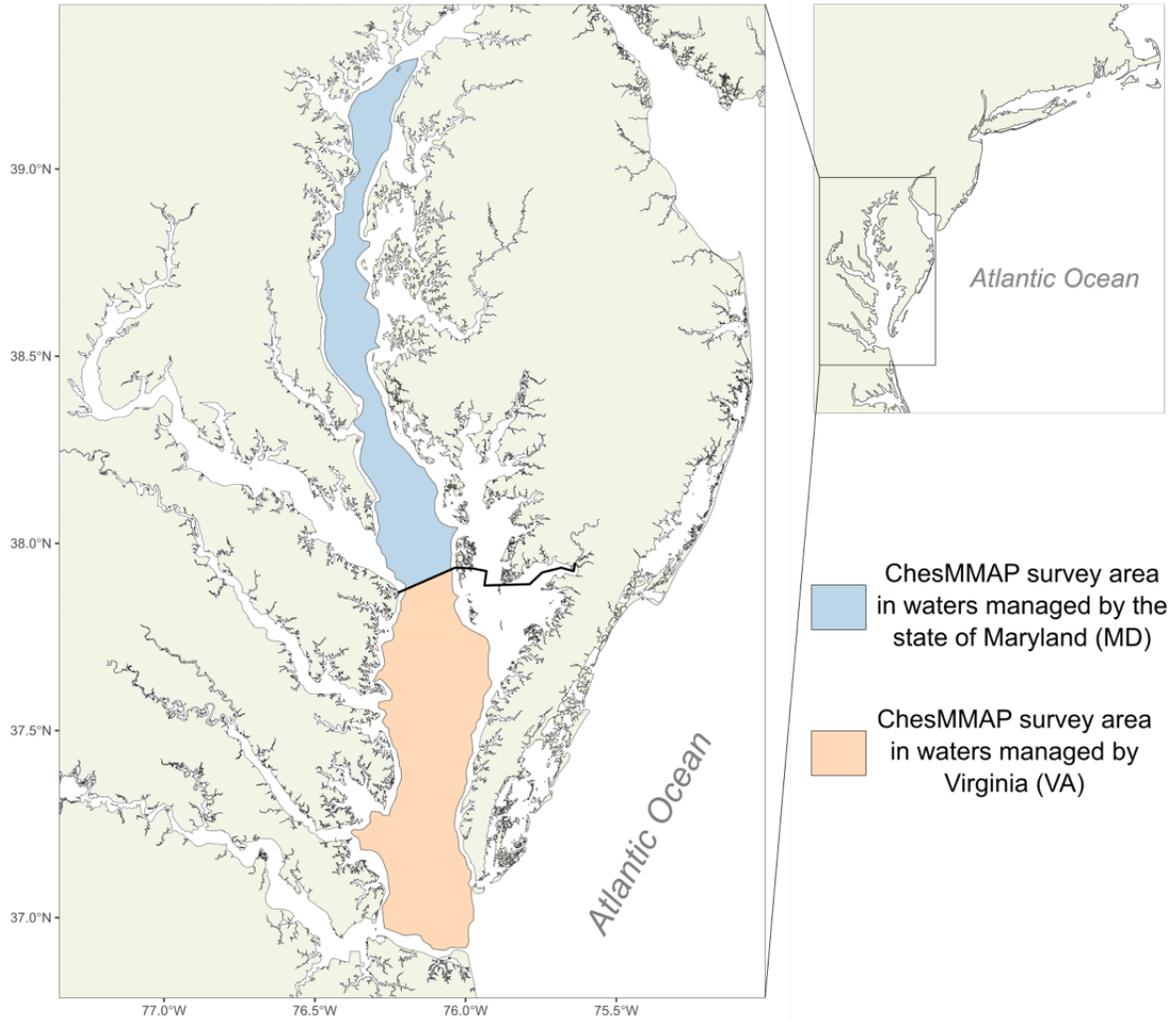
1154 We are grateful to Virginia Sea Grant for their financial support that made this study possible
1155 through a graduate research fellowship, and to the staff and scientists at the Virginia Institute of
1156 Marine Science who collected the data used in this study. We also thank the staff at VMRC and
1157 MD-DMF for providing the commercial harvest data.

1158 **Figures**



1159

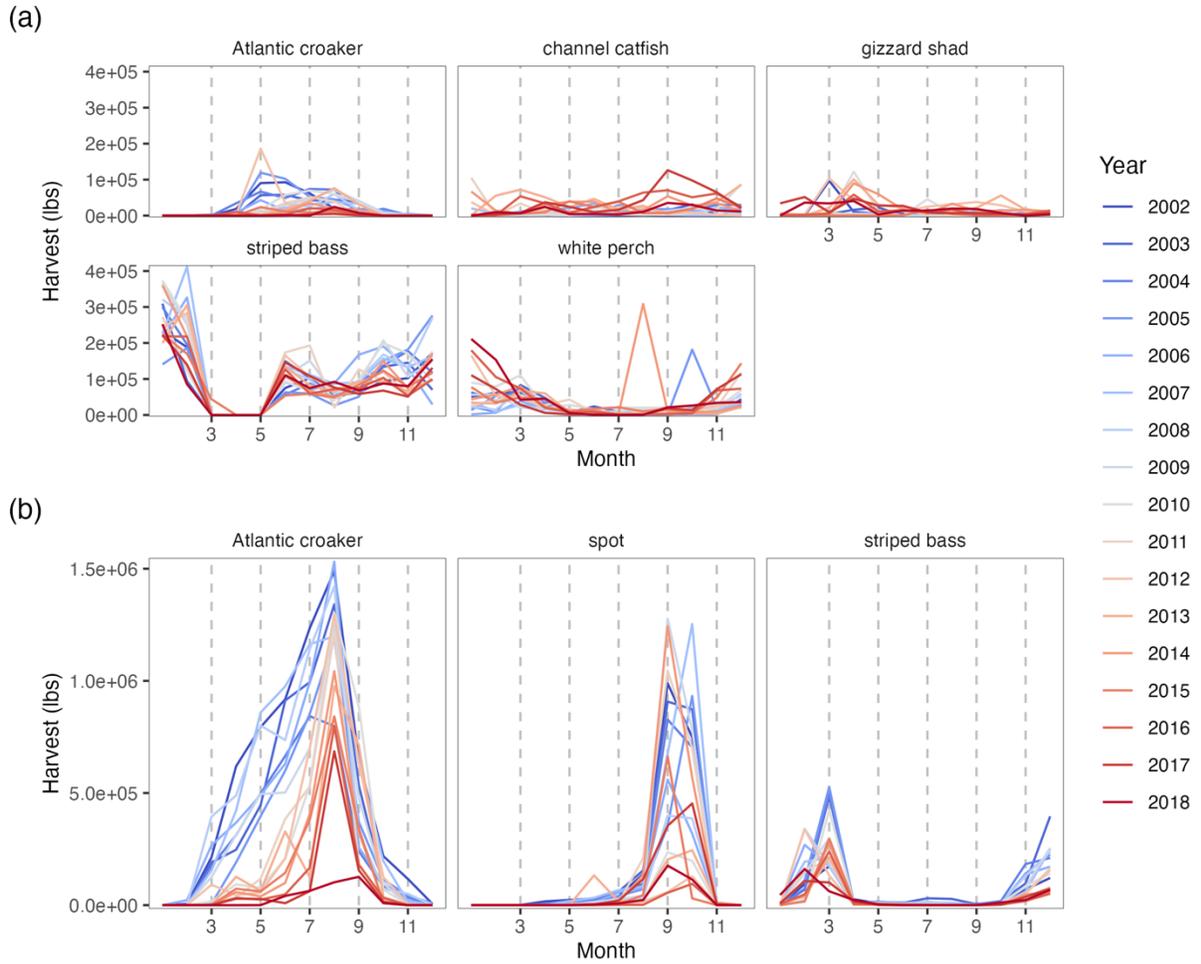
1160 *Figure 1.* Conceptual model describing the generalization of a stability partitioning framework
1161 originally applied in an ecological context (Thibaut & Connolly, 2013), and extended to consider
1162 the asynchrony partitioning method introduced by Zhao et al. 2022. Variables connected by “x”
1163 indicate a multiplicative relation between those variables, where their product is given by the
1164 variable indicated with an arrow.



1165

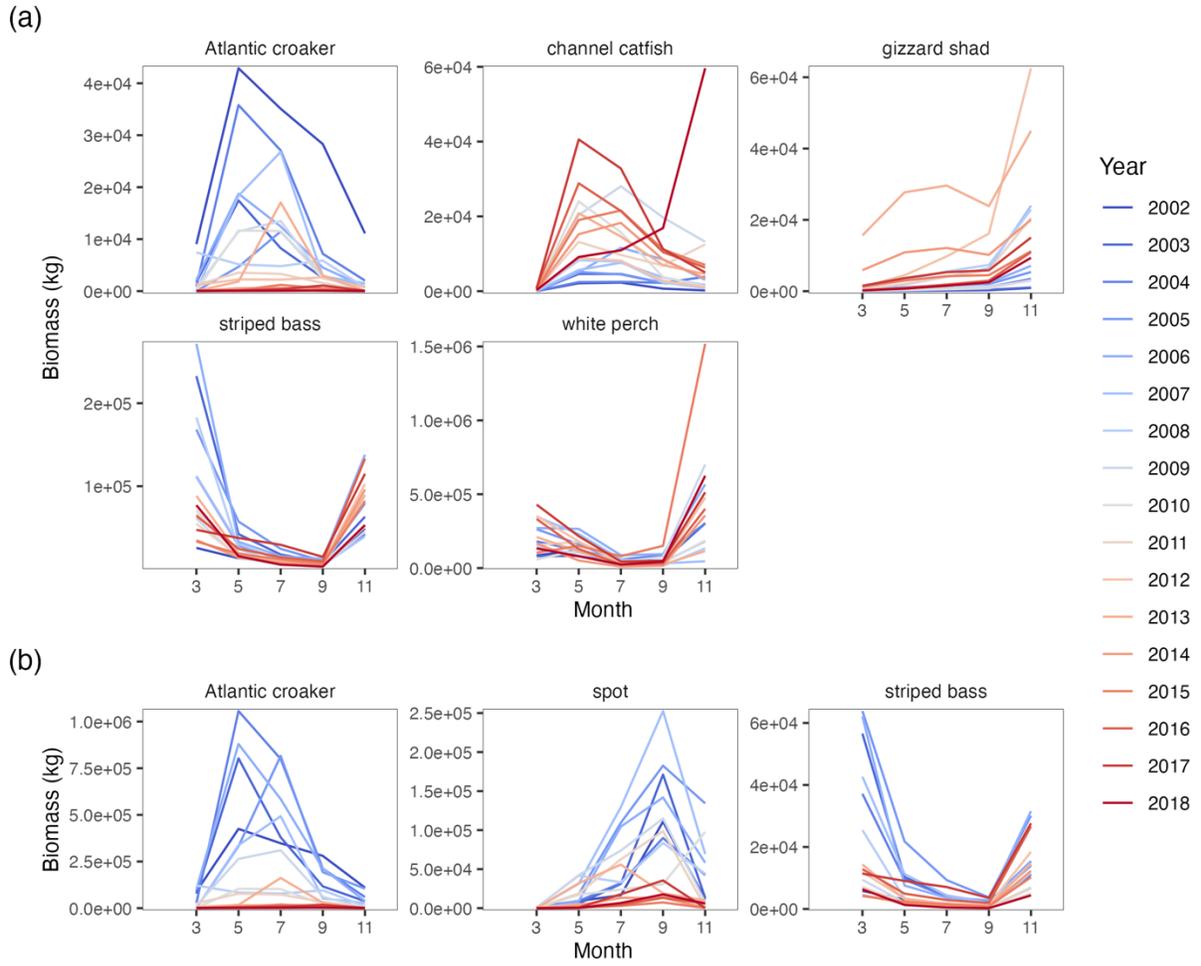
1166 *Figure 2.* Map of Chesapeake Bay and its location along the east coast of the United States
 1167 (inset). The light blue area refers to the area of the ChesMMAP survey falling within waters
 1168 managed by the state government of Maryland (MD), and the light orange area the area falling
 1169 within the management jurisdiction of Virginia (VA). The black line denotes the boundary
 1170 between MD and VA waters.

1171



1172

1173 *Figure 3.* Within-year time series of commercial fishing harvests in Maryland (top, a) and
 1174 Virginia (bottom, b), where each line corresponds to a different year. Vertical lines correspond to
 1175 the months where biomass indices for each species were estimated.

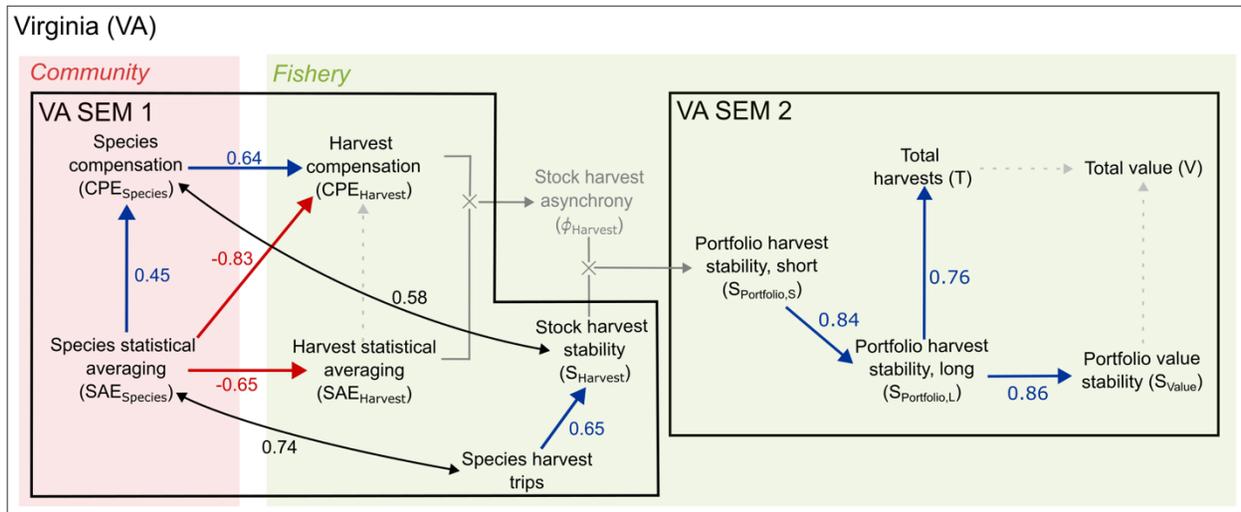
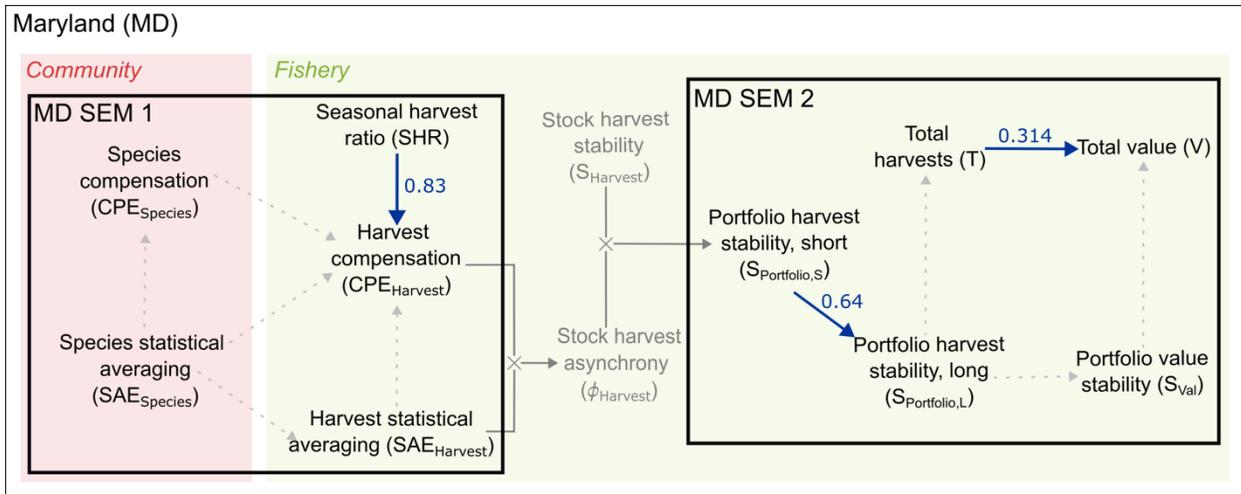


1176

1177 *Figure 4.* Within-year time series of modeled species biomasses in Maryland (top, a) and
 1178 Virginia (bottom, b), where each line corresponds to a different year. Note the change in y-axis
 1179 scales between species.

1180

1181



1182

1183 *Figure 5.* Linked path diagrams showing the relationships between asynchrony in the community

1184 (red) and harvest portfolio (green) in MD (top) and VA (bottom) waters. Variables connected by

1185 an “×” indicate a multiplicative relation between those variables, the product of which is given

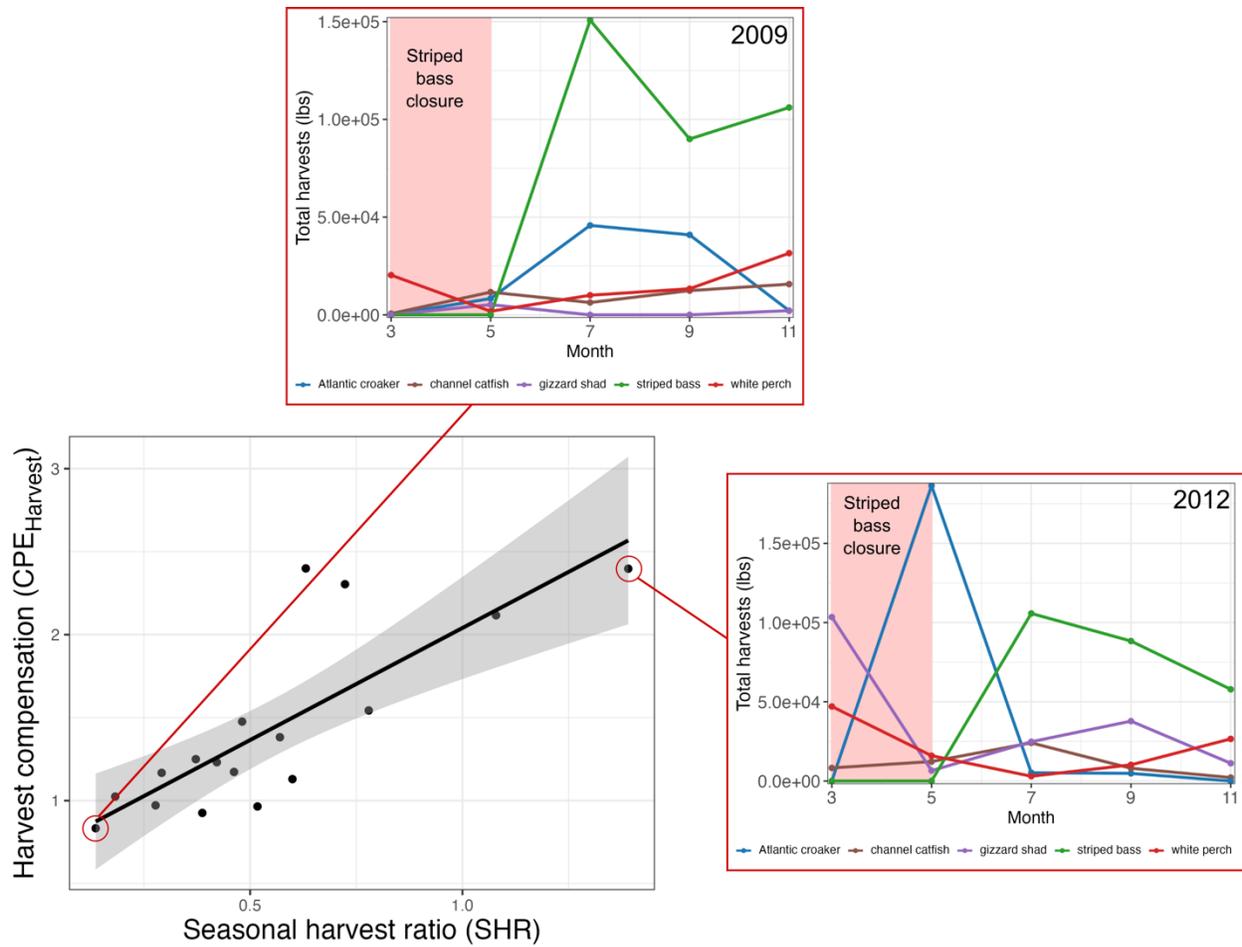
1186 by the variable following the arrow. Statistical relationships were modeled using SEMs, the

1187 hypothesized paths of which are shown in the black boxes: dark blue arrows indicate positive

1188 relationships ($P < 0.05$), red arrows negative, double-headed black arrows correlated errors, and

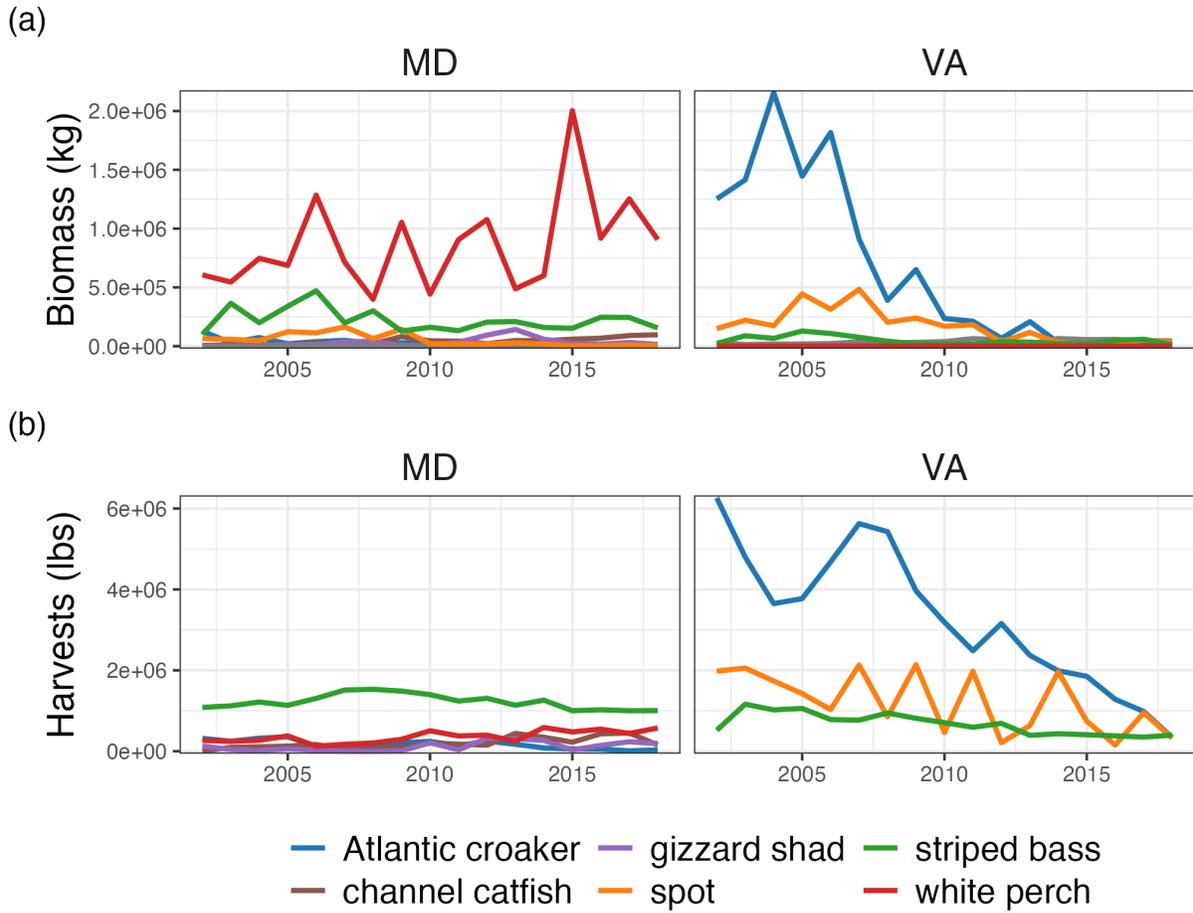
1189 dotted grey lines no relationship. Numbers next to paths show standardized coefficients. See

1190 Table 1 for definitions of terms.



1191

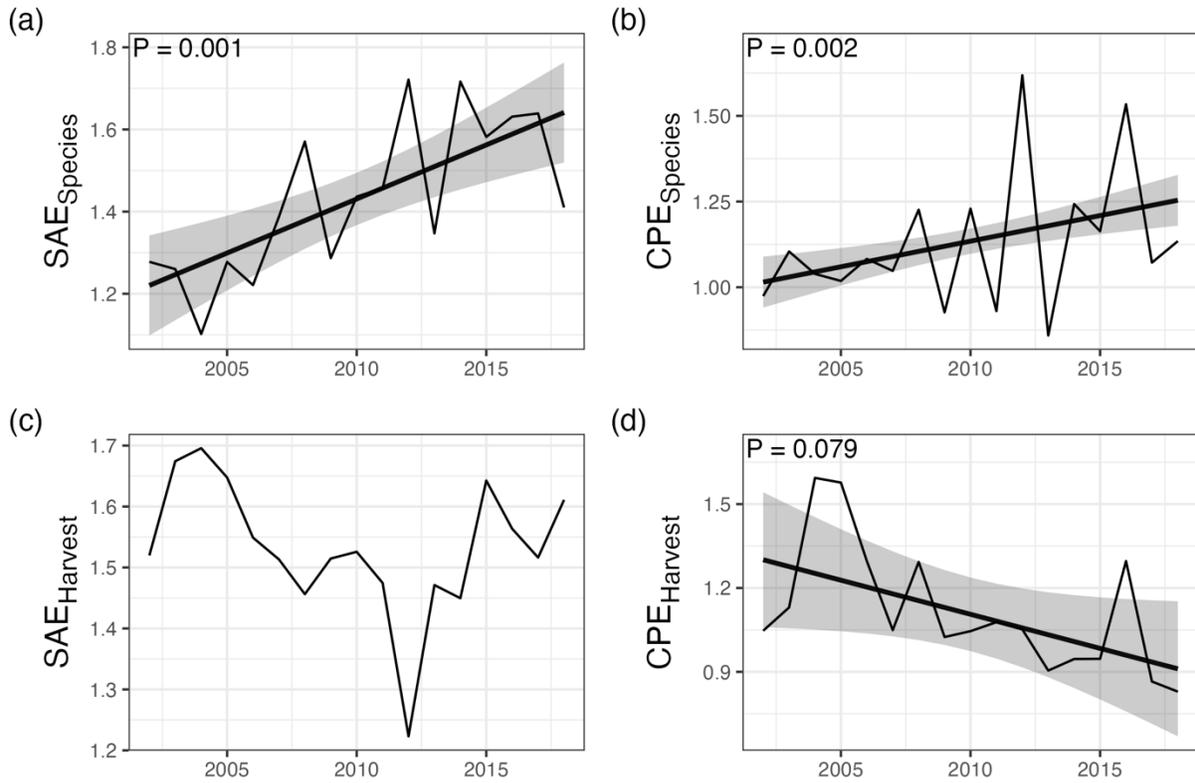
1192 *Figure 6.* The relationship between the harvest compensation effect ($CPE_{Harvest}$) and the seasonal
 1193 harvest ratio (SHR) in MD (bottom left; $R^2 = 0.64$). Also shown are harvest time series from
 1194 Maryland during years when the SHR was highest (right; 2012) and lowest (top; 2009). Red
 1195 shaded areas show the months during which the striped bass fishery was closed.



1196

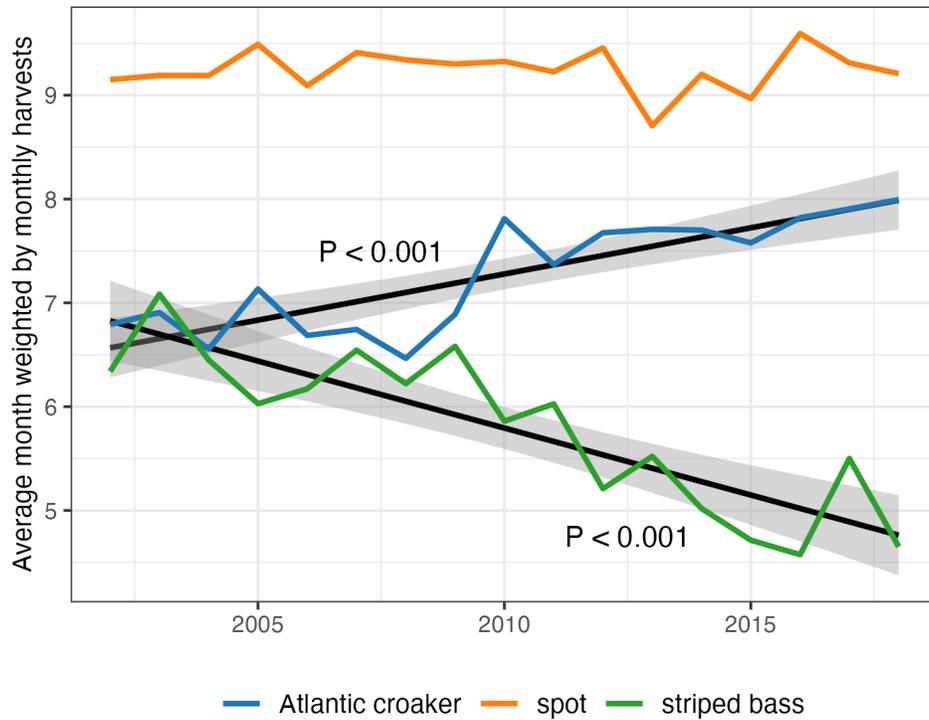
1197 *Figure 7.* Time series of annual biomass (a) and harvests (b) in the Virginia and Maryland

1198 regions of the Chesapeake Bay mainstem between 2002-2018.



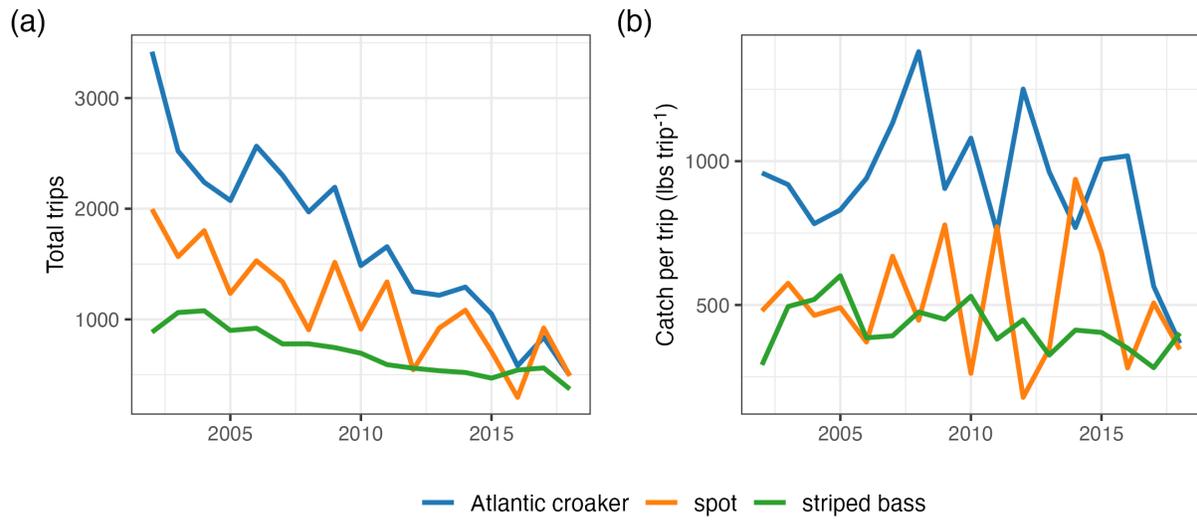
1199

1200 *Figure 8.* Interannual time series of annual species statistical averaging (a, $SAE_{Species}$), species
 1201 compensation effects (b, $CPE_{Species}$), harvest statistical averaging (c, $SAE_{Harvest}$), and harvest
 1202 compensation effects (d, $CPE_{Harvest}$) in Virginia. Trend lines are from linear models shown with
 1203 95% confidence intervals (see Table 2 for details).



1204
 1205 *Figure 9.* Species-specific time series of the weighted-average month of harvest in Virginia,
 1206 where weights are given by monthly harvests. A positive linear trend indicates that the timing of
 1207 harvests shifted to be later in the year, and a negative trend, earlier.

1208



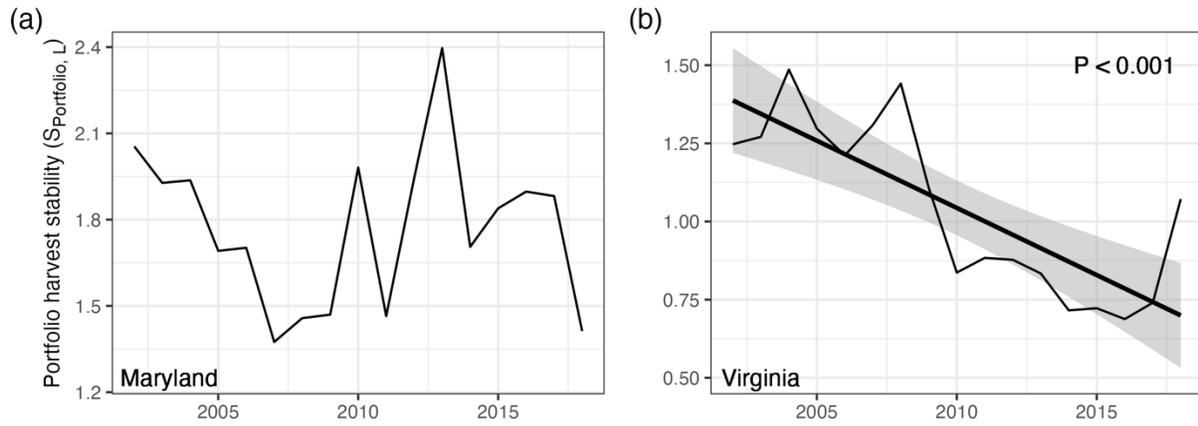
1209

1210 *Figure 10.* Time series of total annual fishing trips across months (a) and total harvest per trip (b)

1211 in the Virginia portion of Chesapeake Bay. The across-species annual average of these time

1212 series is used in VA SEM 1, shown in Fig. 5.

1213



1214

1215 *Figure 11.* Time series of portfolio harvest stability in Maryland (a) and Virginia (b). Virginia
 1216 portfolio harvest stability declined significantly over the study period ($P < 0.001$), as shown by
 1217 the trend line with 95% confidence interval.

1218

1219 Table 1. Definitions of asynchrony and stability terms, their abbreviations, and general mathematical formulas.

<i>Term</i>	<i>Abbreviation</i>	<i>Formula</i>	<i>Description</i>
Total harvest stability (all months)	$S_{Portfolio,L}$	$\left(\frac{\sigma_R}{\mu_R}\right)^{-1}$	The stability of total harvests across all species and months.
Total harvest stability (5 months)	$S_{Portfolio,S}$	$\left(\frac{\sigma_R}{\mu_R}\right)^{-1}$	The stability of total harvests across all species during the five months when biomass estimates are available.
Species harvest stability	$S_{Harvest}$	$\left(\sum_i^n CV_i \times \frac{\mu_i}{\mu_R}\right)^{-1}$	The inverse of weighted-average harvest variability across species. We calculated this index for the five months when biomass estimates are available.
Asynchrony	ϕ	$\frac{\sum_i \sigma_i}{\sigma_R}$	A measure of the stabilizing effects of compensation and statistical averaging; either among harvests of multiple stocks ($\phi_{Harvest}$) or among the biomass dynamics of multiple species ($\phi_{Species}$).
Statistical averaging effect	SAE	$\frac{\sum_i \sigma_i}{\sqrt{\sum_i \sigma_i^2}}$	The enhancement of harvest or community stability following independent fluctuations among species harvests ($SAE_{Harvest}$) or population dynamics ($SAE_{Species}$).
Compensation effect	CPE	$\frac{\sqrt{\sum_i \sigma_i^2}}{\sigma_R}$	The enhancement of harvest or community stability following negatively correlated dynamics among species harvests ($CPE_{Harvest}$) or population dynamics ($CPE_{Species}$).

1220

1221 Table 1. Statistics from linear models evaluating interannual temporal trends in Atlantic croaker,
 1222 spot, and striped bass annual biomass (summed across months) in the Virginia portion of
 1223 Chesapeake Bay over 2002-2018.

Species	Trend (kg year⁻¹)	Std. error	T statistic	P value	Error structure
Atlantic croaker	-122,256	18,549	-6.591	<0.001	iid
spot	-19,140	5,320	-3.598	0.003	iid
striped bass	-3,497	1,546	-2.262	0.039	iid

1224
 1225 Table 2. Statistics from linear models evaluating interannual temporal trends in within-year
 1226 species and harvest statistical averaging ($SAE_{Species}$ and $SAE_{Harvests}$) and within-year species and
 1227 harvest compensation effects ($CPE_{Species}$ and $CPE_{Harvests}$) in the Virginia portion of Chesapeake
 1228 Bay over 2002-2018.

Term	Trend (SAE/CPE year⁻¹)	Std. error	T statistic	P value	Error structure
$SAE_{Species}$	0.026	0.007	3.976	0.001	iid
$CPE_{Species}$	0.014	0.005	2.994	0.009	AR(1)
$SAE_{Harvests}$	-0.005	0.005	-0.994	0.336	iid
$CPE_{Harvests}$	-0.024	0.013	-1.887	0.079	AR(1)

1229
 1230 Table 3. Statistics from linear models evaluating interannual temporal trends in within-year
 1231 weighted-average month of harvest, where weights are monthly harvests in the Virginia portion
 1232 of Chesapeake Bay over 2002-2018.

Species	Trend (month year⁻¹)	Std. error	T statistic	P value	Error structure
Atlantic croaker	0.089	0.014	6.241	<0.001	iid
spot	-0.001	0.011	-0.092	0.928	iid
striped bass	-0.129	0.019	-6.694	<0.001	iid

1233

1234 Table 4. Statistics from linear models evaluating interannual temporal trends in within-year
1235 portfolio harvest stability ($S_{\text{Portfolio,L}}$) in Maryland and Virginia regions of the Chesapeake Bay
1236 over 2002-2018.

Region	Term (stability year⁻¹)	Trend	Std. error	T statistic	P value	Error structure
Maryland	$S_{\text{Portfolio,L}}$	-0.002	0.014	-0.158	0.877	iid
Virginia	$S_{\text{Portfolio,L}}$	-0.043	0.008	-5.130	0.000	iid

1237

1238 Chapter 3

1239 Effects of seagrass restoration on coastal fish abundance and diversity

1240

1241 Citation: Hardison, S., McGlathery, K., & Castorani., C. N. (2023) Effects of seagrass restoration
1242 on coastal fish abundance and diversity. *Conservation Biology*, e14147.

1243

1244 **Abstract**

1245 Restoration is accelerating to reverse global declines of key habitats and recover lost
1246 ecosystem functions, particularly in coastal ecosystems. However, there is high uncertainty about
1247 the long-term capacity of restored ecosystems to provide habitat and increase biodiversity, and
1248 the degree to which these ecosystem services are mediated by spatial and temporal
1249 environmental variability. We addressed these gaps by sampling fishes biannually for 5–7 years
1250 (2012–2018) at 16 sites inside and outside of a rapidly-expanding restored seagrass meadow in
1251 coastal Virginia, USA. Despite substantial among-year variation in abundance and species
1252 composition, seine catches in restored seagrass beds were consistently more numerous (6.4×
1253 more fish, $P < 0.001$) and speciose (2.6× greater species richness, $P < 0.001$; 3.1× greater Hill-
1254 Shannon diversity, $P = 0.03$) than adjacent unvegetated areas, with larger catches particularly
1255 during summer (vs. autumn, $P < 0.01$). Structural equation modeling revealed that depth and
1256 water residence time interacted to control seagrass presence, leading to higher fish abundance
1257 and richness in shallow, well-flushed areas that supported seagrass. Together, our results indicate
1258 that seagrass restoration yields large and consistent benefits for many coastal fishes, but that
1259 restoration and its benefits are sensitive to the dynamic seascapes within which it is established.
1260 Consideration of how seascape-scale environmental variability impacts both the success of

1261 habitat restoration as well as subsequent ecosystem function will improve restoration outcomes
1262 and the provisioning of ecosystem services.

1263 **Introduction**

1264 Human activities have led to global degradation of ecosystem structure and function over
1265 the past century (Ellison et al. 2005; Barbier et al., 2011). In response, habitat restoration has
1266 increasingly been investigated as a tool for combating these declines (Guan et al., 2019),
1267 although there remains uncertainty about the long-term capacity of restored ecosystems to
1268 consistently provide ecosystem services such as habitat provisioning (Lamb 2018). Restored
1269 ecosystems may fail to enhance and sustain ecosystem services due to site-specific
1270 environmental constraints that limit restoration success or the suitability of restored habitat for
1271 associated fauna (Raposa and Talley, 2012; England and Wilkes, 2018). For example, the
1272 direction (positive/negative/zero) and magnitude of effects of salt marsh restoration on nekton
1273 were found to be highly variable and site-dependent at a regional scale due to local-scale
1274 environmental variability and anthropogenic impacts (Raposa and Talley, 2012). Similarly, while
1275 river restorations have been successful in reinstating physical functioning to degraded
1276 waterways, these improvements were not always associated with the rehabilitation of associated
1277 macroinvertebrate communities subject to local- and catchment-scale environmental variability
1278 (Leps et al., 2016; England and Wilkes, 2018). Evaluating how spatial and temporal variation
1279 impacts communities in restored habitats is critical to maximizing the success of restorations and
1280 the ecosystem services they provide.

1281 Following declines of seagrass meadows in shallow coastal ecosystems worldwide
1282 (Dunic et al., 2021), restoration efforts have increased to recover these habitats and their
1283 associated ecosystem services (Duarte et al., 2020; Orth et al., 2020). While numerous studies

1284 have shown the benefits of seagrass for fish and fisheries (e.g., Gilby et al., 2018; Unsworth et
1285 al., 2019; Orth et al., 2020), there is still uncertainty about how spatial environmental variability
1286 directly and indirectly mediate the capacity of restored meadows to yield more numerous and
1287 diverse fish assemblages. Moreover, it is unclear how the enhancement of fish communities by
1288 restoring seagrass meadows varies over longer time scales, as studies that compare habitat
1289 provisioning by restored meadows and reference habitats are almost always less than five years
1290 in duration (e.g., Fonseca et al., 1996; Sheridan et al. 2003; Sheridan, 2004; McSkimming et al.
1291 2016). Fish populations are highly variable over space and time (Auth et al., 2020), and so
1292 positive trends in fish abundance or diversity associated with seagrass restoration may be
1293 overwhelmed by seasonal, interannual, and spatial variability in relation to environmental
1294 conditions, such as nearshore oceanography (Shanks, 2020). Quantifying the spatial and
1295 temporal variability in the enhancement of fish communities by restored seagrass meadows is
1296 valuable to guiding restoration and is especially urgent as climate change intensifies a variety of
1297 factors that may have direct and indirect effects on coastal fishes, including ocean warming, sea-
1298 level rise, and storm disturbance (Ummenhofer and Meehl, 2017; Castorani et al., 2018; von
1299 Biela et al., 2019; Aoki et al., 2020a).

1300 Seagrass restoration may be particularly beneficial to juvenile fishes because meadows
1301 commonly function as nursery habitats (Heck et al., 2003; Lefcheck et al., 2019) that offer
1302 shelter from predation (Hindell et al., 2002; Smith et al., 2011) and enhance food availability
1303 (Nakamura and Sano, 2005; Alfaro, 2006). Numerous coastal fishes recruit to seagrass meadows
1304 as larvae and remain in or near meadows until maturity (Rooker et al., 1998; ASMFC, 2010;
1305 Faletti et al., 2019). Seasonal recruitment results in distinct assemblages of juvenile fishes
1306 inhabiting meadows throughout the year (Rooker et al., 1998; Sobocinski et al., 2013). Overlaid

1307 on these seasonal patterns are typically large year-to-year variations in recruitment and
1308 community composition due to variability in climate drivers, currents, and lower trophic-level
1309 production, among other factors (Wood, 2000; Woodland et al., 2021; Schonfeld et al. 2022).
1310 Further, within seagrass meadows, the distributions and diversities of fishes and mobile
1311 invertebrates vary spatially in response to meadow attributes (Yeager et al., 2011, 2016; Cheng
1312 et al., 2022) that are mediated by environmental context. For example, the density and patchiness
1313 of seagrass meadows are affected by wave exposure, current speed, and water depth (Fonseca
1314 and Bell, 1998; Carr et al., 2010, Uhrin and Turner, 2018), and these meadow characteristics
1315 contribute to the structure of associated fish communities (Yeager et al., 2011, 2016). However,
1316 studies have rarely considered the direct and indirect effects of the abiotic conditions associated
1317 with the seagrass niche on fish communities explicitly (e.g., Giacomazzo et al., 2020).

1318 To address these gaps, we determined (1) how seagrass restoration affects fish abundance
1319 and diversity, (2) the extent to which this effect varies among seasons, years, and locations, and
1320 (3) the direct and indirect effects of spatial variation in abiotic conditions on the distribution of
1321 seagrass and the fish community. We carried out a long-term and spatially replicated study (5–7
1322 years at 16 sites) to characterize fish community dynamics across a rapidly expanding restored
1323 seagrass meadow in coastal Virginia, USA. Our findings indicate that despite pronounced
1324 seasonal and interannual variation, seagrass restoration increases fish abundance and diversity,
1325 and that coastal hydrodynamics (water residence time) and bathymetry (depth) constrain fish
1326 communities by mediating seagrass distributions. A greater consideration of how ecosystem
1327 service provisioning varies in response to direct and indirect environmental forcing will improve
1328 the success and value of future restoration projects.

1329 **Methods**

1330 *Study system*

1331 We conducted our investigation in a network of coastal bays along the Atlantic coastline
1332 of Virginia, USA, studied by the Virginia Coast Reserve Long Term Ecological Research
1333 project. The bays are shallow (1 – 2 m below mean sea level, MSL) and punctuated by deep
1334 inlets connecting them to the Atlantic Ocean (McGlathery et al., 2001; Oreska et al., 2021).
1335 Salinities are euhaline (typically >30 PSU) due to limited freshwater inputs from the Delmarva
1336 Peninsula bordering the western margin of the bays (Oreska et al., 2021), and low nutrient
1337 loading from the coastal watersheds results in general oligotrophy and consistently high water
1338 quality (McGlathery et al., 2001; Anderson et al., 2010). These factors combine to provide an
1339 estimated >100 km² of suitable habitat for the dominant species of seagrass in the coastal bays,
1340 eelgrass *Zostera marina* (Oreska et al., 2021).

1341 Seagrass was extirpated from the coastal bays in the 1930s due to wasting disease and
1342 hurricane disturbance. The discovery of a small patch of naturally occurring seagrass spurred the
1343 creation of a large-scale, seed-based restoration program that since 1999 has deposited >75
1344 million seeds across the coastal bays (Orth et al., 2012, 2020). The total area restored now
1345 exceeds 36 km² and the meadows have contributed to improvements in water quality, carbon and
1346 nitrogen sequestration, and finfish and epifaunal invertebrate biomass and diversity (Lefcheck et
1347 al., 2017; Aoki et al., 2020b; Oreska et al., 2020; Orth et al., 2020).

1348 We focused on two restored bays: South Bay (median depth = 0.85 m below MSL) and
1349 Hog Island Bay (median depth = 1 m below MSL). The two bays have similar tidal ranges of
1350 1.32 m in South Bay and 1.24 m in Hog Island Bay (McGlathery et al., 2012). Aerial surveys
1351 (Moore et al., 2009) show that seagrass cover expanded 83% in South Bay (from 10.7 km² to

1352 19.6 km²) and 80% in Hog Island Bay (1.8 km² to 3.3 km²) from the start to end of our study
1353 (2012–2018). We sampled fish at each of 16 fixed locations (sites): 4 vegetated sites within the
1354 South Bay meadow and 4 nearby unvegetated sites outside the meadow; 4 vegetated sites within
1355 the Hog Bay meadow and 4 nearby unvegetated sites (Fig. 1).

1356 Within the two bays, we used observations of seagrass presence at fish sampling sites to
1357 evaluate relationships between abiotic drivers, seagrass, and the fish community. As the initial
1358 sampling design established a total of eight sites inside and eight sites outside the meadows, we
1359 assumed that seagrass was present or absent at these sites until noted otherwise during field
1360 sampling. When unvegetated sites were overtaken by seagrass encroachment or seagrass was lost
1361 from vegetated sites, we assumed that these altered states persisted at each site until another
1362 change was noted (see Appendix 1 for seagrass presence through time).

1363 *Fish sampling*

1364 To examine the effects of seagrass restoration on fish communities over time, we
1365 sampled fish at each site once or twice per year using beach seines (7.6 m wide × 1.8 m tall; 1.5
1366 m deep pocket with 6.4 mm mesh) pulled along 25 m transects in the summer (May or June) and
1367 autumn (September or October) from 2012 through 2018. Researchers stopped sampling at the 4
1368 initially unvegetated sites in South Bay after 2015 when sites were colonized by seagrass,
1369 although seining occurred once more at these sites during the autumn of 2017. We counted,
1370 measured (total length, TL), and identified fish to the lowest possible taxon in the field prior to
1371 release. All seine pulls occurred during the day and within three hours of low tide for logistical
1372 reasons (N = 204).

1373 Catch data were analyzed in terms of catch per unit effort (CPUE), richness per seine
1374 haul, and diversity per seine haul as Hill-Shannon diversity (D) (Hill, 1973). Hill-Shannon

1375 diversity can be interpreted as an index of mean rarity, where a seine haul containing many
1376 species in low abundance (i.e., many rare species) would have a larger Hill-Shannon diversity
1377 than a seine haul with few species in high abundance. D is given by

$$1378 \quad D = \exp\left(-\sum_{i=1}^S p_i \ln(p_i)\right), \quad (1)$$

1379 where p_i is the proportion of individuals belonging to species i and S is the number of
1380 species (Roswell et al., 2021).

1381 All fish sampling methods have some degree of bias. Our sampling method likely
1382 represents a conservative estimate of the effect of seagrass restoration on fish communities
1383 because the capture efficiency of beach seines is generally lower within structured habitat
1384 relative to unstructured habitats (Connolly 1994; Jenkins and Sutherland 1997). Benthic fishes in
1385 seagrass meadows may avoid capture in seine nets because they are positioned deeper within the
1386 vegetation (Connolly 1994). The strength of undersampling varies among species; for instance,
1387 fish associated with the upper seagrass canopy and middle of the water column are sampled in
1388 seines more efficiently than seafloor-associated species that may avoid the net capture area
1389 (Jenkins and Sutherland 1997). Therefore, reduced capture efficiency within the seagrass
1390 meadow likely resulted in our data underestimating fish densities for segments of the fish
1391 community.

1392 *Abiotic variables*

1393 To address our third goal of determining the influence of spatial variation in abiotic
1394 conditions on seagrass and the fish community, we modeled the effects of wind fetch, water
1395 residence time, and bathymetry on seagrass presence/absence, fish CPUE, and fish species
1396 richness. Specifically, we used piecewise structural equation models (SEMs, described below) to
1397 test the hypotheses that wind fetch, water residence times, and depth have direct effects on

1398 seagrass occurrence, which in turn mediates indirect effects between these environmental
1399 variables and fish community responses. SEMs are a class of modeling techniques that allow for
1400 testing hypotheses about causal relationships within multivariate systems (Grace, 2008).

1401 Recent spatial niche modeling in the Virginia coastal bays found that wind fetch and
1402 water residence times are important predictors of seagrass recruitment and survival, respectively
1403 (Oreska et al., 2020). We chose to include bathymetry in SEMs because water depth constrains
1404 seagrass growth through light limitation in turbid habitats (McGlathery et al., 2012; Aoki et al.,
1405 2020a) and desiccation, temperature stress, and wave disturbance cause mortality in shallow
1406 zones (Marsh et al., 1986; Boese et al., 2005; Castorani et al., 2014). We also tested the
1407 hypothesis that water residence times directly affect fish abundance and richness, as sites closer
1408 to inlets may be more suitable for fishes due to the higher stability of abiotic conditions
1409 associated with more frequent tidal flushing (Martino and Able, 2003).

1410 We created the wind fetch predictor by averaging wind fetch lengths within 150 m of
1411 seine sampling sites from raster layers of fetch calculated across the coastal bays for the
1412 summers of 2014 and 2015 (Kremer and Reidenbach, 2021). Fetch was calculated as the
1413 maximum distance wind may travel unobstructed over water in a constant direction and was
1414 weighted by observed wind directions from a nearby meteorological station. Despite variability
1415 in fetch between 2014 and 2015 due to changes in wind direction, fetch lengths at sampling sites
1416 were highly correlated ($R^2 = 0.97$), and so we assumed that the average fetch length across years
1417 was representative of site-specific fetch lengths over the study period. We followed a similar
1418 approach to estimate water residence times, which were developed with a three-dimensional
1419 numerical coastal circulation model (Safak et al., 2015). Here, residence time refers to the
1420 amount of time a Lagrangian particle, released during high tide, spent at a particular location

1421 within the coastal bays before leaving through a tidal inlet (Safak et al., 2015; data from Wiberg
1422 et al., 2015). Variability in residence times within the inlet bays—where our seining sites were
1423 located—is largely mediated by distance to the bay inlet, bay geometry, and tidal phase (Safak et
1424 al., 2015), factors that do not show large interannual variation on the timescale of our study.
1425 Therefore, we assumed that the residence times used here were broadly representative of
1426 residence times over the study period. We extracted depths within 150 m of each site using
1427 bathymetric maps of the Virginia coastal bays (Richardson et al., 2014; Barnes and Wiberg,
1428 2020).

1429 *Statistical analyses*

1430 We used generalized linear mixed models (GLMMs) to quantify the effects of seagrass
1431 (inside vs. outside of meadows) and season (summer vs. autumn) on fish CPUE, richness, and
1432 Hill-Shannon diversity. We also evaluated whether the meadow where sampling occurred
1433 mediated changes in the fish community by modeling the interactions between meadow location
1434 (Hog Island or South Bay) and abundance, richness, and diversity, with sampling sites and years
1435 treated as random intercepts (Zuur et al., 2009). Hill-Shannon diversities had a high frequency of
1436 ones due to seine hauls where zero or one species of fish were caught, so we subtracted one from
1437 calculated catch diversities to model the data as zero-inflated (Zuur et al., 2009) and modeled
1438 non-zero Hill Shannon diversities using the gamma distribution with log-link. We used a
1439 negative binomial distribution with log-link in community- and species-level models of CPUE,
1440 and a Poisson distribution with log-link to model richness. When comparing the catches of
1441 multiple species across habitats or seasons, we adjusted P values according to Benjamini and
1442 Hochberg (1995). We used a residual simulation approach to confirm that fitted model residuals
1443 met expectations of dispersion, zero-inflation, the abundance of outliers, and residual normality

1444 using the *DHARMA* package (Hartig, 2017). We fit GLMMs using the *glmmTMB* package in R
1445 (Brooks et al., 2017).

1446 We next applied permutational multivariate analysis of variance using distance matrices
1447 (PERMANOVA) to the fish catch data to test the hypothesis that fish community composition
1448 differed between meadow locations (South Bay vs. Hog Island Bay) and sampling habitats
1449 (inside vs. outside of meadows) (Anderson, 2005). We log transformed catches according to
1450 $\ln(\text{CPUE} + 1)$ for $\text{CPUE} > 0$ within the community matrix and then converted the catch matrix to
1451 Bray-Curtis dissimilarities for hypothesis testing. We included sampling habitat, meadow
1452 location, season, and year as model covariates, and blocked permutations within sampling sites
1453 ($N = 5000$). We used the R package *vegan* for these analyses (Oksanen et al., 2020).

1454 Lastly, we used piecewise SEMs to evaluate the strengths of relationships among direct
1455 and indirect predictors of CPUE and species richness. We used piecewise SEMs because they
1456 provided for the inclusion of hierarchical data structures and non-Gaussian response distributions
1457 in evaluating model paths (Lefcheck, 2016). For SEMs of both CPUE and richness, we created
1458 path diagrams encoding hypothesized relationships between seagrass presence, physical
1459 variables (fetch, residence time, and depth), and fish community variables (CPUE, richness) (see
1460 *Physical Variables*, above). All covariates in SEMs were scaled and centered prior to model
1461 fitting.

1462 We used Fisher's C to determine if the proposed model structures were appropriate given
1463 the data and that no important paths were excluded (Shipley, 2000; Lefcheck, 2016), in which
1464 large P values derived from Fisher's C ($P > 0.05$) indicated a well-fitting model. We modeled the
1465 probability of seagrass occurrence using a binomial GLMM with a logit-link and modeled CPUE
1466 and richness using negative binomial and Poisson GLMMs with log-links, respectively. We

1467 found standardized path coefficients for the binomial GLMM submodel using the latent-
1468 theoretical approach, and for the Poisson and negative binomial models using the observation-
1469 empirical approach (Grace et al., 2018). We fitted piecewise SEMs using the R package
1470 *piecewiseSEM* and mixed models using *lme4* (Bates et al., 2015; Lefcheck 2016).

1471 **Results**

1472 *Total fish CPUE and diversity*

1473 On average, total fish CPUE was 6.4× greater inside than outside of seagrass meadows
1474 (Wald $\chi_1^2=59.14$, $P < 0.001$). Likewise, estimates of Hill-Shannon diversity were 3.1× greater
1475 (Wald $\chi_1^2=4.68$, $P = 0.03$), and estimates of species richness were 2.6× greater (Wald $\chi_1^2=38.67$,
1476 $P < 0.001$) inside than outside of seagrass meadows (Fig. 2a, c, e). We also found that CPUE,
1477 diversity, and species richness were consistently elevated within the seagrass meadow relative to
1478 outside the meadow regardless of whether samples were collected in Hog Island or South Bay.
1479 However, the interaction model for species richness showed that richness was significantly
1480 higher overall in South Bay compared to Hog Island Bay (Wald $\chi_1^2=5.39$, $P = 0.02$). Summary
1481 tables for community-level statistical models are available in Appendices 2 and 3.

1482 *Community composition across habitats and meadows*

1483 Fish community composition differed between vegetated and unvegetated habitats (SS =
1484 4.61, Pseudo $F = 15.84$, $P = 0.02$). Inside of seagrass meadows, pipefish (*Syngnathus* spp.; N =
1485 610) and silver perch (*Bairdiella chrysoura*; N = 486) were the most common taxa. Outside of
1486 seagrass meadows, anchovies (*Anchoa* spp.; N = 210) and Atlantic silversides (*Menidia menidia*;
1487 N = 78) were most common (Figs. 3 and 4). Together with pinfish (*Lagodon rhomboides*) and
1488 spot (*Leiostomus xanthurus*), these six taxa accounted for 91.8% of the total catch. 28 additional
1489 taxa comprised the remaining 8.2% of catch (see Appendix 4 for complete list). Of the 2664 fish

1490 collected over the 2012–2018 sampling period, 86.7% were collected in seagrass habitats (Fig.
1491 3a). Significant ($P < 0.05$) species-level differences in CPUE between habitats were also evident
1492 in four of the six most common taxa. Species-specific pairwise comparisons showed that 11.7×
1493 more pinfish, 57× more pipefish, 24.8× more silver perch, and 14× more spot were collected
1494 inside vs. outside of seagrass (all $P_{\text{adj.}} < 0.01$; Fig. 4b). Despite increased richness in the South
1495 Bay meadow, we did not detect any differences in community composition between the two
1496 meadows ($P = 0.3$).

1497 *Fish lengths*

1498 The lengths of collected fishes ranged from 2.5 mm (Gobiidae) to 400 mm (bluespotted
1499 cornetfish, *Fistularia commersonii*). We found that the lengths of the majority of Atlantic
1500 silverside, pinfish, silver perch, and spot were below respective lengths-at-maturity as indicated
1501 by the literature (anchovy: Stevenson, 1958; Atlantic silverside: Fay et al., 1983; pipefish: Ripley
1502 and Foran, 2006; pinfish: Ohs et al., 2011; silver perch: Grammer et al., 2009; spot: ASMFC,
1503 2010), suggesting that the majority of collected fish were juveniles or post-larvae (Fig. 5).

1504 *Season and interannual variation in fish communities*

1505 Pairwise comparisons showed that CPUE during the summer months significantly
1506 exceeded those in the autumn (Wald $\chi_1^2 = 7.45$, $P < 0.01$), and catch data suggested that this
1507 effect was mostly due to a decline in October CPUE rather than September (Fig. 2a). In contrast,
1508 Hill-Shannon catch diversity was not significantly different between seasons ($P = 0.8$), although
1509 species-level models of catch across sampling seasons revealed strong within-year patterns in
1510 catch composition. Within the top six most collected taxa, we found that pipefish, pinfish, and
1511 spot were significantly more abundant in the summer rather than autumn months (all $P_{\text{adj.}} <$
1512 0.001; Fig. 4a).

1513 Species-specific catches also varied greatly among years. For example, pinfish made up
1514 69% of catch during June 2015, but were absent from collections in June 2016. Similar patterns
1515 emerged for catches of silver perch, which were absent from May 2017 collections but made up
1516 69% of catch in May 2018. Pipefish were the only taxa present across all years and sampling
1517 months, while Atlantic silversides were present in all sampling years and months except for
1518 October of 2018.

1519 *Seascape-scale environmental drivers*

1520 Fisher's C indicated that no paths were missing from SEMs modeling CPUE ($C = 0.9$, P
1521 $= 0.9$, $df = 4$), but identified a missing path in the richness SEM between depth and richness that
1522 we updated the model to include ($C = 1.04$, $P = 0.6$, $df = 2$). We found that the best fitting sub-
1523 model of seagrass presence included an interaction between depth and residence time ($\beta_{\text{unstd.}} = -$
1524 3.13 , $SE = 1.24$, $P = 0.01$) with a significant main effect of depth ($\beta_{\text{unstd.}} = -2.55$, $SE = 1.0$, $P =$
1525 0.01), indicating that seagrass presence was more likely in shallow areas, and particularly when
1526 those shallower areas were more frequently flushed.

1527 In the CPUE SEM, seagrass presence was the most important contributor to CPUE ($\beta_{\text{unstd.}}$
1528 $= 1.82$, $SE = 0.23$, $P < 0.001$), whereas residence time did not influence catch. However, in the
1529 richness SEMs, seagrass presence ($\beta_{\text{unstd.}} = 0.83$, $SE = 0.13$, $P < 0.001$), residence time ($\beta_{\text{unstd.}} = -$
1530 0.11 , $SE = 0.05$, $P = 0.03$), and depth ($\beta_{\text{unstd.}} = -0.17$, $SE = 0.07$, $P = 0.01$) were each found to be
1531 significant contributors to richness. Standardized coefficients for the richness sub-model showed
1532 that seagrass presence had a larger influence on mean richness ($\beta_{\text{std.}} = 0.44$) than residence time
1533 ($\beta_{\text{std.}} = -0.12$) or depth ($\beta_{\text{std.}} = -0.18$).

1534 **Discussion**

1535 Using a spatially replicated, multiyear seine survey, we found that restored seagrass
1536 meadows were associated with large gains in fish abundance and diversity relative to adjacent
1537 unvegetated areas. Catches within restored meadows were 6.4 times more abundant and 3.1
1538 times more diverse, with substantial differences in community composition attributable to the
1539 presence of seagrass and seasonal patterns in abundance. The seine method that we used
1540 undersamples fishes that closely associate with the seafloor, where they may avoid the net
1541 capture area (Connolly, 1994; Jenkins & Sutherland, 1997); therefore, our assessment of the
1542 benefits of seagrass restoration for the abundance and diversity of juvenile fishes is likely an
1543 underestimate. We found that patterns of enhanced abundance and richness were stable through
1544 both space and time despite environmental variability across seagrass meadows and seasonal-to-
1545 interannual variability in species composition. SEM results showed that seagrass presence was
1546 influenced by coastal hydrodynamics and bathymetry. Deeper areas with longer residence times
1547 (less frequent flushing) were less likely to have seagrass, and an absence of seagrass was
1548 associated with less diverse and abundant catches. By considering the environmental constraints
1549 limiting restoration success and the provisioning of ecosystem services by restored habitats,
1550 practitioners may better optimize site selection during the initial phases of restoration and reduce
1551 the risk of restoring habitats that fail to meet benchmarks for recovering lost ecosystem
1552 functions. For instance, our results from coastal Virginia indicate that the success of seagrass
1553 restoration and the associated benefits for fish production and diversity will be maximized at
1554 shallow, frequently flushed sites; optimal site selection guidelines for other regions will differ
1555 with variation in the relationships between environmental conditions and seagrass restoration
1556 success, and the associated direct and indirect effects on fishes.

1557 Our findings showed that fish were more numerous within the restored seagrass meadows
1558 than outside of them, supporting short-term studies documenting higher densities of fishes in
1559 seagrass meadows relative to unvegetated habitats (Orth and Heck, 1980; Heck et al., 1989;
1560 Arrivillaga and Baltz, 1999; Gilby et al., 2018). Importantly, we found that these patterns held
1561 over long periods (5–7 years) despite spatial, between-year, and within-year variability in
1562 community composition. Catches were largest during the summer months when new and
1563 recently recruited fishes could be found in high densities. For example, pinfish (*Lagodon*
1564 *rhomboides*) recruit to seagrass meadows from offshore spawning grounds in the winter and
1565 spring months (Faletti et al., 2019). Likewise, the most abundant taxa in our collections, pipefish
1566 (*Syngnathus* spp.), migrate from offshore into estuaries during the spring and summer where
1567 spawning and brooding of eggs occurs (Campbell and Able, 1998). We also documented spot
1568 (*Leiostomus xanthurus*) in higher abundance in seagrass meadows during the summer months,
1569 which aligns with the timing of spot larval migration from offshore into nearshore nurseries
1570 (spring to early summer in the Mid-Atlantic) (ASMFC, 2010).

1571 Silver perch (*Bairdiella chrysoura*) were also collected in higher abundance within the
1572 meadow, although catches were similar across seasons, aligning with previous work from the
1573 northern Gulf of Mexico identifying silver perch as permanent residents of estuaries with a peak
1574 spawning period between mid-March and June (Grammer et al., 2009). Similarly, there was a
1575 weakly significant, positive effect of meadow presence on Atlantic silverside abundance ($P_{Adj.} =$
1576 0.06), and silverside catches did not differ between seasons. Within the U.S. mid-Atlantic region,
1577 both juvenile and adult Atlantic silverside are found in coastal habitats in high abundance in all
1578 seasons except winter, when migration to deeper or offshore waters occurs (Fay et al., 1983).
1579 Our results and the life histories of these species suggest that silver perch and Atlantic silverside

1580 may contribute less to seasonal differences in community composition than other species
1581 inhabiting the meadows.

1582 Length-frequency distributions for these common fishes suggest that the majority of
1583 collected Atlantic silverside, pinfish, silver perch, and spot were juveniles, with fish lengths also
1584 tending to be smaller during the summer months than in the autumn (Appendix 5). This
1585 interpretation was less well-supported for pipefish, possibly because two species of pipefish with
1586 distinct length-frequencies—*Syngnathus fuscus* and *Syngnathus floridae* (Ripley and Foran,
1587 2006)—are found in this region and we only identified pipefish to genus. Given increases in
1588 water temperatures in the Mid-Atlantic that favor the more southerly distributed *S. floridae*
1589 (Sobocinski et al., 2013), understanding the relative abundance dynamics of these species in the
1590 restored meadows may provide insights into the responses of seagrass-associated fish
1591 communities to climate change. While our findings suggest that the restored meadows act as
1592 preferred habitat for juvenile fishes over nearby unvegetated areas, the vegetated-versus-
1593 unvegetated dichotomy oversimplifies the complexity of fish movement patterns within the
1594 meadow and broader seascape mosaic (Nagelkerken et al., 2015; Litvin et al., 2018). The coastal
1595 bay seascape is composed of a variety of structured habitats used by fishes in addition to seagrass
1596 meadows, including intertidal marshes and oyster reefs, as well as hydrodynamically and
1597 bathymetrically complex areas like oceanic inlets, channels, and tidal creeks. The proximity of
1598 these habitats to seagrass meadows, as well as the capacity for fishes to move between habitats,
1599 influences fish abundance and biodiversity patterns (Yeager et al., 2016; Gilby et al. 2018).
1600 Expanding these analyses to consider interactions between restored meadows and features of the
1601 seascape more broadly would further refine our understanding of how environmental context
1602 shapes ecosystem services associated with seagrass restoration.

1603 SEM analyses showed how seascape-scale environmental covariates affected fish-habitat
1604 relationships indirectly by influencing seagrass distributions. We found a negative effect of depth
1605 on the probability of seagrass occurrence, indicating a lower likelihood of seagrass presence in
1606 deeper areas, where deep, turbid water restricts seagrass growth through light limitation (Carr et
1607 al., 2010; Aoki et al., 2020a). Further, we identified an interaction effect of depth and water
1608 residence time on the probability of seagrass occurrence, indicating a greater likelihood of
1609 seagrass presence in shallower areas, especially those with frequent tidal flushing. Frequently
1610 flushed areas tend to have larger sediment grain sizes that are favorable for seagrass growth
1611 (Koch, 2001; Wiberg et al., 2015) and may be less susceptible to the disturbance effects of
1612 marine heatwaves through more frequent exchange with the cooler coastal ocean (Aoki et al.,
1613 2021, Berger et al. 2020).

1614 While we show that depth and an interaction between depth and residence time mediate
1615 seagrass presence and habitat provisioning for juvenile fishes, our analyses were limited by only
1616 considering seagrass presence within the immediate vicinity of sampling sites. Seagrass-
1617 associated fish communities will also organize along gradients of seagrass patchiness and areal
1618 coverage (Yeager et al., 2011, 2016; McCloskey and Unsworth, 2015). In the species richness
1619 SEM, we found that depth had a direct negative effect on species richness, indicating a greater
1620 diversity of fishes at shallower sites. This could be due to correlations between depth and
1621 unmeasured characteristics of the meadow, such as meadow area, patchiness, or shoot density
1622 (Aoki et al., 2020a; Belgrad et al., 2021). For example, in the shallower and more frequently
1623 flushed South Bay meadow, average seagrass shoot densities were 4.4× greater than in Hog
1624 Island Bay during the study period (365 vs. 84 shoots/m²; Appendix 6; McGlathery 2017; Aoki
1625 et al., 2020a). Alternatively, physical properties of the water column that scale with depth, such

1626 as light availability or water temperature (Aoki et al., 2020a) may alter local habitat suitability
1627 and subsequent community structure (Belgrad et al., 2021).

1628 We also found that catches at sites with shorter water residence times (more frequent tidal
1629 flushing) tended to be more species rich, a pattern largely due to high richness in the South Bay
1630 meadow where residence times were approximately three hours. Comparatively, residence times
1631 at all other sites exceeded 20 hours (Appendix 7). Beyond the consistent presence of seagrass in
1632 South Bay, this finding may be due to higher stability in environmental conditions afforded by
1633 more frequent flushing by the coastal ocean (Martino and Able, 2003), which has been
1634 hypothesized to mitigate the negative effects of marine heatwaves on seagrass (Aoki et al.,
1635 2021). Alternatively, more frequent water exchange with the coastal ocean may simply introduce
1636 more larvae into the meadow, where the presence of dense structured habitat (Grol et al., 2011)
1637 and an abundance of food (Lefcheck et al., 2017) could improve survivorship of recruits and lead
1638 to a more diverse species assemblage relative to Hog Island Bay.

1639 Through long-term study, we discovered a diverse and abundant coastal fish community
1640 that is highly dynamic over space and time yet is consistently enhanced by restored seagrass.
1641 Such increases in abundance and diversity add to the numerous ecosystem services shown to
1642 recover with large-scale seagrass restoration (Orth et al. 2020). Continued long-term study may
1643 further contextualize how local to regional scale environmental variation mediates the benefits of
1644 seagrass restoration for fish production and biodiversity. The acceleration of marine habitat
1645 restoration in recent decades (Duarte et al. 2020) should serve as motivation for new studies to
1646 evaluate the functional services of restored habitats, advise restoration planning, and inform
1647 policies that incentivize restoration (e.g., Smith and Castorani, 2023). For example, the seagrass
1648 restoration that we studied will soon be the first of its kind to enter a ‘blue carbon’ market with

1649 payments for ecosystem services, which involves a spatial accounting of carbon sequestered by
1650 the restored meadow (Oreska et al., 2021; Shilland et al., 2021). Understanding how ecosystem
1651 services beyond carbon sequestration, such as provisioning habitat and supporting biodiversity,
1652 could contribute to such markets is nascent but interest in these services is growing (Shilland et
1653 al. 2021).

1654 **Acknowledgements**

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1657 DEB-1832221). We thank the many participants of the VCR LTER for their assistance in
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1660 and M. Cornish for comments that improved this manuscript.

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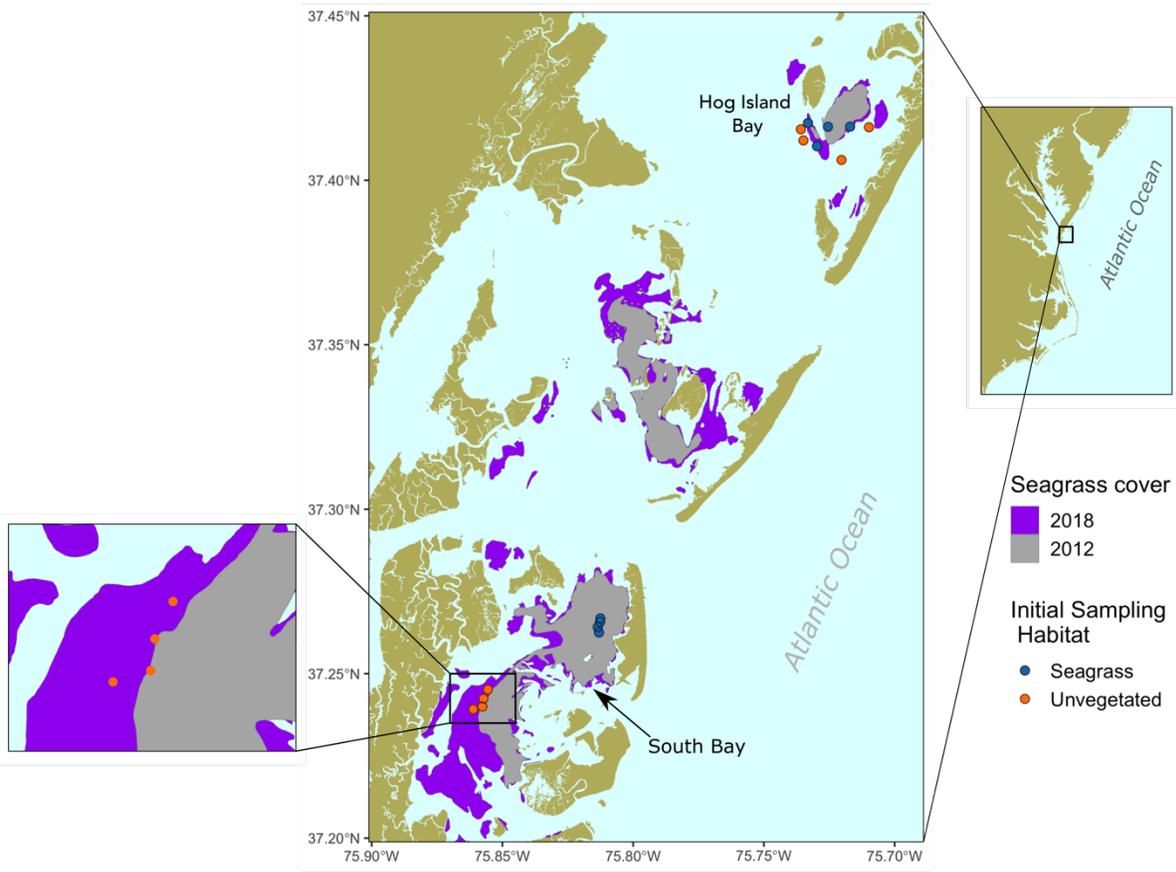
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1670 **Figures**

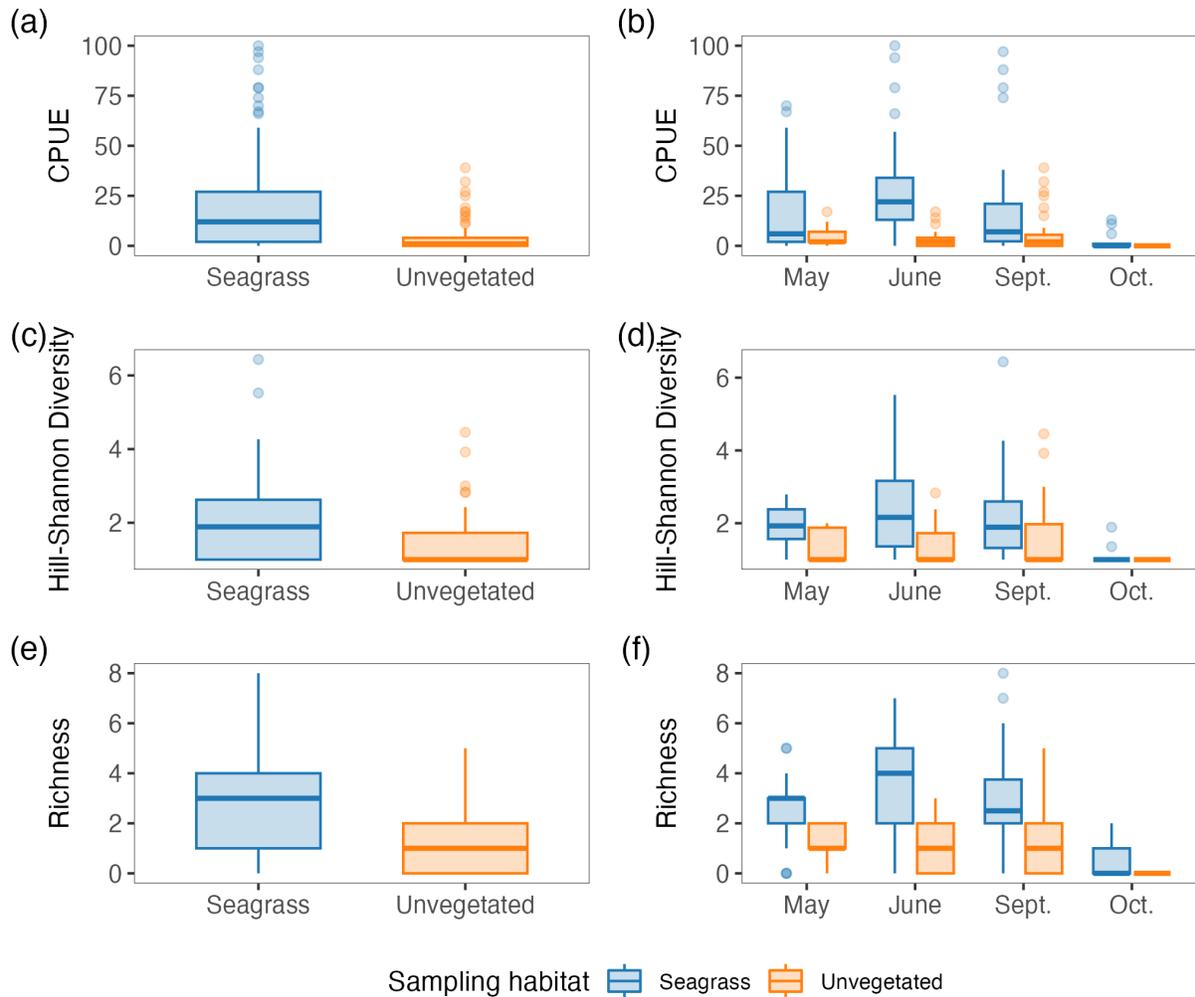


1671
1672 *Figure 1.* Fish sampling sites in Hog Island and South Bay seagrass meadows within the Virginia
1673 Coast Reserve (VCR) along the Atlantic coast of Virginia, USA. Grey and purple shading shows
1674 the extent of seagrass in 2012 and 2018, respectively, from aerial imagery. Points show sampling
1675 sites. The color of points indicate the presence (blue) or absence (orange) of seagrass at sampling
1676 sites when fish surveys began in 2012. Note that some sites that were initially unvegetated
1677 became vegetated during the study period (e.g., see bottom left inset).

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1683 *Figure 2.* Boxplots showing fish catch per unit effort (CPUE; a, b), Hill-Shannon diversity (c, d),

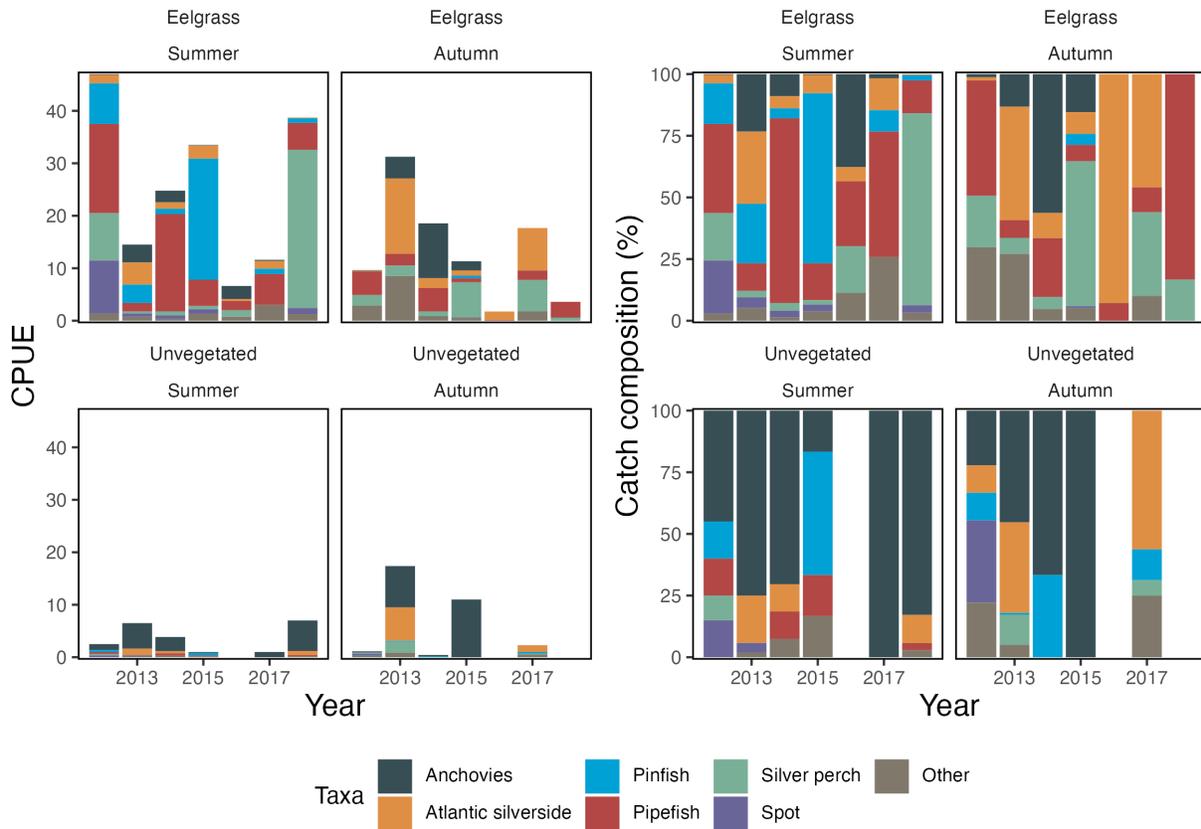
1684 and species richness (e, f) inside and outside seagrass meadows across all sampling months (a, c,

1685 e) and specific to each sampling month (b, d, f). Boxplots include all years of data ($N_{\text{May}} = 28,$

1686 $N_{\text{June}} = 72, N_{\text{Sept.}} = 80, N_{\text{Oct.}} = 24$), and show median values (bold lines) and interquartile ranges

1687 (IQR; boxes), with outliers (transparent points) greater than $1.5 \times$ the interquartile range

1688 (whiskers).



1689

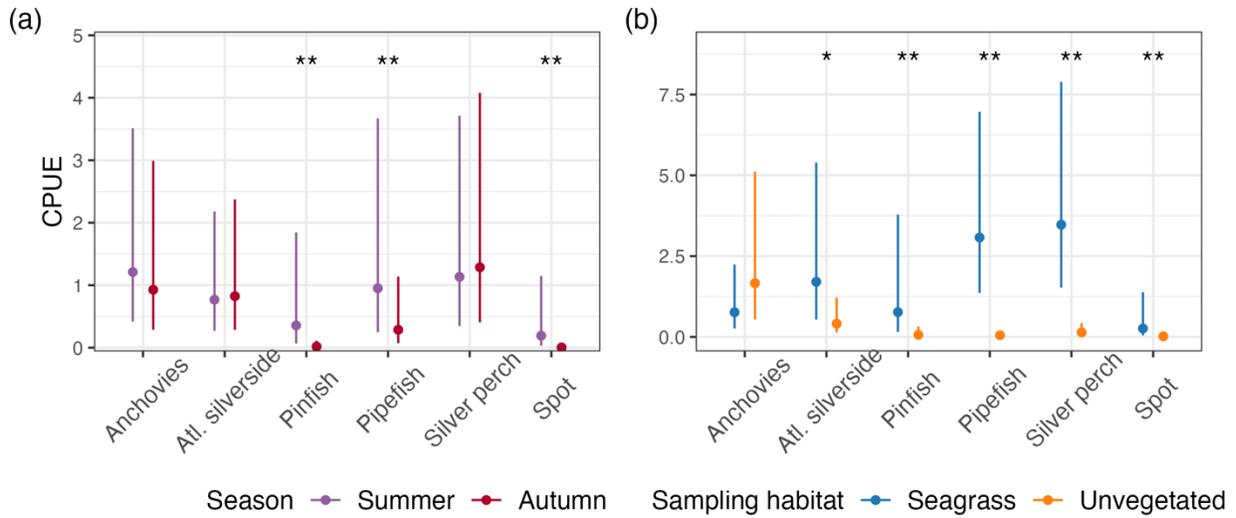
1690 *Figure 3.* Catch per unit effort (CPUE) of the six most common taxa and all other taxa (“Other”)

1691 in the summer (May and June) and autumn (September and October) inside and outside of

1692 restored seagrass meadows. Panels (a) and (b) show the total and proportional contribution,

1693 respectively, of taxa to aggregate CPUE. Empty bars indicate a year-season-habitat combination

1694 in which sampling occurred but no fish were captured.



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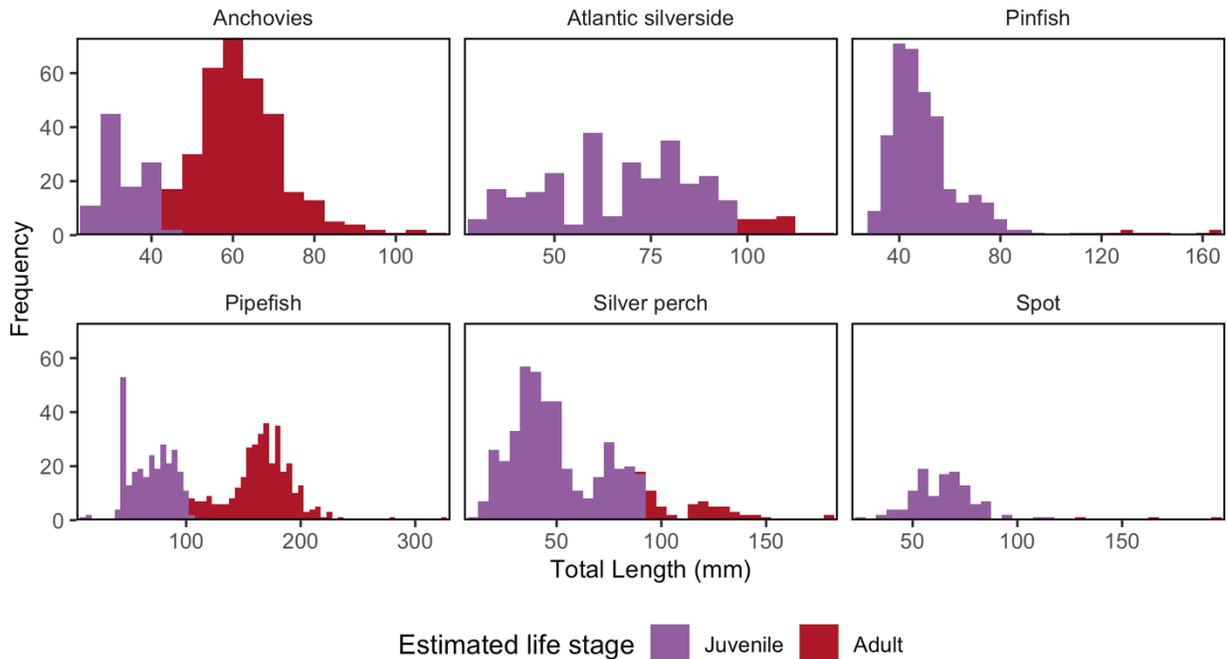
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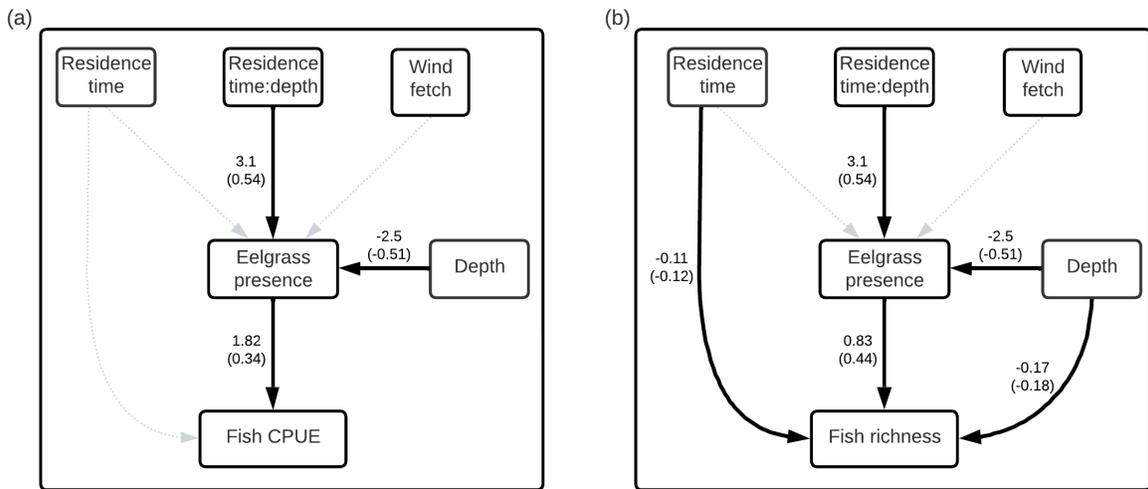
Figure 4. Predicted catch per unit effort (CPUE) for the six most commonly collected taxa in summer and autumn months (a) and inside and outside of seagrass habitat (b). Points indicate marginal means and bars show 95% confidence intervals. Asterisks indicate significant differences in mean CPUE between seasons or habitats at $P < 0.1$ (*) and $P < 0.05$ (**) after adjusting for multiple comparisons.

1702



1703

1704 *Figure 5.* Length-frequency histograms for the six most common taxa in seine collections,
1705 segregated into estimated life stages based on available estimates of lengths at maturity (purple
1706 vs. red). Life stage splits are reported as total lengths-at-maturity except for pipefish and silver
1707 perch where length-length conversions were unavailable and reported as standard lengths-at-
1708 maturity. The pipefish length-at-maturity is the minimum of reported lengths for mature
1709 *Syngnathus floridae* females (Ripley and Foran 2006). For anchovies, we report length-at-
1710 maturity for the highly abundant bay anchovy, *Anchoa mitchilli*. These data represent lengths
1711 measured during all sampling periods and locations both inside and outside of seagrass.



1712

1713 *Figure 6.* Path diagrams for community-level fish catch per unit effort (CPUE; a) and taxa

1714 richness (b). Black arrows with solid lines are significant at $P < 0.05$. Grey dotted lines indicate

1715 non-significant paths. Coefficients are reported next to black lines where $P < 0.05$; standardized

1716 coefficients are included in parentheses next to unstandardized coefficients. The interaction

1717 between residence time and depth is indicated by 'Residence time:depth'.

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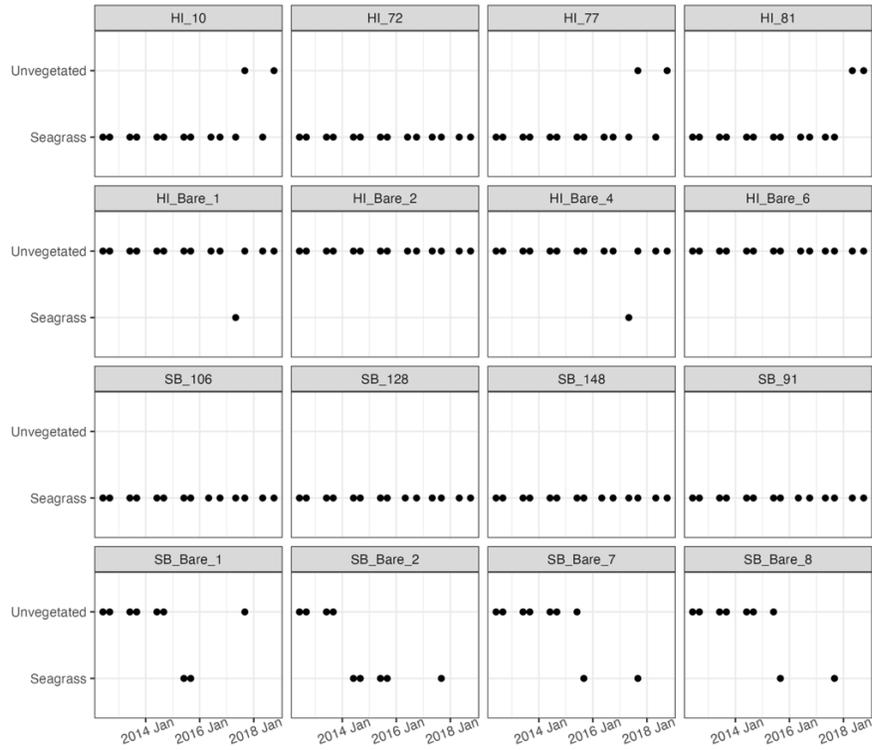
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1728 **Appendix**



1729

1730 *Appendix 1.* Time series of presence-absence of seagrass at each seine sampling site from 2012-
 1731 2018. Each subplot represents a different site, and those with “Bare” in the subplot title were
 1732 unvegetated in 2012 at the beginning of the study. “SB” or “HI” refers to the South Bay or Hog
 1733 Island Bay meadows, respectively. Note that sampling did not occur consistently at initially
 1734 unvegetated South Bay sites after 2015.

1735

1736 *Appendix 2.* Model summaries from comparisons of catch abundance, Hill-Shannon diversity,
 1737 and species richness inside and outside of seagrass, the interaction between habitat and meadow
 1738 location on species richness, and the effect of sampling season on total CPUE. Factor reference
 1739 levels are given in brackets, and reported P values are those returned by Wald z tests.

Model 1: Catch (across species) inside vs. outside seagrass

<i>Predictors</i>	<i>Log-Mean</i>	<i>95% CI</i>	<i>Wald χ_1^2</i>	<i>Wald z</i>	<i>P</i>
Intercept	2.87	2.27 – 3.46	-	-	<0.001
Vegetated vs. unvegetated [Unvegetated]	-1.86	-2.33 – -1.38	59.14	-7.69	<0.001
Random Effects					
σ_{site}^2	0.04				
σ_{year}^2	0.49				

Model 2: Catch diversity (Hill-Shannon) inside vs. outside seagrass

<i>Predictors</i>	<i>Log-Mean</i>	<i>95% CI</i>	<i>Wald</i>	<i>Wald z</i>	<i>P</i>
Count Model					
Intercept	0.31	0.08 – 0.54	-	2.63	0.008
Vegetated vs. unvegetated [Unvegetated]	-0.29	-0.55 – -0.03	4.68	-2.16	0.03

Probability Model

Intercept	-1	-1.41 – -0.59	-	-2.86	0.004
Vegetated vs. unvegetated [Unvegetated]	1.76	1.15 – 2.38	-	5.67	<0.001

Random Effects

σ_{site}^2	0
σ_{year}^2	0.07

Model 3: Catch richness inside vs. outside seagrass

<i>Predictors</i>	<i>Log-Mean</i>	<i>95% CI</i>	<i>Wald χ_1^2</i>	<i>Wald z</i>	<i>P</i>
Intercept	0.9	0.54 – 1.25	-	4.96	<0.001
Vegetated vs. unvegetated [Unvegetated]	-0.95	-1.25 – -0.65	38.67	-6.22	<0.001

Random Effects

σ_{site}^2	0.02
σ_{year}^2	0.17

Model 4: Catch richness inside vs. outside seagrass (× meadow)

<i>Predictors</i>	<i>Log-Mean</i>	<i>95% CI</i>	<i>Wald χ_1^2</i>	<i>Wald z</i>	<i>P</i>
Intercept	0.77	0.40 – 1.14	-	4.1	<0.001

Vegetated vs. unvegetated [Unvegetated]	-0.93	-1.29 – -0.58	47.00	-5.16	<0.001
Meadow [SB]	0.26	0.02 – 0.51	5.39	2.12	0.034
Vegetated vs. unvegetated [Unvegetated] × Meadow [SB]	-0.05	-0.55 – 0.46	0.03	-0.182	0.856

Random Effects

σ_{site}^2	0
σ_{year}^2	0.17

Model 5: Catch (across species) by season (summer vs. autumn)

<i>Predictors</i>	<i>Log-Mean</i>	<i>95% CI</i>	<i>Wald χ_1^2</i>	<i>Wald z</i>	<i>P</i>
Intercept	2.39	1.68 – 3.10	-	6.56	<0.001
Season [Fall]	-0.62	-1.06 – -0.17	7.45	-2.73	0.006

Random Effects

σ_{site}^2	0.72
σ_{year}^2	0.44

1741 *Appendix 3. Summary from PERMANOVA test for differences in sample species composition*
 1742 *inside vs. outside seagrass habitat, across seasons and years.*

PERMANOVA					
<i>Predictors</i>	<i>DF</i>	<i>Sum of Sqs.</i>	<i>Pseudo-R²</i>	<i>F</i>	<i>P</i>
Vegetated vs. unvegetated	1	4.61	0.09	15.84	0.02
Season	1	1.72	0.03	5.91	<0.001
Meadow (HI vs. SB)	1	1.27	0.02	4.37	0.32
Year	1	0.72	0.01	2.4	<0.01
Residual	148	43.1	0.82		
Total	152	52.58	1		

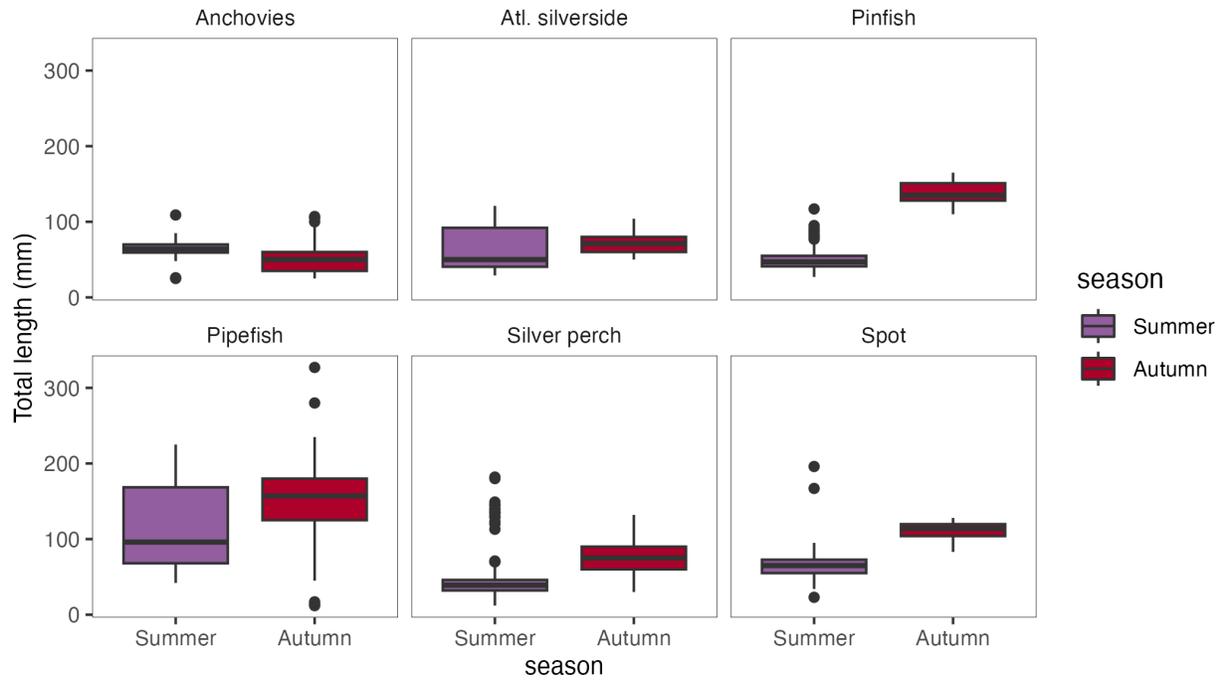
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1745 *Appendix 4. Catches inside and outside of eelgrass meadows with mean, minimum, and*
 1746 *maximum lengths for all collected fishes.*

Fish taxa	Unvegetated	Eelgrass	Mean length (mm)	Min. length (mm)	Max length (mm)
<i>Syngnathus</i> spp.	8	610	123	12	327
<i>Bairdiella chrysoura</i>	20	486	56.3	12	182
<i>Anchoa</i> spp.	210	220	55.7	25	109
<i>Menidia menidia</i>	78	333	68.8	29	121
<i>Lagodon rhomboides</i>	11	352	52.4	27	165
<i>Leiostomus xanthurus</i>	8	110	66.9	23	196
<i>Archosargus probatocephalus</i>	1	43	76.8	53	115
<i>Orthopristis chrysoptera</i>	0	37	103.7	15	127
<i>Micropogonias undulatus</i>	1	34	36.6	21	240
<i>Centropristis striata</i>	1	17	86.2	20	160
<i>Hippocampus erectus</i>	1	14	92.9	80	111
Gobiidae	4	10	36.6	2	61
<i>Eucinostomus argenteus</i>	1	6	45.3	31	88
<i>Chilomycterus schoepfii</i>	0	5	167.6	80	236
<i>Pomatomus saltatrix</i>	2	2	92	65	155
<i>Tautoga onitis</i>	0	4	74.5	18	105
<i>Fistularia commersonii</i>	1	2	325	195	400

Fish taxa	Unvegetated	Eelgrass	Mean length (mm)	Min. length (mm)	Max length (mm)
Mugilidae	0	3	83	55	109
<i>Paralichthys dentatus</i>	0	3	78.7	71	86
Sciaenidae	0	3	30	30	30
<i>Sphoeroides maculatus</i>	0	3	37.7	13	80
<i>Strongylura marina</i>	1	2	171.7	80	285
<i>Hyporhamphus meeki</i>	0	2	84	71	97
<i>Lolliguncula brevis</i>	2	0	43	35	51
<i>Lutjanus griseus</i>	0	2	58	56	60
<i>Opsanus tau</i>	0	2	113.5	112	115
<i>Stephanolepis hispidus</i>	0	2	35	30	40
<i>Brevoortia tyrannus</i>	0	1	52	52	52
<i>Chasmodes bosquianus</i>	0	1	60	60	60
<i>Cynoscion nebulosus</i>	0	1	38	38	38
<i>Gobiesox strumosus</i>	1	0	120	120	120
<i>Peprilus triacanthus</i>	1	0	10	10	10
<i>Sphyaena borealis</i>	0	1	75	75	75
<i>Stenotomus chrysops</i>	1	0	174	174	174
Total	353	2,311	-	-	-

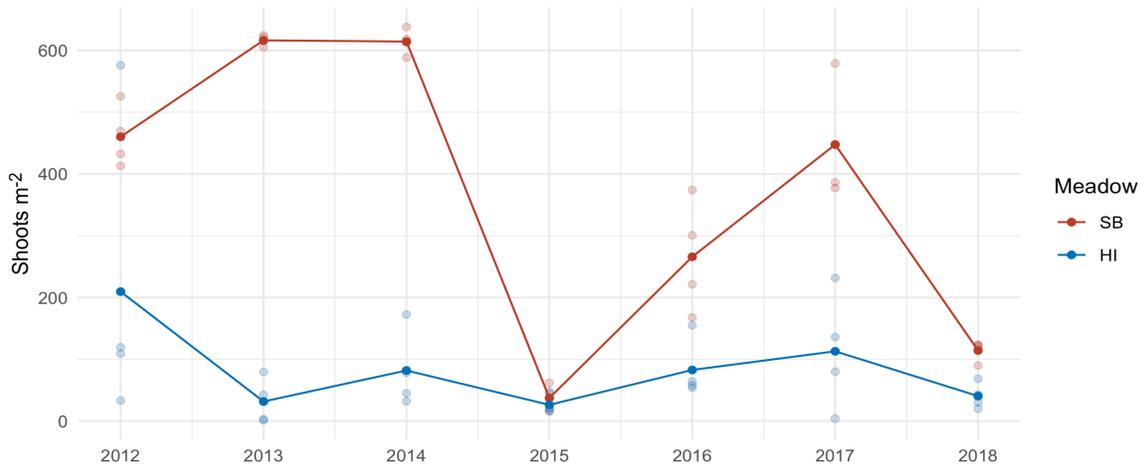


1748

1749 *Appendix 5.* Boxplots showing the distribution of fish lengths (TL) in the summer (purple) and
 1750 autumn (red) months by most common species. The data shown here include fish collected both
 1751 inside and outside the meadow and across years/seasons.

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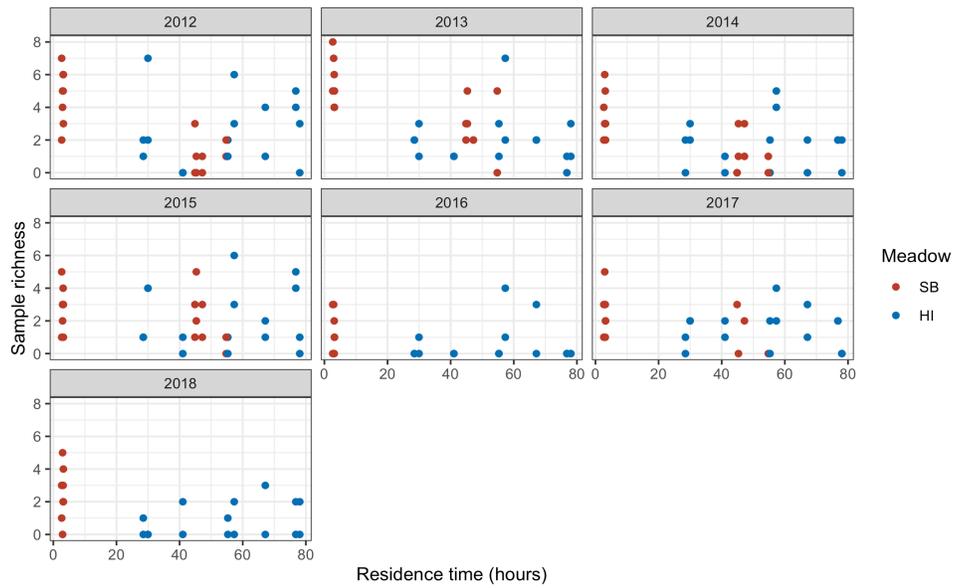
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1755 *Appendix 6.* Eelgrass shoot densities at sites within South Bay (red) and Hog Island Bay (blue
 1756 where fish seining occurred between 2012-2018. Density observations were made annually
 1757 during July. Bold points indicate annual mean densities. See Aoki et al. (2020) for more
 1758 information.

1759



1760

1761 *Appendix 7. Sample richness against residence times across years and meadows. Samples*

1762 *collected in South Bay are shown in red and Hog Island Bay in blue.*

1763 Discussion

1764 In this work, I explored how SST anomalies influence inshore fish populations on the
1765 spatial scale of the Mid-Atlantic Bight (MAB), how stability and asynchrony within commercial
1766 fisheries relates to asynchrony in the ecological communities that they depend upon in the
1767 Chesapeake Bay, and how seagrass restoration enhances the abundance and diversity of juvenile
1768 fishes in the coastal bays of Virginia. A unifying aspect of these studies is that they yield insights
1769 into the give and take of societal interactions with fish populations. Society manages the
1770 extraction of fish biomass from inshore ecosystems while also advocating for the restoration of
1771 fish habitats that support these populations. The effects of ocean warming on fish populations,
1772 which are an indirect effect of societal greenhouse gas emissions (Frölicher et al., 2014), are
1773 another pathway through which humans have impacted inshore fish populations. However, while
1774 these research foci are connected in the sense that they are all related to human drivers of
1775 ecological change, the differing spatial scales of their human-ecological interactions cannot be
1776 overstated.

1777 In Chapter 1, I showed that the responses of inshore fish population dynamics within the
1778 MAB were weakly mediated by species life history characteristics, displaying high variability in
1779 their responses to warming from bay to bay. However, the species whose populations responded
1780 positively to warming and maintained the “fastest” life histories, meaning that they were smaller-
1781 bodied and faster-lived, tended to be warm-adapted. Alternatively, the species whose populations
1782 responded negatively to warming and were the “slowest” lived, or were larger-bodied and
1783 longer-lived, tended to be cool-adapted. For species whose populations are “moving” northwards
1784 in the MAB (I was not able to differentiate between the possible drivers of population responses,

1785 so an increase/decline could indicate the movement of individuals to new habitats or changes in
1786 productivity or phenology) along with their suitable thermal habitats, the conservation and
1787 restoration of nursery habitats like seagrass meadows is unlikely to be enough to prevent
1788 population declines in the long-term. This is because the spatial scale of seasonally averaged
1789 warming occurs on hundreds to thousands of kilometers, whereas habitat restoration, even when
1790 hugely successful, falls on the scale of tens of kilometers, and can be further complicated by
1791 temperature variability associated with climate change like a greater occurrence of marine
1792 heatwaves (Spillman et al. 2021). This is not to say that seagrass restoration efforts are unlikely
1793 to benefit fish communities in a warming climate; quite the opposite, as many warm-adapted
1794 species whose populations are increasing will likely benefit from an abundance of inshore
1795 nursery habitat in the future. However, the composition of species occupying inshore nursery
1796 habitats may shift towards those species that are more well-suited to the local thermal
1797 environment.

1798 Nowhere is confronting and understanding variability in fish population dynamics more
1799 important than in fisheries management. In the Chesapeake Bay, for example, many of the most
1800 commercially valuable finfish move into and out of the Bay mainstem seasonally, and in recent
1801 years, the biomasses of these species have declined, particularly in the southern regions of the
1802 Bay. These declines have been attributed to reductions in the exchange of finfish biomass with
1803 the adjacent continental shelf due to climatic variability—e.g., the North Atlantic Oscillation and
1804 its impacts on regional oceanographic dynamics—processes that manifest on spatial scales far
1805 exceeding that of Chesapeake Bay, but also due to coastwide declines in abundance for species
1806 like Atlantic croaker (Schonfeld et al. 2022).

1807 In Chapter 2, I showed that these biomass declines mitigated the stabilizing effects of
1808 within-year harvest diversification for the commercial fishing portfolio in the Virginia portion of
1809 Bay, as the within-Bay biomasses of targeted species in this region displayed synchronous
1810 declines on interannual time scales, and small-scale fishing operations in the Bay typically do not
1811 extend their operations into adjacent offshore habitats where populations may be more abundant.
1812 Stated differently, the within-year stability of commercial harvest portfolios in the lower Bay
1813 benefit less from diversification if the species in the portfolio, while retaining unique migratory
1814 patterns leading to within-year biomass asynchrony, are synchronously declining in their within-
1815 Bay availabilities from year to year. This situation requires creative management strategies to
1816 ensure the stability of commercial fisheries operating in the Virginia portion of the Bay. These
1817 strategies are being enacted in the form of legislative initiatives supporting the development of
1818 the blue catfish fishery, but other approaches like management strategies that support fishers
1819 transitioning from finfish harvesting to related industries like shellfish and crab fisheries are
1820 likely to be beneficial for the stability of the regional harvest portfolio.

1821 In addition to the population dynamics of targeted fishes impacting the stability of the
1822 commercial harvest portfolio in the Virginia region of the Bay, we found that the stability of the
1823 harvest portfolio in both Virginia and Maryland was impacted by fishing behaviors in the form
1824 of declines in harvest effort in Virginia and target species diversification during the striped bass
1825 closure in Maryland. In Virginia, the social context of this declining effort is important to
1826 consider, and declines in fishing effort may also be rooted in changes in the composition of the
1827 fishing fleet, as recent research has shown that the number of licensed fishers in Virginia has
1828 declined in recent decades (White & Scheld, 2021), but also changes in market forces and
1829 demand for certain fishes. In Maryland, behavioral patterns in fishing effort were stabilizing for

1830 the portfolio, but within-year harvest portfolio effects did not translate to economic stability:
1831 during the striped bass closure, the targeting of other species like Atlantic croaker, white perch,
1832 and gizzard shad created harvest compensation that stabilized within-year fishery harvests, but
1833 the relatively low value of these species compared to striped bass ultimately meant that
1834 diversification during the closure did not stabilize the within-year value of the harvest portfolio.
1835 These findings showed that asynchrony in population dynamics was not guaranteed to be
1836 associated with asynchrony and stability in the portfolio of harvests derived from those
1837 populations, as social and management-related factors can influence how fishing effort is
1838 distributed across target stocks.

1839 For many fishes that inhabit inshore environments for a portion or all of their life cycles,
1840 the use of nursery habitats like seagrasses enhances the survivorship of juveniles (Lefcheck et al.,
1841 2019a) and so indirectly supports the fisheries that depend on those species (Unsworth et al.,
1842 2019). For this reason, and that the high diversity and abundance of invertebrates and non-target
1843 fish species that inhabit seagrasses contribute to fisheries production through trophic pathways
1844 (Jackson et al., 2001), funding seagrass restoration efforts is one way that governments can
1845 support commercial and recreational fisheries. In Chapter 3, I considered how fish biodiversity
1846 and abundance differed between restored seagrass (*Zostera marina*) meadow and unvegetated
1847 habitats in the coastal bays of Virginia. I showed that both the abundance and diversity (species
1848 richness and Hill-Shannon diversity) were significantly higher within the restored meadow
1849 compared to unvegetated areas. I also found that this pattern of increase abundance and diversity
1850 was temporally stable throughout the year when sampling occurred, and that most of the
1851 collected fish were juveniles. In the context of my other chapters, this work showed that
1852 restoring seagrass meadows was effective for increasing local fish biodiversity in the coastal

1853 bays of Virginia, although the composition of species inhabiting the restored meadows may
1854 change under continued ocean warming.

1855 In this dissertation, I have provided novel insights into how a warming continental shelf
1856 has impacted populations of fishes in the inshore habitats of the MAB, how within-year species
1857 asynchrony among commercially fished species in Chesapeake Bay impacts the stability of
1858 fisheries portfolios associated with those species, and lastly, that restored seagrasses in the
1859 coastal bays of Virginia were associated with substantial and temporally consistent increases in
1860 fish biodiversity and abundance within those meadows relative to nearby unvegetated areas.
1861 Understanding these interactions is essential to the successful management of these highly
1862 interconnected systems in the Anthropocene.

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